



The Abdus Salam
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Advanced School on Non-linear Dynamics and Earthquake Prediction

28 September - 10 October, 2009

**Space-Time Earthquake prediction:
The Error Diagrams**

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Space-Time Earthquake prediction: The Error Diagrams

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PAGEOPH (2009), accepted

OUTLINE

- The simplest space-time prediction problem
 - Generalized error diagrams
 - Statistical problems
 - Prediction efficiency at the research stage

MOTIVATION

The main prediction characteristics:

- ◆ n rate of failures-to-predict
- ◆ τ rate of alarm time

Nonuniqueness of τ for space-time prediction:

$$\tau = \frac{1}{T} \int_0^T \mu(G_t) dt, \quad T \gg 1$$

G_t – alarm zone at t

$$\mu(G) = \begin{cases} \text{area of } G \\ \text{rate of target events in } G \\ \text{population in } G \end{cases}$$



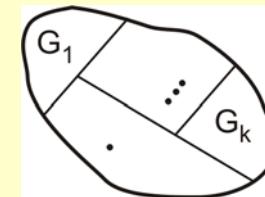
Non-adequate interpretation of predictability in terms of (n, τ)

THE SIMPLEST SPACE-TIME PREDICTION PROBLEM

NOTATION

$G = \bigcup_1^k G_i$ – area, $I(t)$ – data for prediction at t

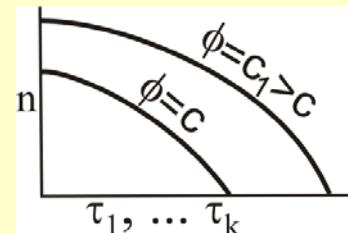
$\pi = \{ \pi(t) \}$ – **Strategy** based on $I(t)$



Solution $\pi(t | G_i) = \begin{cases} \text{alarm} & \text{in } G_i \times (t, t + \Delta), \Delta \ll 1 \\ \text{no alarm} & \end{cases}$

$(n, \tau_1, \dots, \tau_k)$ – **Errors**, τ_i – rate of alarm time in G_i

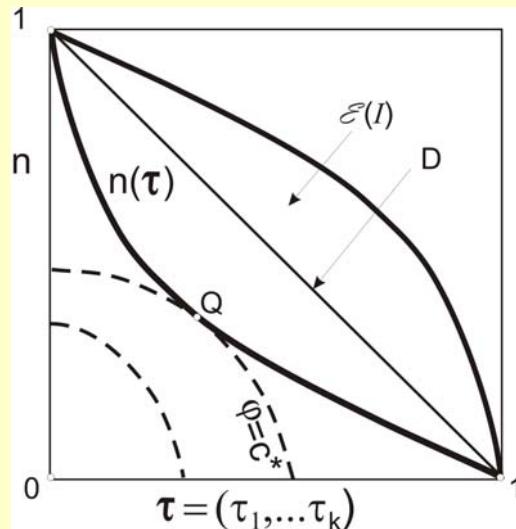
Goal: $\varphi(n, \tau_1, \dots, \tau_k) \Rightarrow \min$



PROBLEM: π^* (optimal strategy) – ?

$$\inf_{\pi} \varphi - ?$$

SOLUTION: the simplest prediction problem



$\mathcal{E}(I) = \{(n, \tau_1, \dots, \tau_k)_\pi\}$ error set, convex

$D = \{n + \tau_\lambda = 1\}$ trivial strategies (random guessing)

$$\tau_\lambda = \sum_{i=1}^n \tau_i \lambda_i / \Lambda$$

$$\lambda_i \quad \text{rate of target events in } G_i \\ \Lambda = \sum \lambda_i \quad \text{rate of target events in } G$$

