

*Stanislav (Stas) Babak.*

*AstroParticule et Cosmologie, CNRS (Paris)*

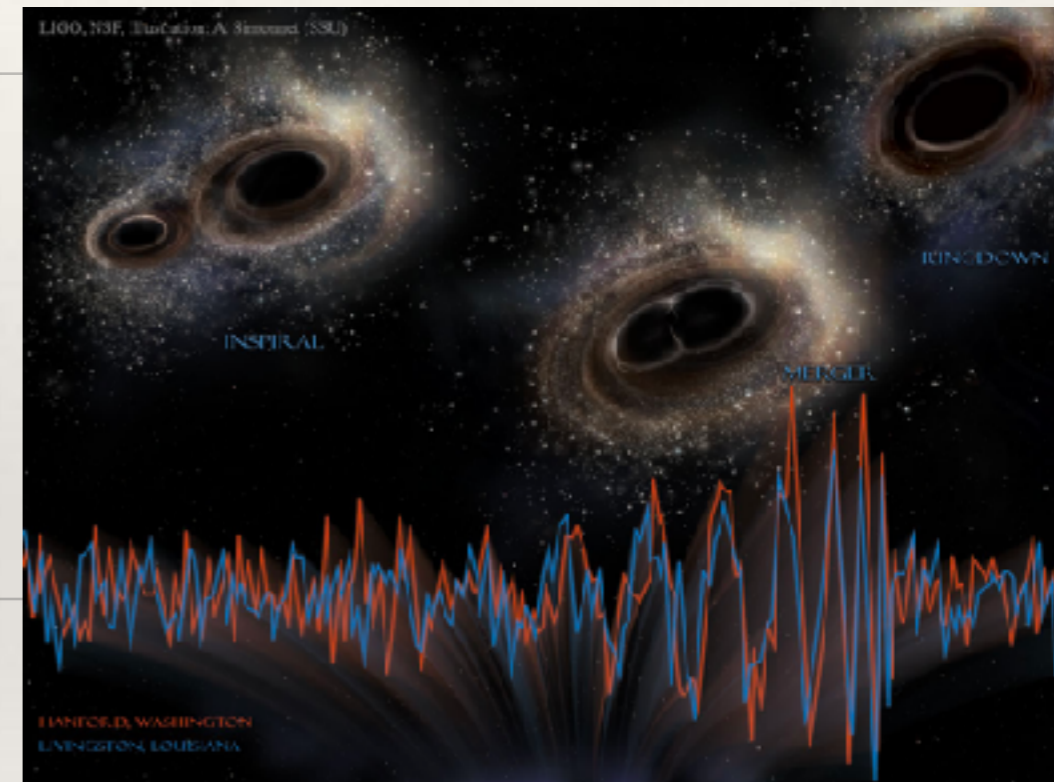


---

# Gravitational waves

## Part I

---



ICTP, 18-22 June 2018

# Lecture 4

---

- 📌 Data analysis: Part II
- 📌 Testing GR with GWs
- 📌 Pulsar Timing Array: detecting GWs in the nano-Hz band



# Likelihood

Let us assume that the data contains the signal.: hypothesis (model)  $H_1$

$$d(t) = n(t) + s(t, \theta_i) \quad \text{signal "s" depends on parameters } \theta_i$$

data = noise + signal

If the template matches the GW signal exactly  $h(t, \theta_i) = s(t, \theta_i) \longrightarrow d(t) - h(t, \theta_i) = n(t)$

$$p(d(t)|H_1, \vec{s}(t, \lambda)) = p(d(t) - s(t, \vec{\lambda})) = p_n$$

- Assume that the noise is Gaussian (but not necessarily white): non white noise has different variance at different frequencies. The the likelihood can be written as

**Likelihood:**  $p(d|H_1, \theta_i) \propto e^{-\frac{1}{2}(d-h(\theta_i)|d-h(\theta_i))}$

The inner product: matched filtering  $(a|b) \equiv 4\Re \int_0^\infty \frac{\tilde{a}(f)\tilde{b}^*(f)}{S_n(f)} df$

- We search for parameters which maximize the likelihood: making the residuals most noise-like — maximum likelihood estimators for parameters  $\hat{\theta}_i$



# Likelihood

○ For a given noise realization the maximum likelihood estimators.  $\hat{\theta} \neq \theta_{true}$   
 How close the estimated parameters to the true depends on the noise realization and on the strength of the signal: stronger the signal (high signal-to-noise ratio, SNR) closer estimators to the true values — less influence of the noise. Unbiased estimator: equal to true if averaged over the noise realizations.  $\langle \hat{\theta} \rangle = \theta_{true}$

○ If  $s \neq h(\theta_{true})$  — lack of accuracy in the signal modelling: systematic bias in parameter estimation.

$$\min_{\theta_i} (s - h(\theta_i) | s - h(\theta_i)) \rightarrow \tilde{\theta}_i \quad |\theta_{true} - \tilde{\theta}_i| = \delta\theta_i \quad \text{bias}$$

If there is a bias, we still can detect the GW signal on expense of making error in parameters characterizing the system (binary): *effectualness*

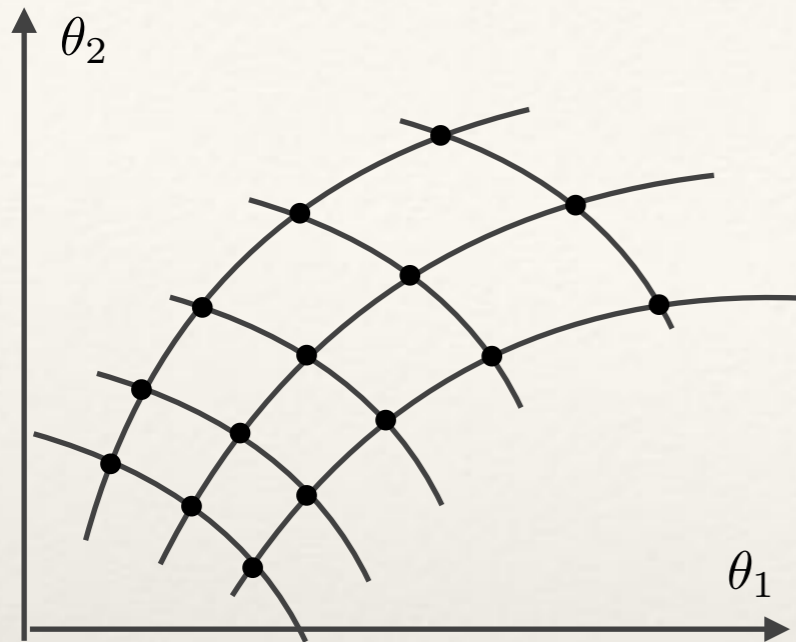
$$\text{Faithfulness} \quad \frac{(s|h(\theta_{true}))}{\sqrt{(s|s)(h(\theta_{true})|h(\theta_{true}))}} \quad \text{Overlap} \quad (\hat{s}|\hat{h}), \quad \hat{h} = \frac{h}{\sqrt{(h|h)}}$$

↓  
normalized

Overlap varies [-1, 1]: 1 is a perfect match, related to the loss in SNR



# Likelihood maximization

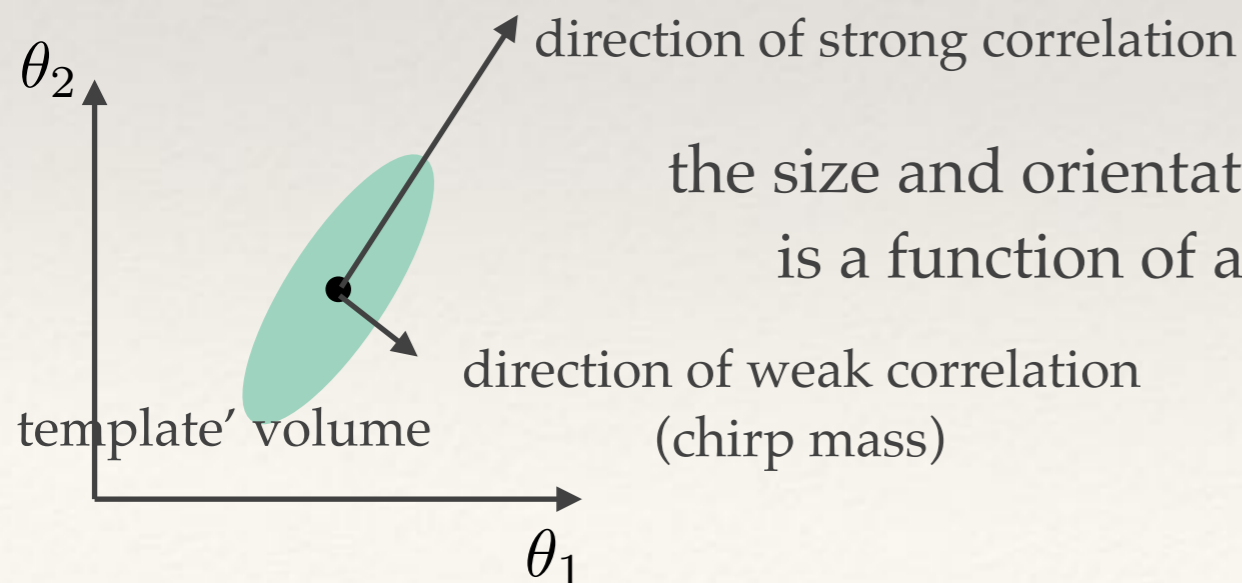


- We want to cover the parameter space (N-dim) by grid of points at equal distance from each other.
- Grid: not too coarse, not too fine
- The distance is determined **not** by a coordinate distance but by “proper” distance — correlation between nearby templates: introduce interval and metric

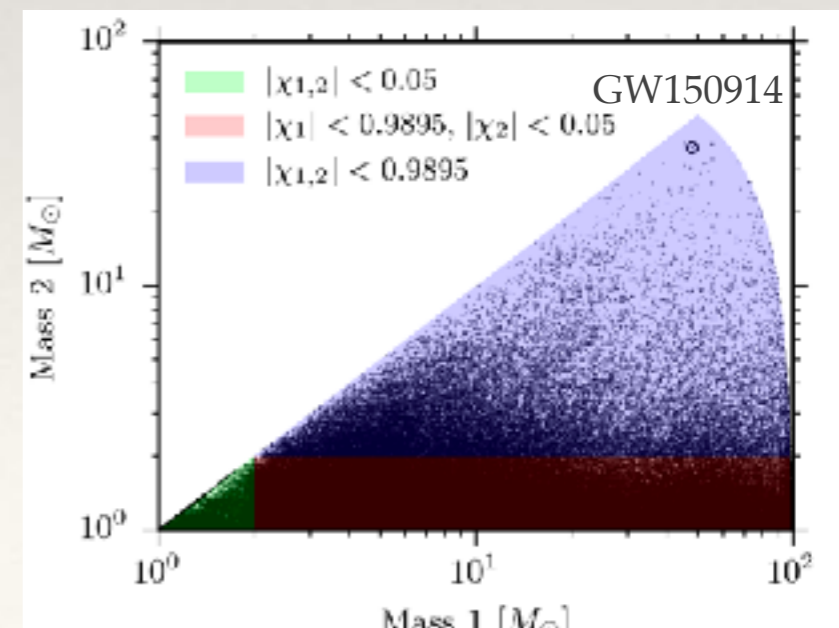
$$ds^2 = |\hat{h}(\theta_i + \delta\theta_i) - \hat{h}(\theta_i)| \approx (\hat{h}(\theta_i + \delta\theta_i) - \hat{h}(\theta_i)) \left( \frac{\partial \hat{h}}{\partial \theta_i} \mid \frac{\partial \hat{h}}{\partial \theta_j} \right) \delta\theta_i \delta\theta_j$$