$n = n(\tau_1, \dots, \tau_k)$ – correct version of the error diagram

- ♦ $\inf_{\pi} \varphi = \varphi(Q)$

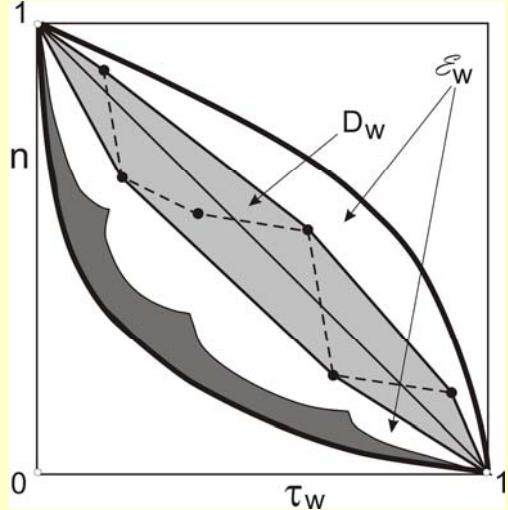
- ♦ $\pi^*(t/G_i) = \text{alarm}$ in $G_i \times (t, t + \Delta)$ as soon as $r(t/G_i) \geq r_i^0$

$$r(t | G_i) = P\{dN(t | G_i) = 1 | I(t)\} / \Delta \quad \text{hazard function for } G_i$$

$$r_i^0 = b_i / a \quad \text{for} \quad \varphi = a \Lambda n + \sum \tau_i b_i$$

- ♦ $n(\tau_1, \dots, \tau_k) = \sum \lambda_i n(\tau_i | G_i) / \Lambda$

REDUCED ERROR DIAGRAM



$$\tau_w = \sum_{i=1}^k w_i \tau_i, \quad \sum w_i = 1$$

$$\begin{aligned}\gamma : (n, \tau_1 \dots \tau_k) &\rightarrow (n, \tau_w) \\ \mathcal{E} &\rightarrow \mathcal{E}_w \\ D &\rightarrow D_w \text{ (grey zone)} \\ n(\tau) &\rightarrow n(\tau)_w \text{ (dark zone)}\end{aligned}$$

- ♦ D_w – convex hull of the points (n, τ_w) :

$$n = 1 - \sum \lambda_i \varepsilon_i / \Lambda, \quad \tau_w = \sum w_i \varepsilon_i, \quad \varepsilon_i = \{0, 1\}$$

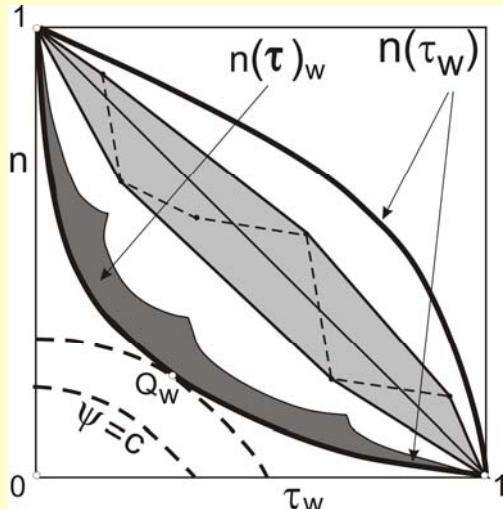
- ♦ $w_1 = \dots = w_k \Rightarrow$

lower boundary of D_w = convex minorant of the points:

$$(n, \tau_w) : (1, 0), (1 - \lambda_{(k)} / \Lambda, 1/k), \dots, (1 - \sum_{i=1}^p \lambda_{(k-i+1)} / \Lambda, p/k) \dots (0, 1)$$

$$\lambda_{(1)} \leq \lambda_{(2)} \dots \leq \lambda_{(k)}, \text{ ordered } \{\lambda_i\}$$

◆ D_w , grey zone $=\{n+\tau_w=1\} \leftrightarrow w_i=\lambda_i/\Lambda, i=1,\dots,k$