Consider 2-D parameter space and fix  $ds = 0.01$

metric on the parameter manifold



the size and orientation of the ellipse is a function of a central point



# Bayesian approach

- Expensive computationally: often used when the signal is detected using the grid-based method (or something else) :  $H_1$  is true. Allows to test several models (different signal' models, non-GR theories).
- We have to assign the *prior* probability to our models and parameters of each model. We treat parameters describing a signal as random variables and trying to estimate probability distribution function(s) for each parameter based on the observed data (*posterior*)
- Consider several models  $M_i$  each parametrized by set of parameters  $\vec{\theta}_i$

Bayes' theorem 
$$P(M_i|d) = \frac{P(d|M_i)\pi(M_i)}{p(d)}$$

For a given model  $M_i$

$$\underbrace{P(\vec{\theta}_i|M_i, d)}_{\text{posterior}} = \frac{\underbrace{P(d|\vec{\theta}_i, M_i)}_{\text{likelihood}} \underbrace{\pi(\vec{\theta}_i)}_{\text{prior}}}{\underbrace{p(d|M_i)}_{\text{Evidence of model } M_i}}$$



# Bayesian approach

$$p(d|M_i) = \int d\vec{\theta}_i p(d|\vec{\theta}_i, M_i)\pi(\vec{\theta}_i) \quad \text{— Evidence of model } M_i: \text{ important for the model selection}$$

$$P(M_i|d) = \left[ \int d\vec{\theta}_i p(d|\vec{\theta}_i, M_i)\pi(\vec{\theta}_i) \right] \frac{\pi(M_i)}{P(d)} \quad \text{— probability of the model } M_i \text{ given a data}$$

○ **Odds ratio:** The problem to evaluate the normalization  $P(d)$  - requires full set of models which are mutually exclusive. We can evaluate the ratio of probabilities:

$$O_{a,b} = \frac{p(M_a|d)}{p(M_b|d)} = \underbrace{\frac{p(d|M_a)}{p(d|M_b)}}_{\text{Bayes factor}} \underbrace{\frac{\pi(M_a)}{\pi(M_b)}}_{\text{prior odds}}$$



# Markov Chain Monte Carlo (MCMC)

- So for a given model  $M_i$  we need to evaluate posterior pdf for all parameters and the evidence: posterior pdf tells us about parameters of the GW signal (system) and evidence tells us how good this models fits the observations.

$$p(\vec{\theta}_i | M_i, d) = \frac{p(d | \vec{\theta}_i, M_i) \pi(\vec{\theta}_i)}{p(d | M_i)} \quad \text{— posterior pdf}$$

$$p(d | M_i) = \int d\vec{\theta}_i p(d | \vec{\theta}_i, M_i) \pi(\vec{\theta}_i) \quad \text{— evidence of model } M_i$$

- Markov Chain Monte Carlo (MCMC)* approach:

We construct Markov chain: stochastic process where the next point in the chain depends only on the previous one. And:

- we want chain to move towards the region of parameter space with high likelihood
- we need to introduce the transitional probability: way to move from one point to another
- we want a transitional probability to satisfy the ballance equation

$$\text{ballance eqn.:} \quad P(\vec{\theta}_{(k)}) P(\vec{\theta}_{k+1} | \vec{\theta}_{(k)}) = P(\vec{\theta}_{(k+1)}) P(\vec{\theta}_k | \vec{\theta}_{(k+1)})$$

distribution we want to sample  
(posterior)

transitional probability





# MCMC

- Consider a particular implementation: Metropolis-Hastings
  - particular way of building transitional probability which satisfies the balance equation
  - we start with introducing a proposal distribution (arbitrary\*)  $q(\vec{\theta}_{(k+1)}|\vec{\theta}_{(k)})$
  - then we build the chain by introducing the acceptance probability

$$\alpha(\vec{\theta}_{(k+1)}|\vec{\theta}_{(k)}) = \min \left\{ 1, \frac{p(d|\vec{\theta}_{(k+1)})}{p(d|\vec{\theta}_{(k)})} \frac{q(\vec{\theta}_{(k)}|\vec{\theta}_{(k+1)})}{q(\vec{\theta}_{(k+1)}|\vec{\theta}_{(k)})} \frac{\pi(\vec{\theta}_{(k+1)})}{\pi(\vec{\theta}_{(k)})} \right\}$$

likelihood ratio

ratio of  
proposals

ratio of priors

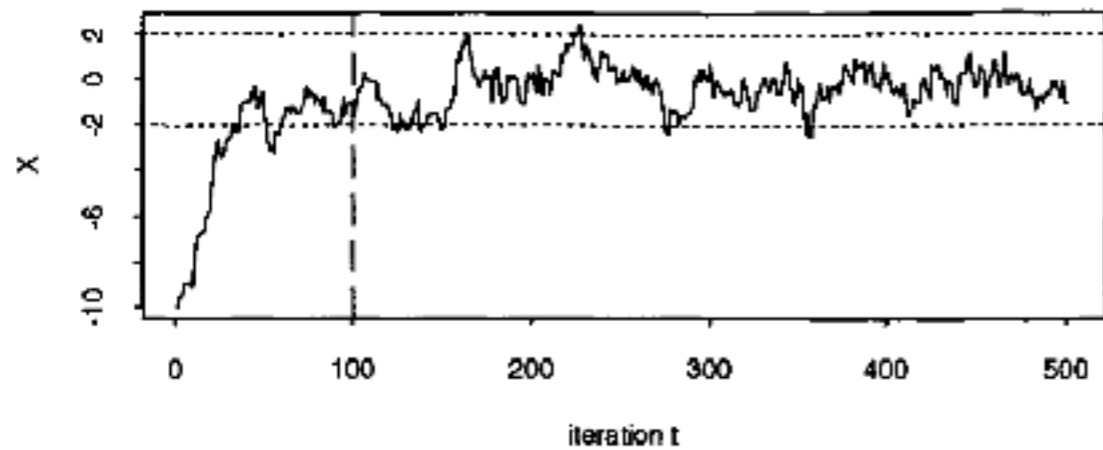
It is easier to understand if we use symmetric proposal and uniform priors: likelihood ratio

$\alpha$  — probability of accepting new point  $\vec{\theta}_{(k+1)}$

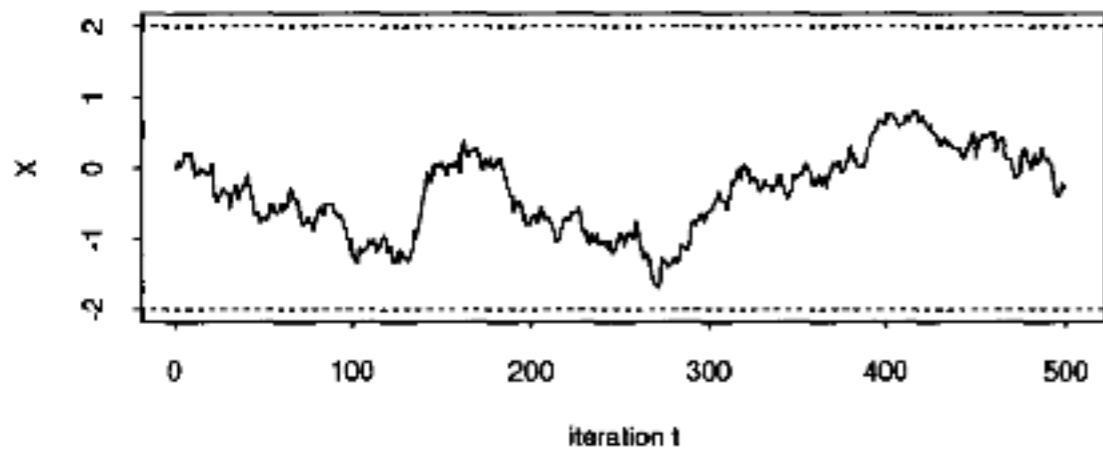
The chain moves predominantly in the direction of high likelihood.



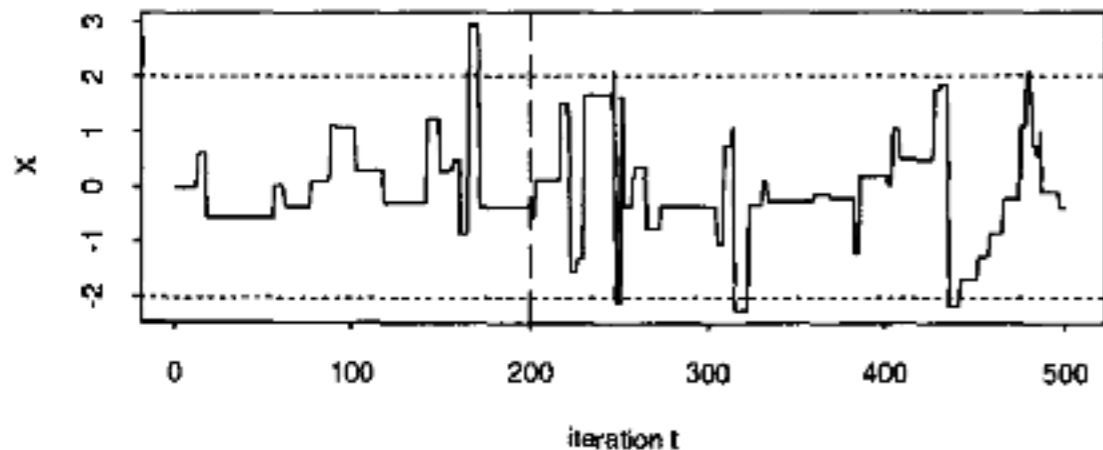
# MCMC



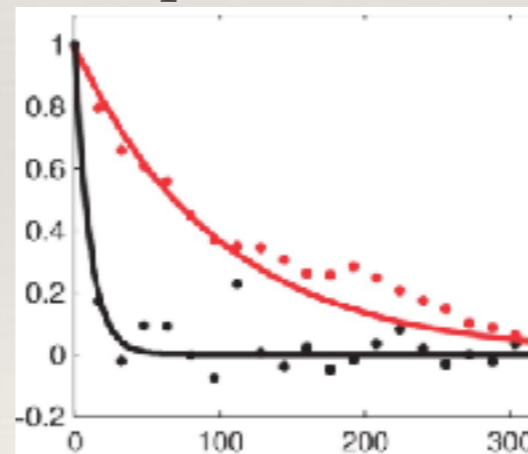
b



c



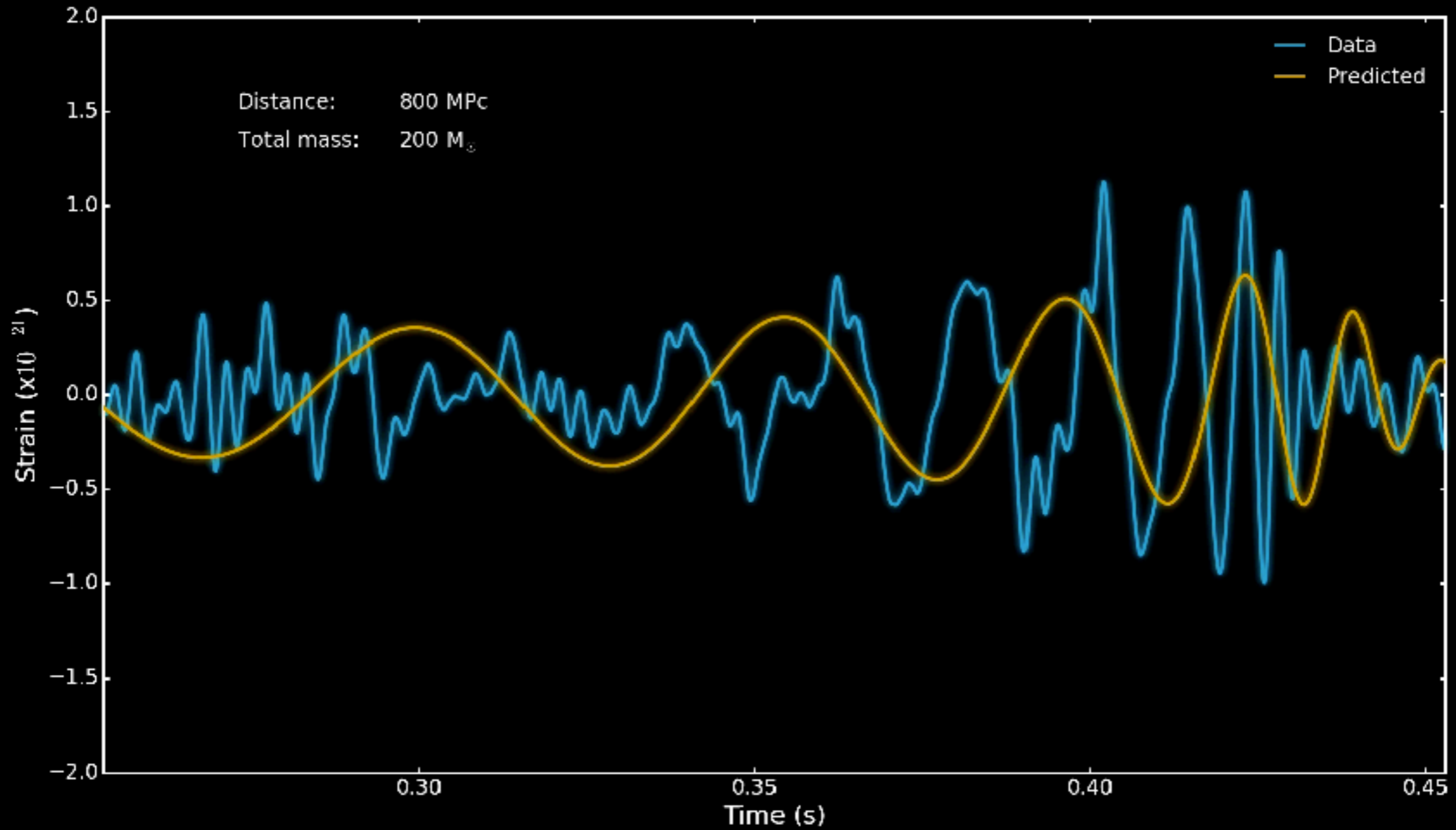
- The theorem tells us that the chains will sample the posterior pdf (after some burn-in length) independent of the proposal distribution, BUT
- The efficiency of the sampling strongly depends on the proposal (proposal should resemble the posterior)
- Number of samples vs. number of independent samples (defined by autocorrelation length)



- Multimodal posterior require special treatment! (simulated annealing, parallel tempering)



# GW data analysis



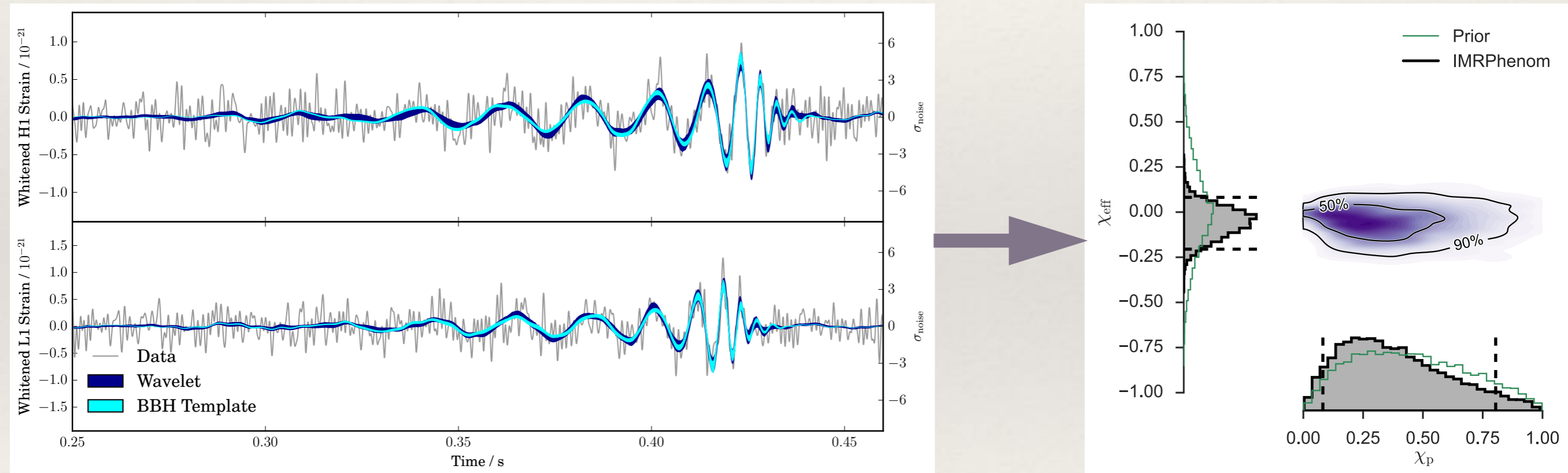
Data & Best-fit Waveform: LIGO Open Science Center ([losc.ligo.org](http://losc.ligo.org)); Prediction & Animation: C.North/M.Hannam (Cardiff University)



# GW data analysis

$$p(\theta|d) = \frac{p(d|\theta) p(\theta)}{p(d)}$$

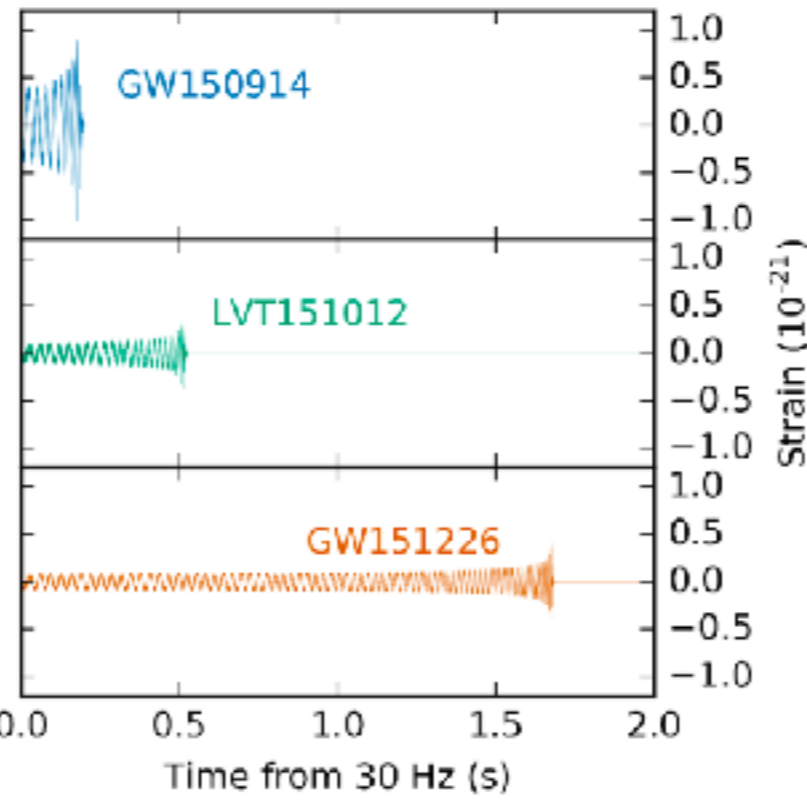
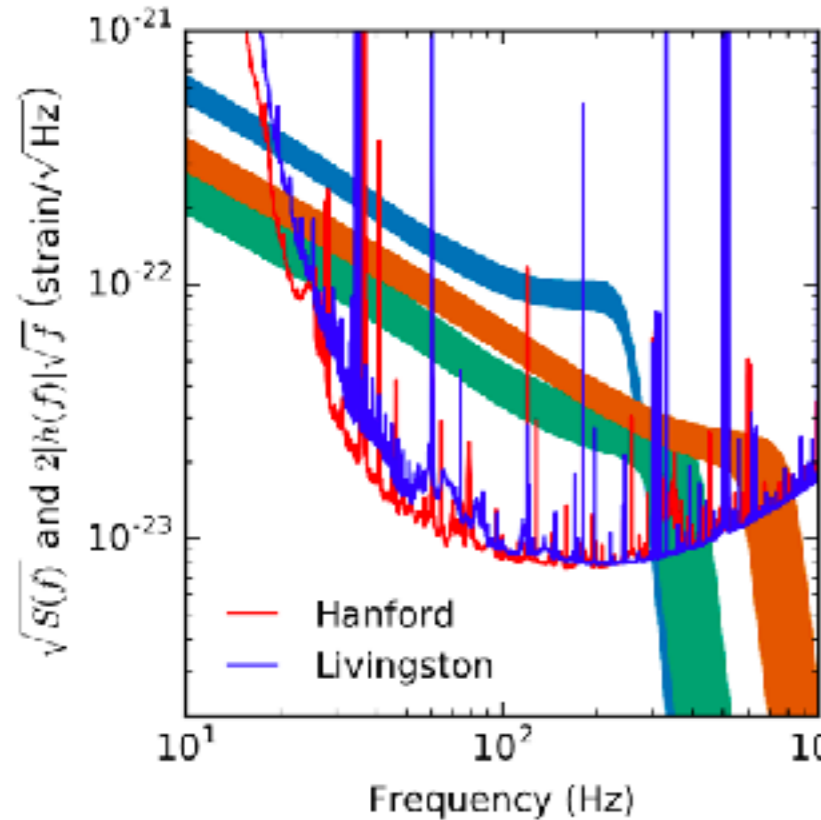
Fit to the data not a single line but multiple  
(for each parameter set in the posterior)



[GW150914, LSC+VIRGO PRL (2016)]



# GW data analysis



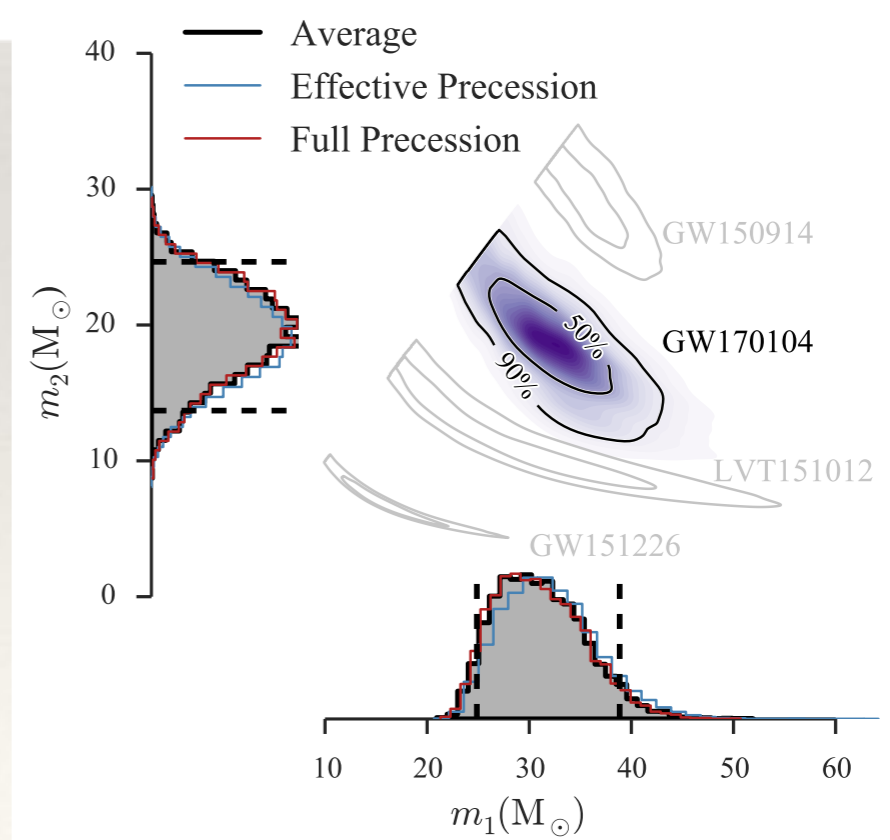
Chirp mass is the best measured parameter for the signal dominated by inspiral