◆ $\{n_i(\tau)\}$ are piecewise smooth, $i=1,\dots,k$

$n(\tau_w)$, dark zone $= n(\tau_w) \leftrightarrow I(t)$ is trivial

◆ upper bound of $n(\tau_w)$:

$$n^+(x) = \max_{i,\varepsilon} \{\lambda_i/\Lambda \cdot n_i(x/w_i - a_i(\varepsilon)) + b_i(\varepsilon)\}$$

$$\varepsilon = (\varepsilon_1 \dots \varepsilon_k), \quad \varepsilon_i = (0,1)$$

$$a_i(\varepsilon) = \sum_{j \neq i} w_j \varepsilon_j / w_i, \quad b_i(\varepsilon) = \sum_{j \neq i} \lambda_j (1 - \varepsilon_j) / \Lambda$$

◆ optimal property of $n(\tau_w)$:

$$\varphi(n, \tau_1 \dots \tau_k) = \psi(n, \tau_w) \Rightarrow \min_{\pi} \varphi = \psi(Q), \quad Q \in n(\tau_w)$$

STATISTICAL PROBLEMS

1. Confidence zone for D_w of level α , $D_w(\alpha)$

PROBLEM: find $D_w(\alpha)$ using a confidence zone $U(p/\alpha)$
for unknown $\{p_i = \lambda_i/\Lambda\}$

SOLUTION

$$\diamond U(p|\alpha): \sum_{i=1}^k (p_i - \hat{p}_i)^2 / \hat{p}_i < q(\alpha)$$

$\{\hat{p}_i\}$, unbiased estimates of $\{p_i\}$, based on N events, $\hat{p}_i N \gg 1$

$$q(\alpha) \approx [(k-1) + \rho \sqrt{2(k-1)} / N]$$

$$\rho(\alpha) \approx 2.5 \text{ for } k \geq 5, \alpha > 97.5\%$$

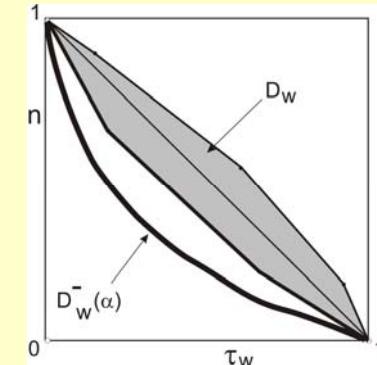
$\diamond D_w^-(\alpha)$ is the convex minorant of (n, τ_w)

$$\bullet \quad n(\varepsilon) = f\left(\sum_{i=1}^k \hat{p}_i \varepsilon_i\right), \quad \tau_w(\varepsilon) = \sum_{i=1}^k w_i \varepsilon_i, \quad \varepsilon_i = \begin{cases} 0 \\ 1 \end{cases}, \quad \#\{\varepsilon\} = 2^k$$

$$f(x) = \begin{cases} 1-x-\sqrt{qx(1-x)}, & (1+q)x < 1 \\ 0 & (1+q)x > 1 \end{cases}$$

$$\bullet \quad n(j) = f\left(\sum_{i=1}^j \hat{p}_{(k-i+1)}\right), \quad \tau_w(j) = j/k, \quad j = 1, \dots, k \quad \text{if } w_1 = \dots = w_k$$

$\hat{p}_{(1)} \leq \hat{p}_{(2)} \leq \dots \leq \hat{p}_{(k)}$, ordered $\{\hat{p}_i\}$



STATISTICAL PROBLEMS (continuation)

- ◆ Confidence α -zone for $\{n+\tau_\lambda=1\}$

$$w_i = \hat{p}_i \quad \Rightarrow \quad n(\varepsilon) = f(x), \quad \tau(\varepsilon) = x = \sum \hat{p}_i \varepsilon_i, \quad \varepsilon_i = \begin{cases} 0 \\ 1 \end{cases}$$