$$\mathcal{M} = \frac{(m_1 m_2)^{3/5}}{(m_1 + m_2)^{1/5}}$$

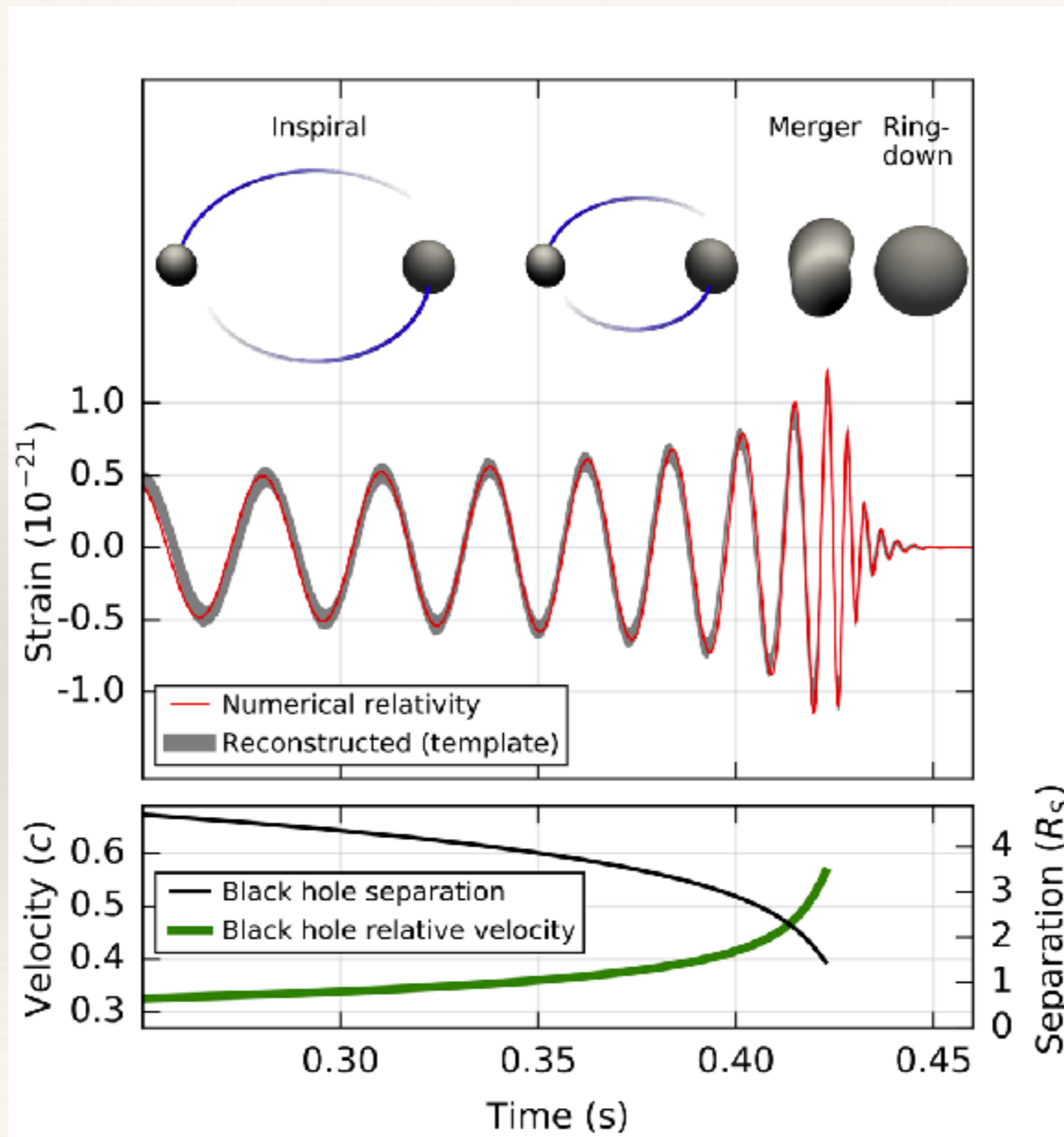
Frequency domain signal

Time domain signal

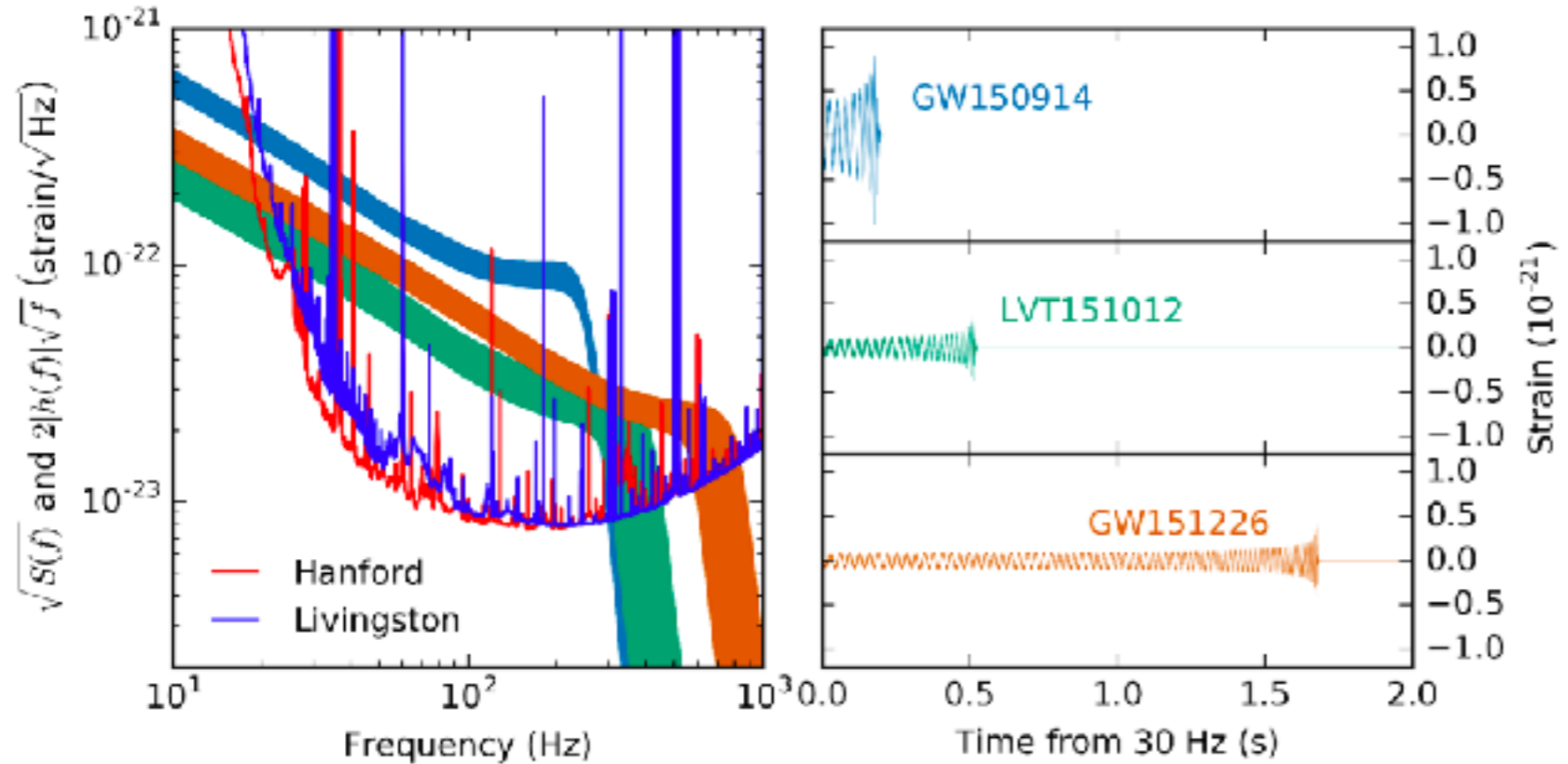
Estimation of BHs masses



# GW data analysis



# GW data analysis

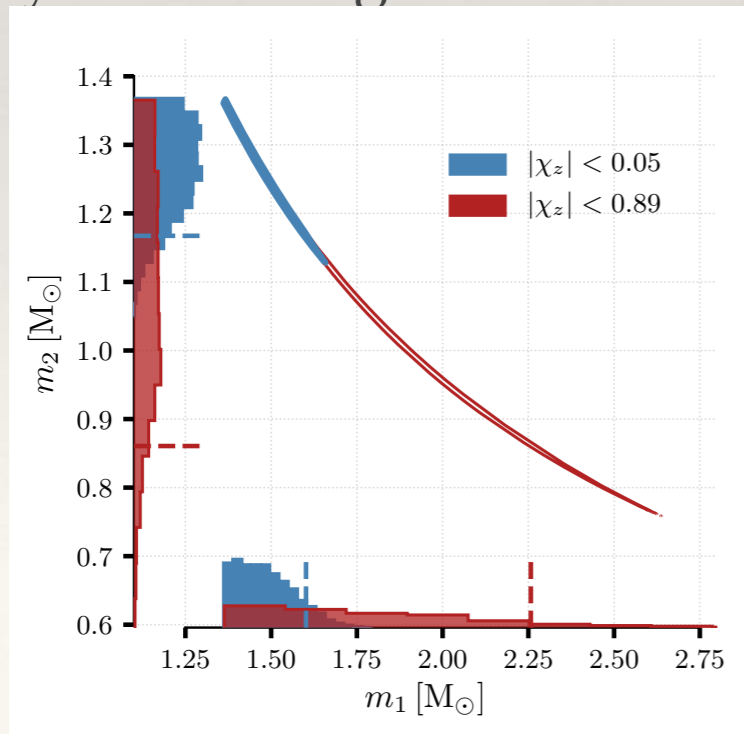


Chirp mass is the best measured parameter for the signal dominated by inspiral

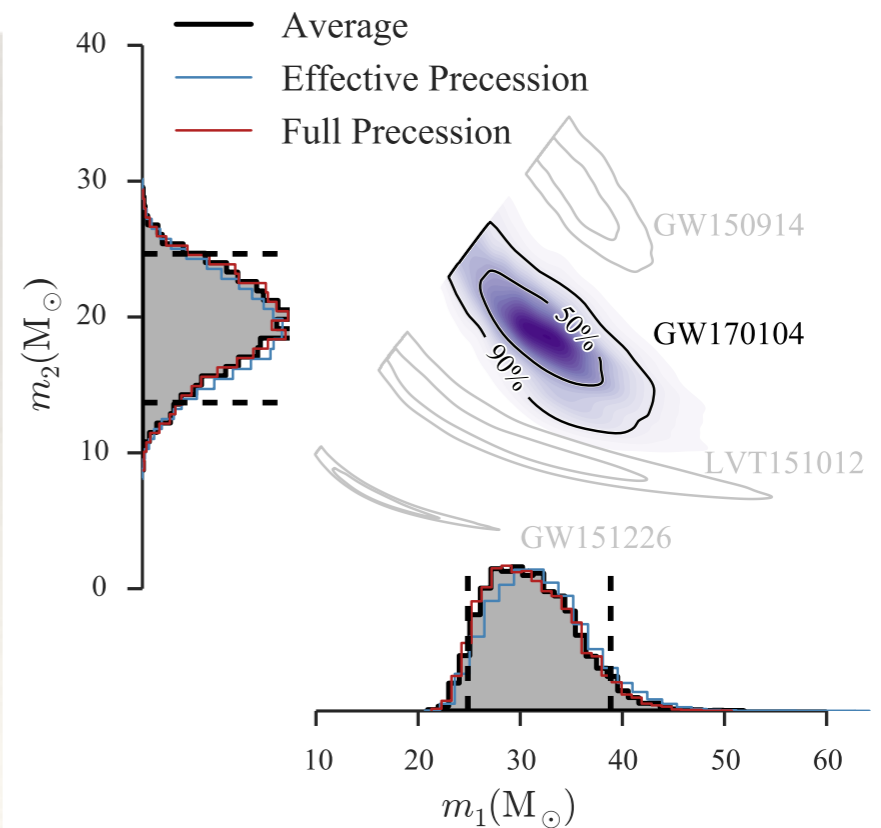
$$\mathcal{M} = \frac{(m_1 m_2)^{3/5}}{(m_1 + m_2)^{1/5}}$$

Frequency domain signal

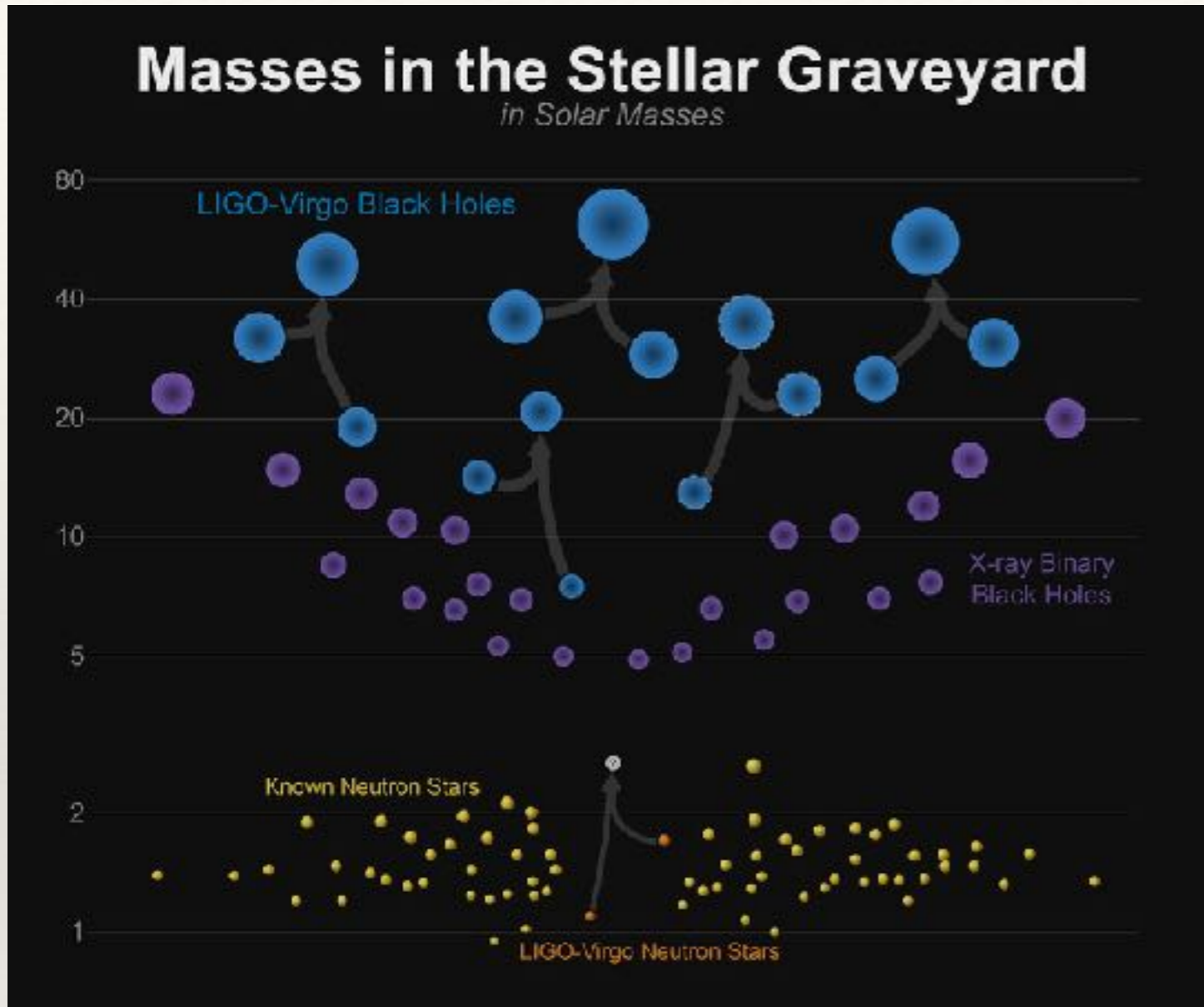
Time domain signal



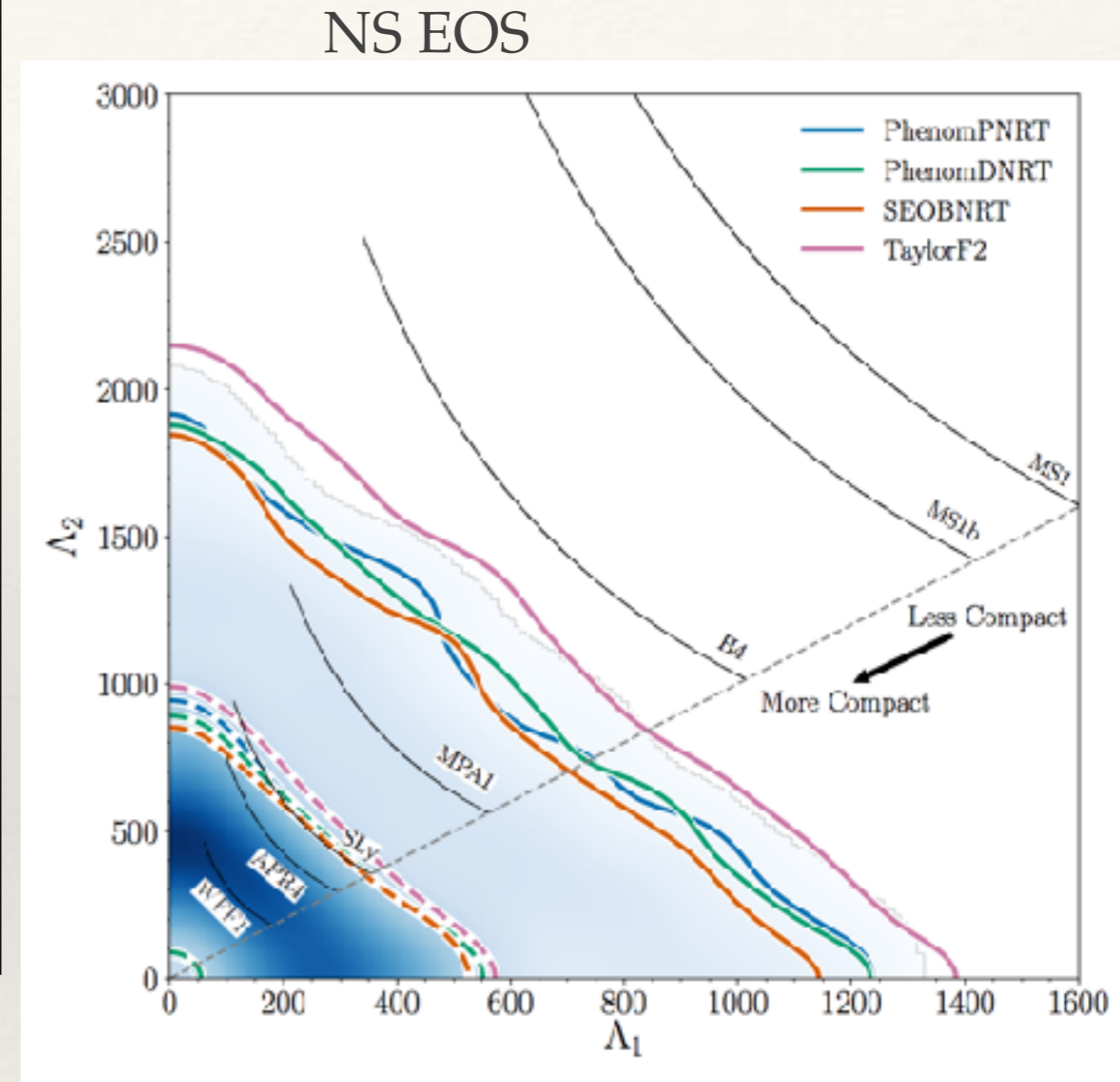
Estimation of BHs masses



# Detected GWs

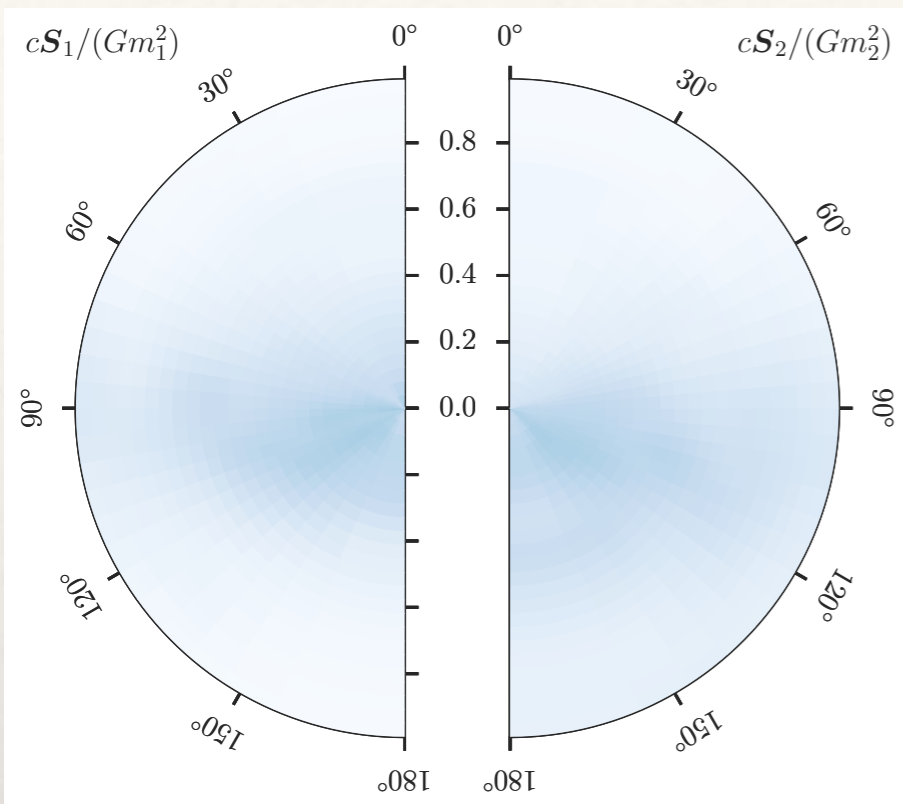


Mass distribution (LIGO)