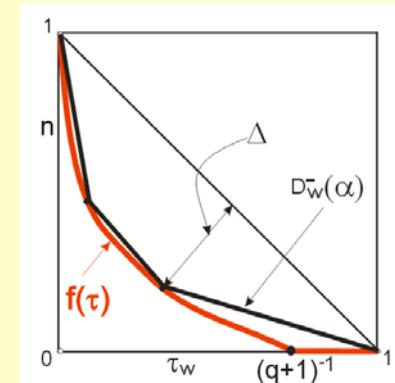
- $D_w^-(\alpha)$ given $\{w_i = \hat{p}_i\}$

$$D_w^-(\alpha) \geq f(\tau) = \begin{cases} 1 - \tau - \sqrt{q\tau(1-\tau)}, & (1+q)\tau < 1 \\ 0, & (1+q)\tau \geq 1 \end{cases}$$

- Zone width:

$$h = \Delta \sqrt{2} = \max_{\varepsilon} (1 - n(\varepsilon) - \tau_w(\varepsilon)) \leq \begin{cases} 0.5\sqrt{q}, & q \leq 1 \\ q/(1+q), & q \geq 1 \end{cases}$$

- $\alpha \approx 97.5\% \Rightarrow h \leq 0.5\sqrt{(k + 2.5\sqrt{2k})/N_-} = \begin{cases} h \leq 0.05 & k = 10, N_- = 2000 \\ h \leq 0.15 & k = 10, N_- = 200 \end{cases}$

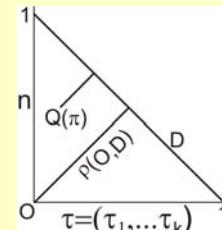


Prediction efficiency: $1-n-\tau$ (HK score)

- ◆ HK with $\tau = \tau_w$, $w_i \neq \lambda_i / \Lambda$ say nothing about the prediction potential of π

- ◆ correct version of HK score

$$\begin{aligned} HK(\pi) &= 1 - n - \sum \lambda_i \tau_i / \Lambda = 1 - n - \tau_\lambda \\ &= \rho(Q, D) / \rho(O, D) \end{aligned}$$



$$D = \{n + \sum \lambda_i \tau_i / \Lambda = 1\}$$

- ◆ $(1-n) - \tau_\lambda > 0$, the rate of not random successes

- ◆ Optimization of $HK(\pi)$

$$\pi(t | G_i) = \text{alarm in } G_i \times (t, t + \Delta) \text{ as soon as } r(t | G_i) > \lambda_i = \text{Er}(t | G_i)$$

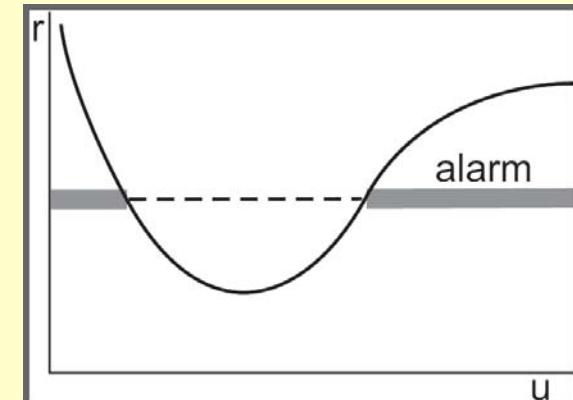
Typical HK-values

- **Renewal** model (Molchan, 1991)

$I(t)$ is the time since the last event

$F(x) = P(t_{k+1} - t_k < x)$
interevent time distribution

$$r(t) = mF'(u)/(1 - F(u)), \quad u = I(t)$$



- If F is Weibull, Log-Normal or Gamma distribution with mean m and standard deviation σ then

σ/m	.25	.50	.75
HK score	.52-.60	.32-.38	.15-.22

↑
San Andreas

- ETAS model with five day lag of the data (Helmstetter & Sornette, 2003)

HK score=20%, M=6;

HK score=30%, M=7

- M8 (Kossobokov, 2005)

$$\text{HK score} = 1 - n - \tilde{\tau} = \begin{cases} 50\%, M = 8 \\ 20\%, M = 7.5 \end{cases}$$

$\delta n \approx \frac{1}{N} \approx 10\%$ for a single future failure-to-predict