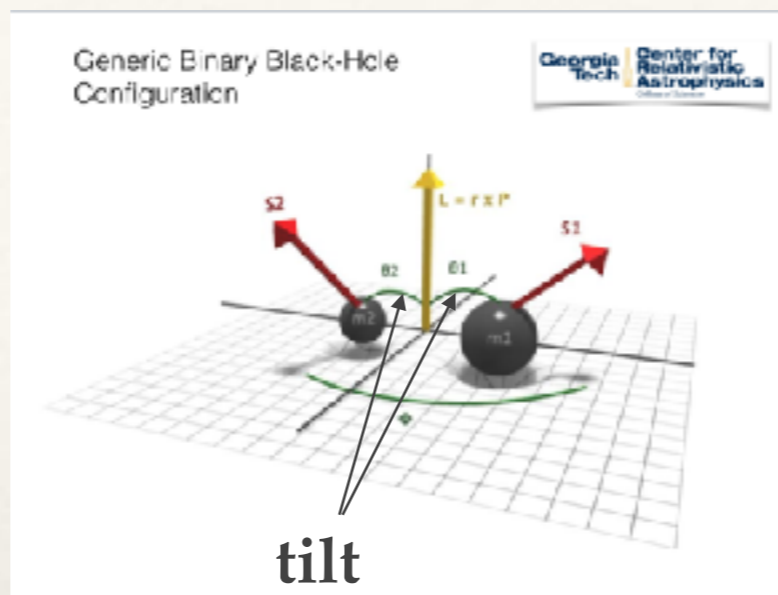




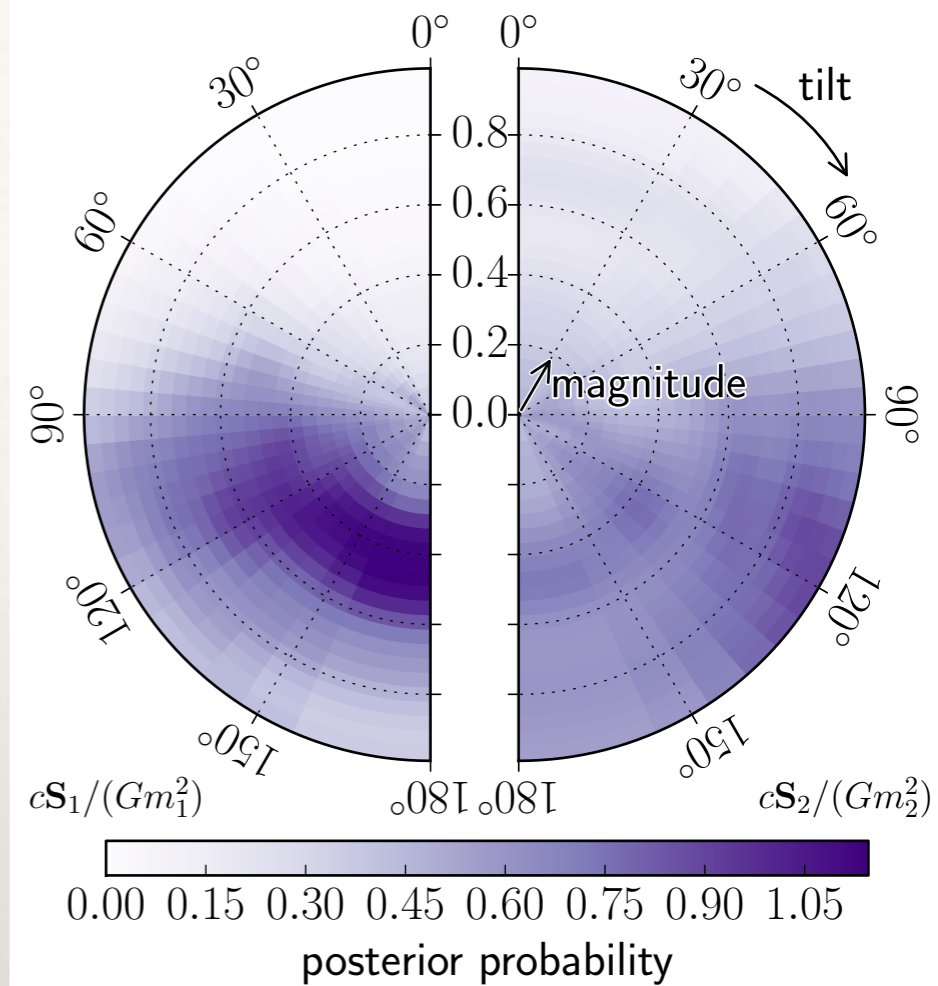
# Black Hole spins



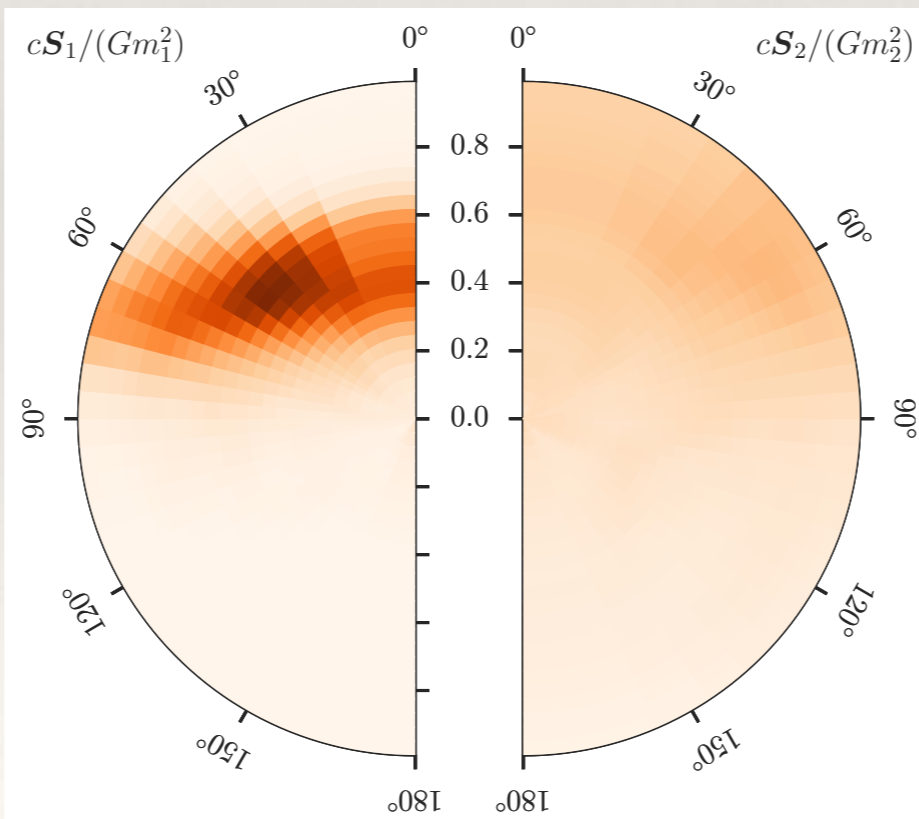
GW150914



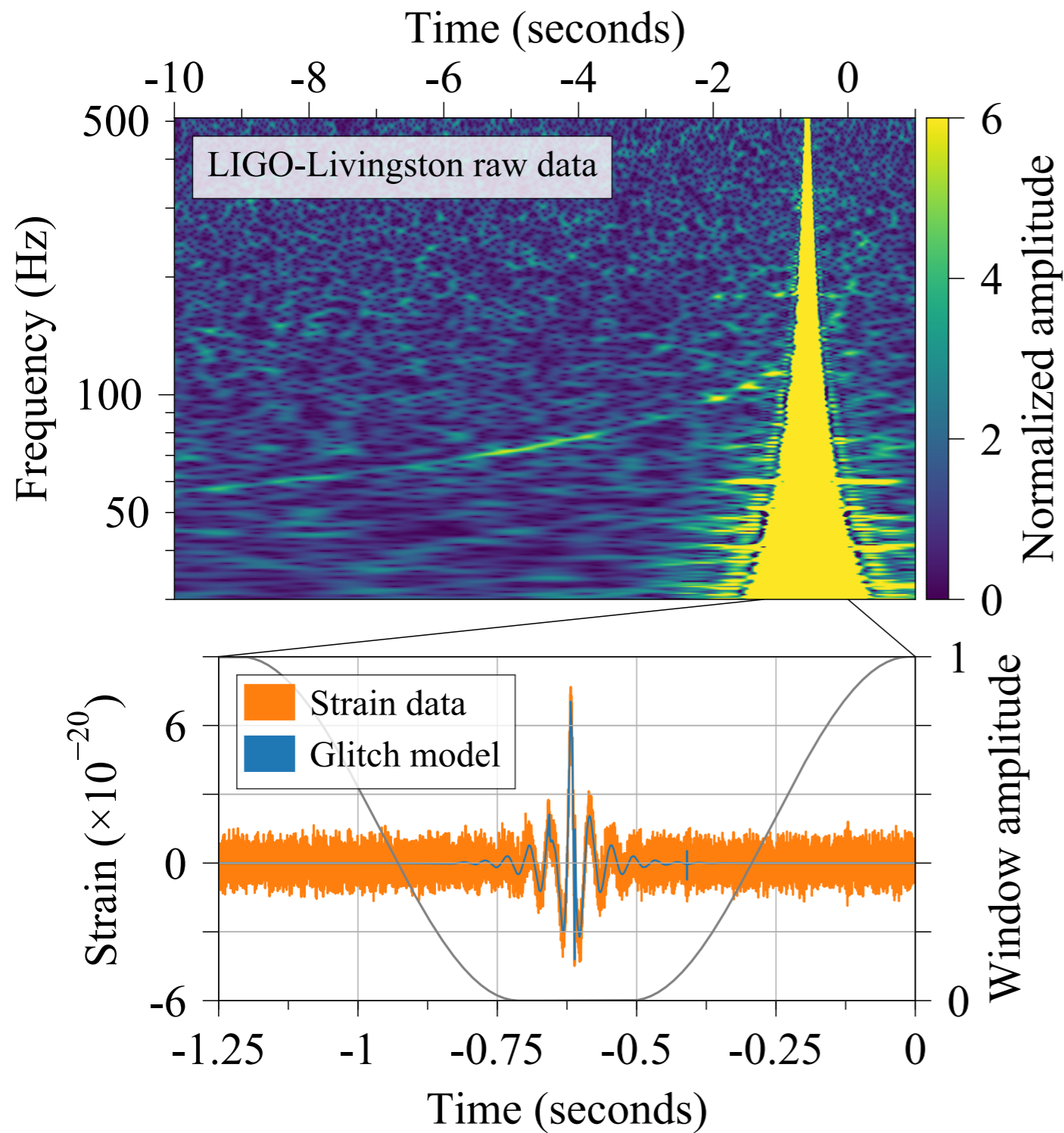
GW151226



GW170104



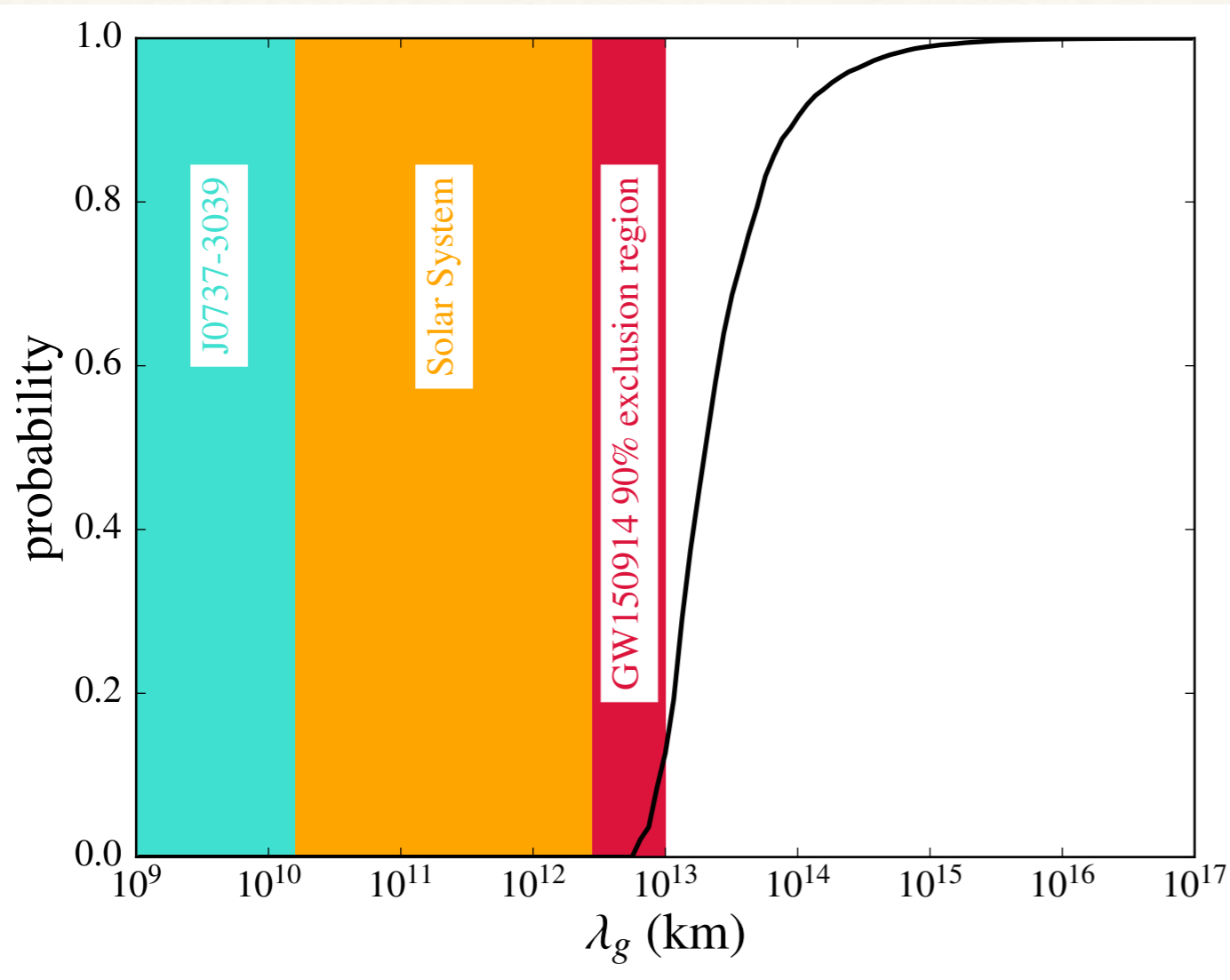
# Again non-stationarity



Strong glitch during the GW signal from BNS system



# Testing GR



- In GR graviton is massless (propagates with the speed of light)
- If graviton has non-zero mass ( $m_g$ ) then we should observe dispersion of GWs
- The observed data does not contradict GR: we can set upper limit on the mass of graviton (lower bound on the Compton wavelength)

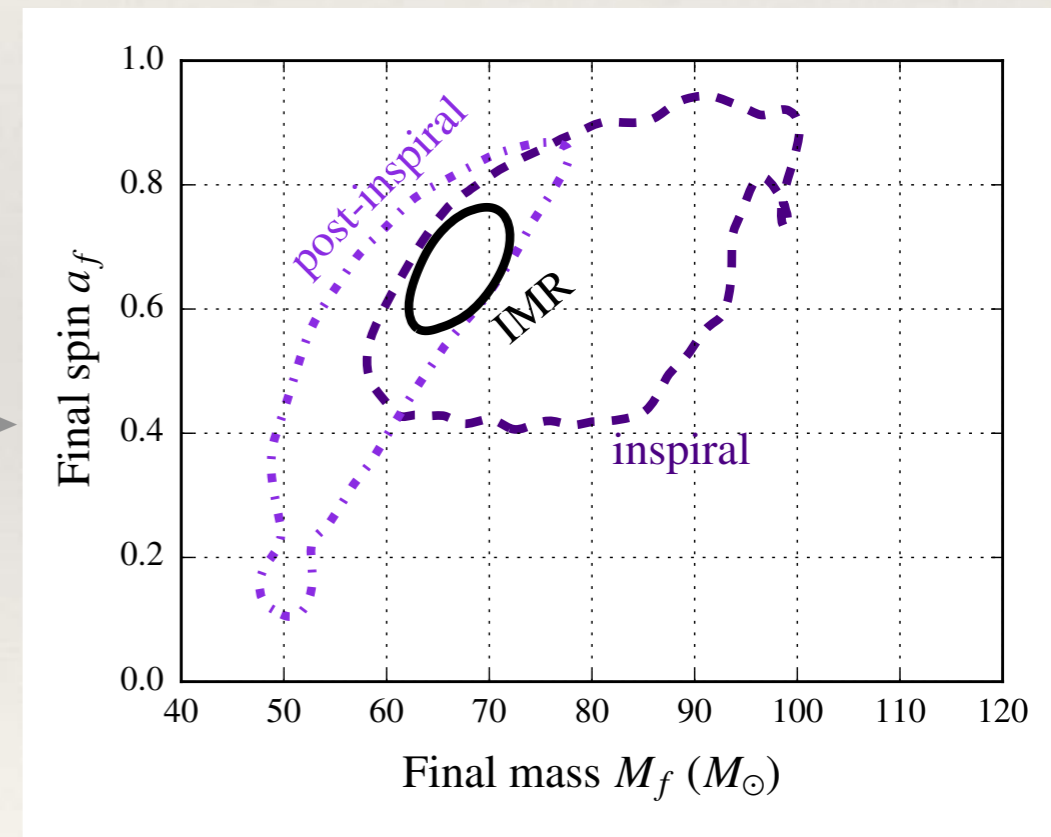
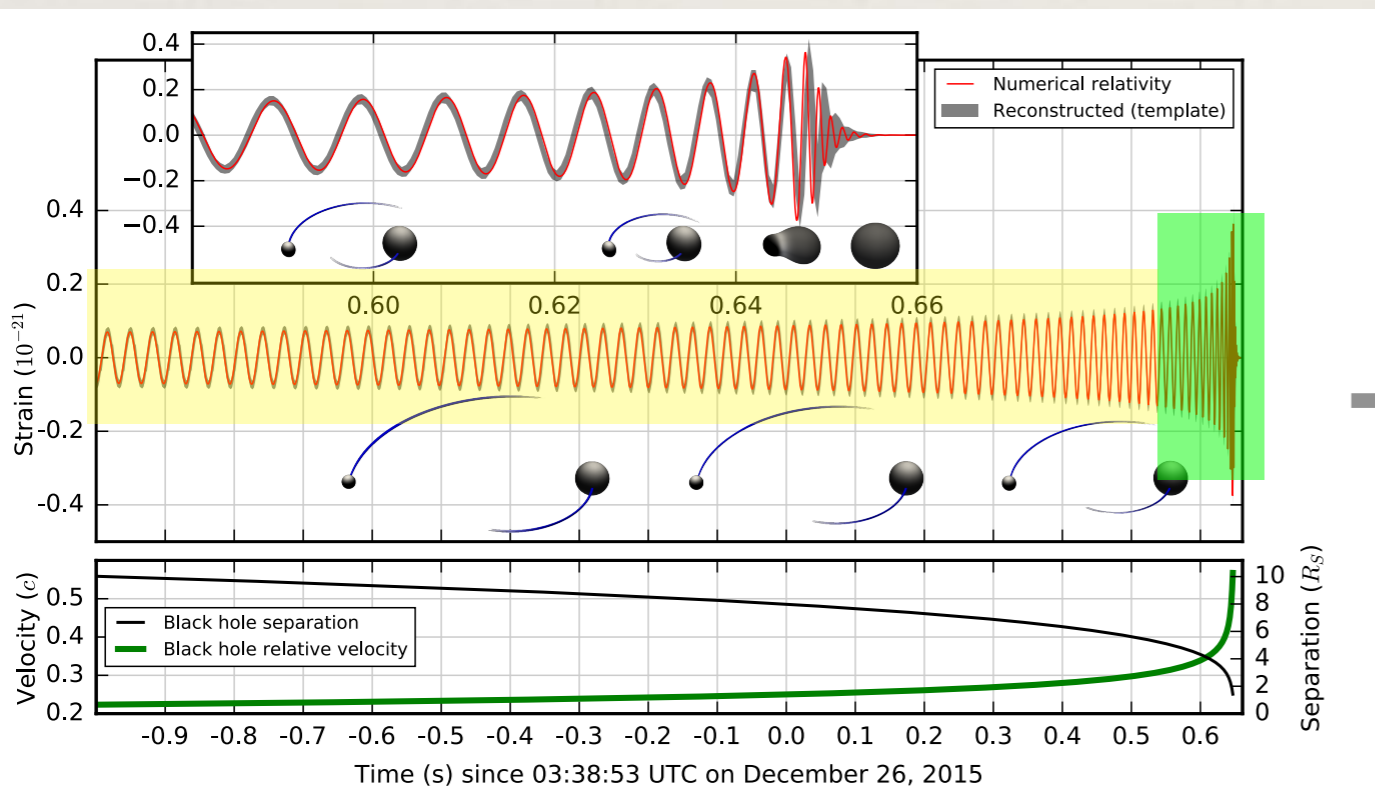
$$\lambda_g = h/(cm_g)$$

[LVC PRL (2016)]



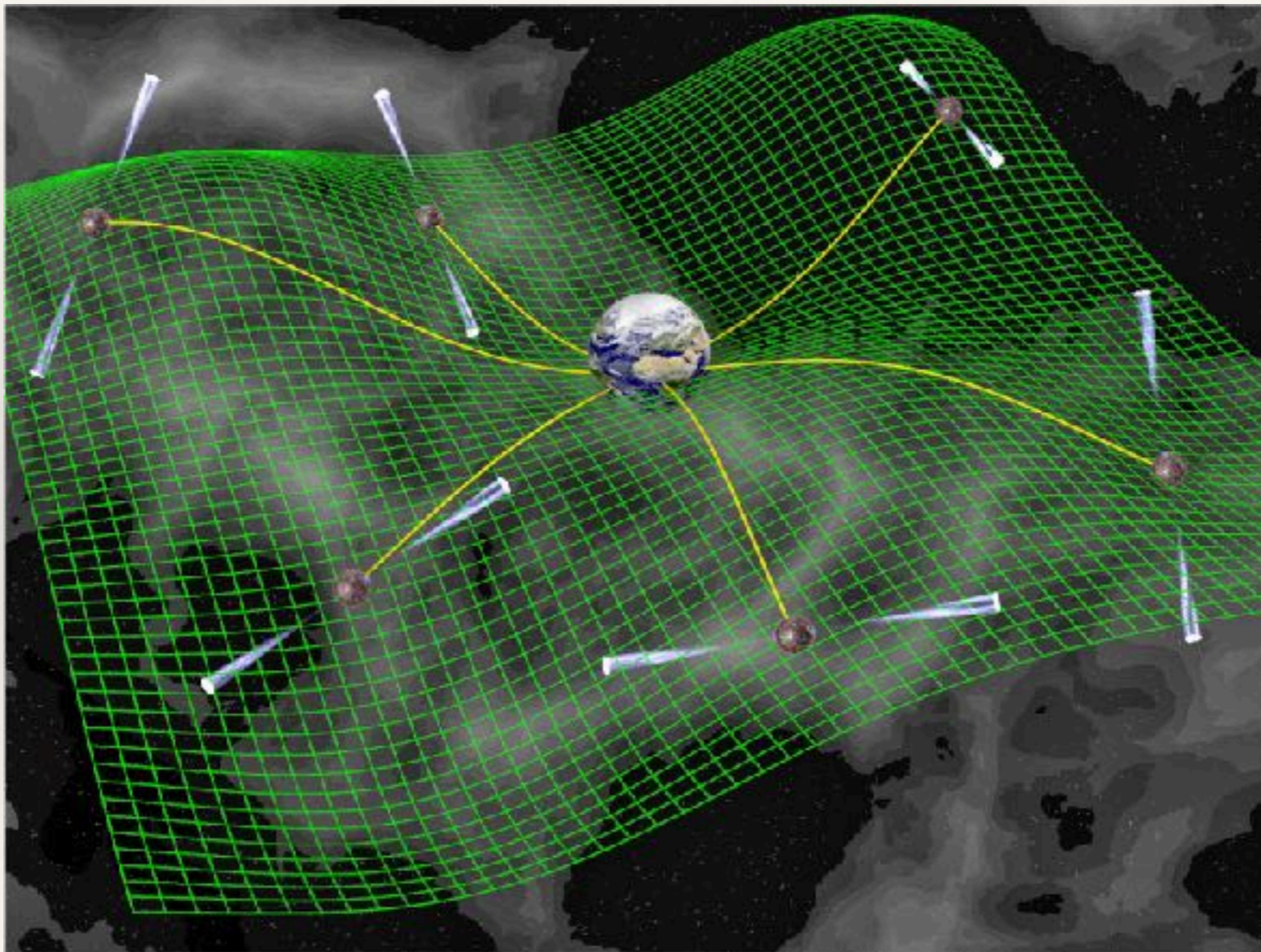
# Testing GR

- We can test self-consistency of GR
- We split the signal in two parts and analyze both parts separately: estimate parameters
- We compare if parameter estimation from each part are consistent