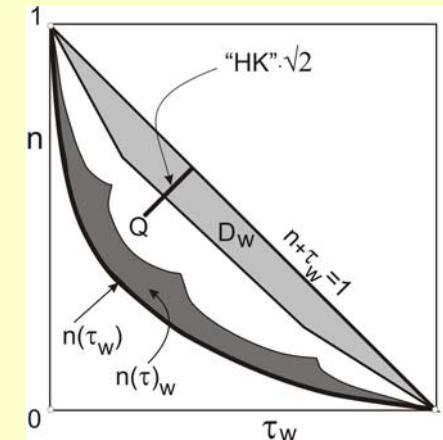
$\delta \tilde{\tau} \approx 5 - 10\%$ if $M=4, 5, 6$ or 7 is used to estimate the rate of $M=8$

Corollary:

- HK(Renewal) \approx HK(M8) \approx HK(ETAS)
- prediction of target events should not necessarily be based on a detailed model of seismicity

Conclusion

- be careful looking at (n, τ_w) diagram, because
 - ♦ $D_w = \{n + \tau_w = 1\}$ iff $w = (\lambda_1 / \Lambda, \dots, \lambda_k / \Lambda)$
 - ♦ $\{n(\tau_w)\} = \{n(\tau)\}_w$ iff $I(t)$ is trivial
 - ♦ $dist((n, \tau_w), n + \tau_w = 1)$ say nothing about the prediction potential of π if $\{w_i\} \neq \{\lambda_i / \Lambda\}$



- $HK(\pi) = 1 - n - \sum \tau_i \lambda_i / \Lambda$, but not " HK " = $1 - n - \tau_w$
is useful as a measure of prediction efficiency at the research stage
 $HK \approx 30\%$: M8 algorithm for $M=8$ (careful estimation)

Renewal models for San Andreas, $\sigma / m < 0.6$ (Molchan, 1997)

- the confidence area of D_w can be useful in statistical analysis of forecasting methods

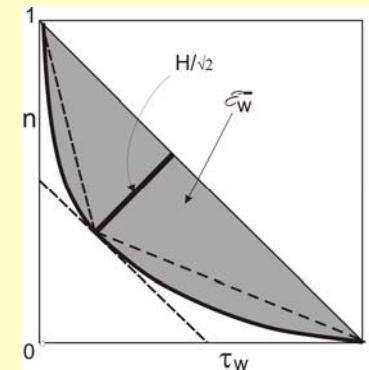
THANK YOU

- $H = \max_{\pi} HK(\pi) = \max_{\tau_\lambda} (1 - n(\tau_\lambda) - \tau_\lambda)$ or $A = 2 \cdot \text{area of } \mathcal{E}_w^-$?

$$|A - \hat{A}| < H(1-H)/2 < 1/8$$

$$\hat{A} = H(3-H)/2$$

→ H and the Area skill scores are equivalent

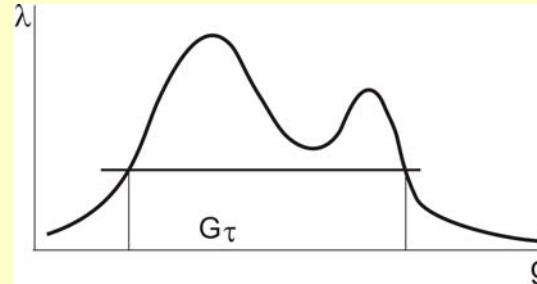


2. Distribution of $\hat{H} = \max(1 - \hat{n}(\tau_\lambda) - \tau_\lambda)$: Poissonian case

$\pi(t)$ = alarm in $G_\tau \times [0, T]$

$$\lambda(G_\tau) / \lambda(G) = \tau$$

$N_\tau = \#\{\text{success}\}$



$$\hat{H} = \max_{\tau} (N_\tau / N - \tau) \stackrel{\text{law}}{=} D_N^+$$

$N = \#\{\text{target event}\}$

D_N^+ = one sided Kolmogorov – Smirnov statistic