# Pulsar Timing Array

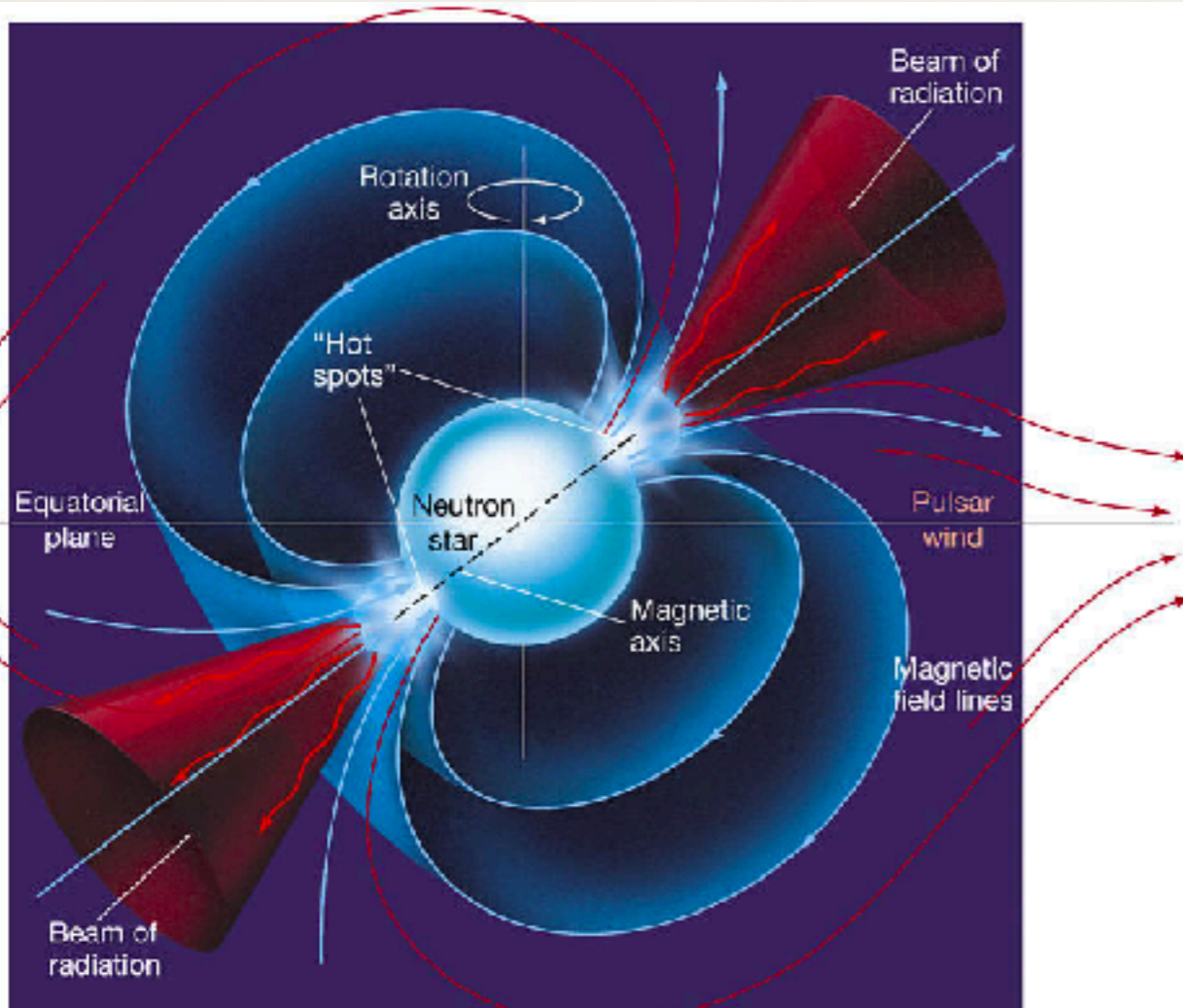
The main idea behind pulsar timing array (PTA) is to use ultra-stable millisecond pulsars as beacons (clocks sending signals) for detecting GW in the nano-Hz range ( $10^{-9}$  -  $10^{-7}$  Hz).



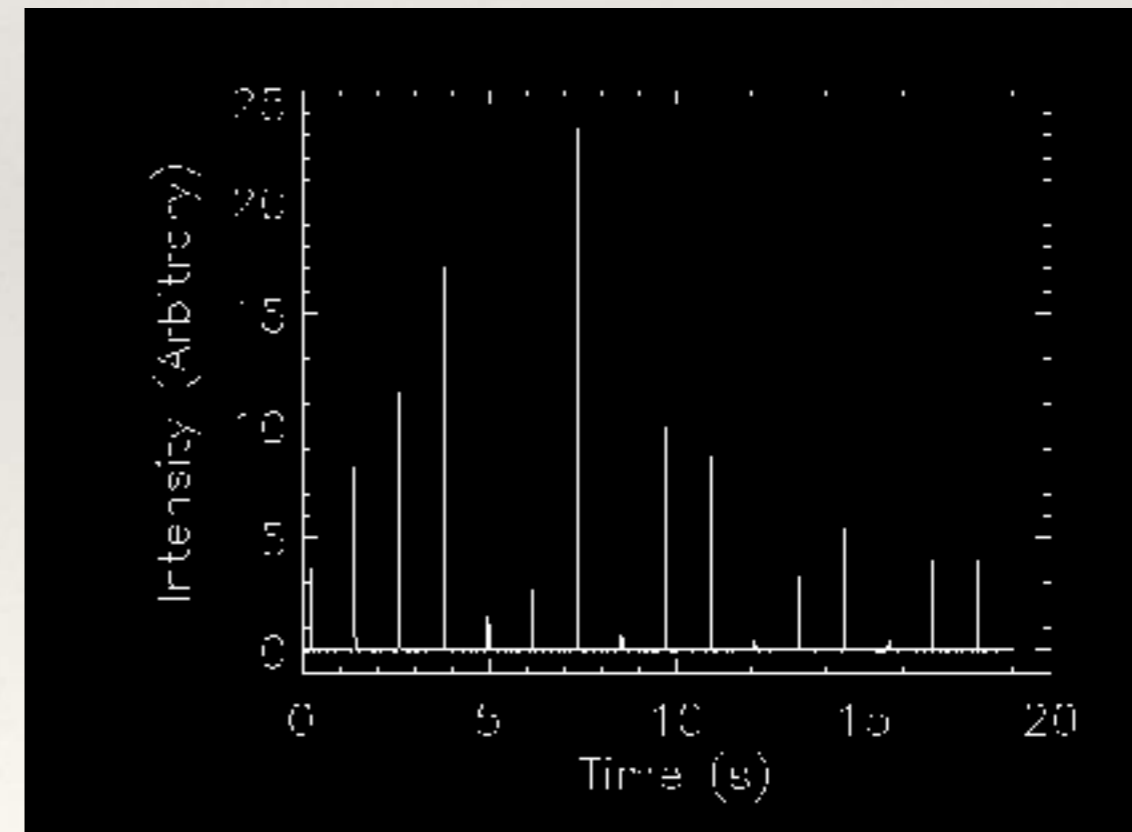
[Credits: D. Champion]

# Millisecond pulsars

- Pulsars - neutron stars (end product of evolution of stars with the mass  $> 7$  solar) with rapid rotation and strong magnetic field
- Emit beamed e/m radiation from the magnetic poles. Powered by rotation: spinning down.
- Beamed radio emission swaps across the line of sight — seen as pulses in observations (similar to the lighthouse)



Copyright © 2005 Pearson Prentice Hall, Inc.

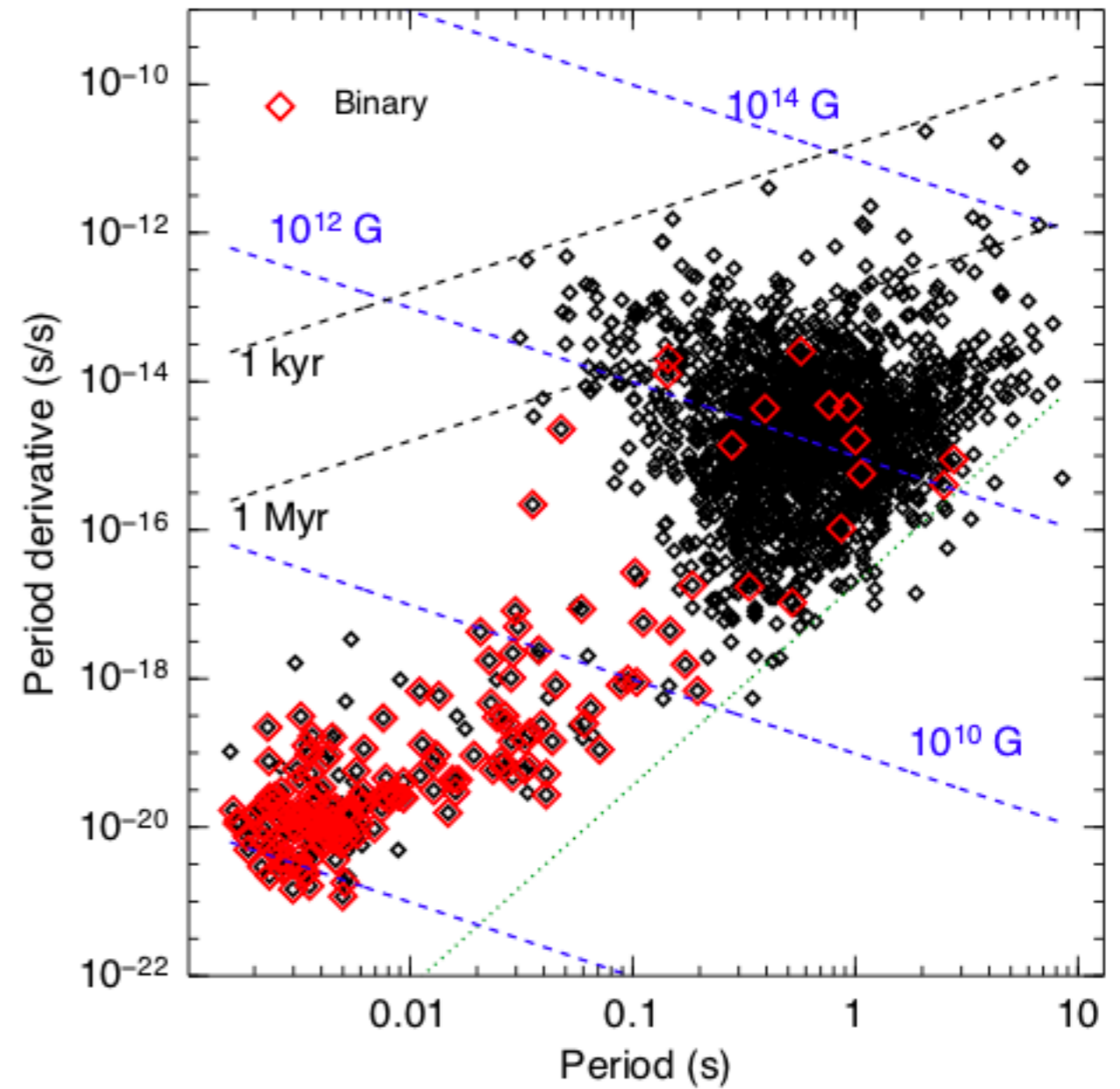


# Millisecond pulsars

- Millisecond pulsars: period of rotation  $\sim$  millisecc
- Often in binaries
- Very old NSs, very stable rotation
- The most accurate clock on the long time scale (decades)



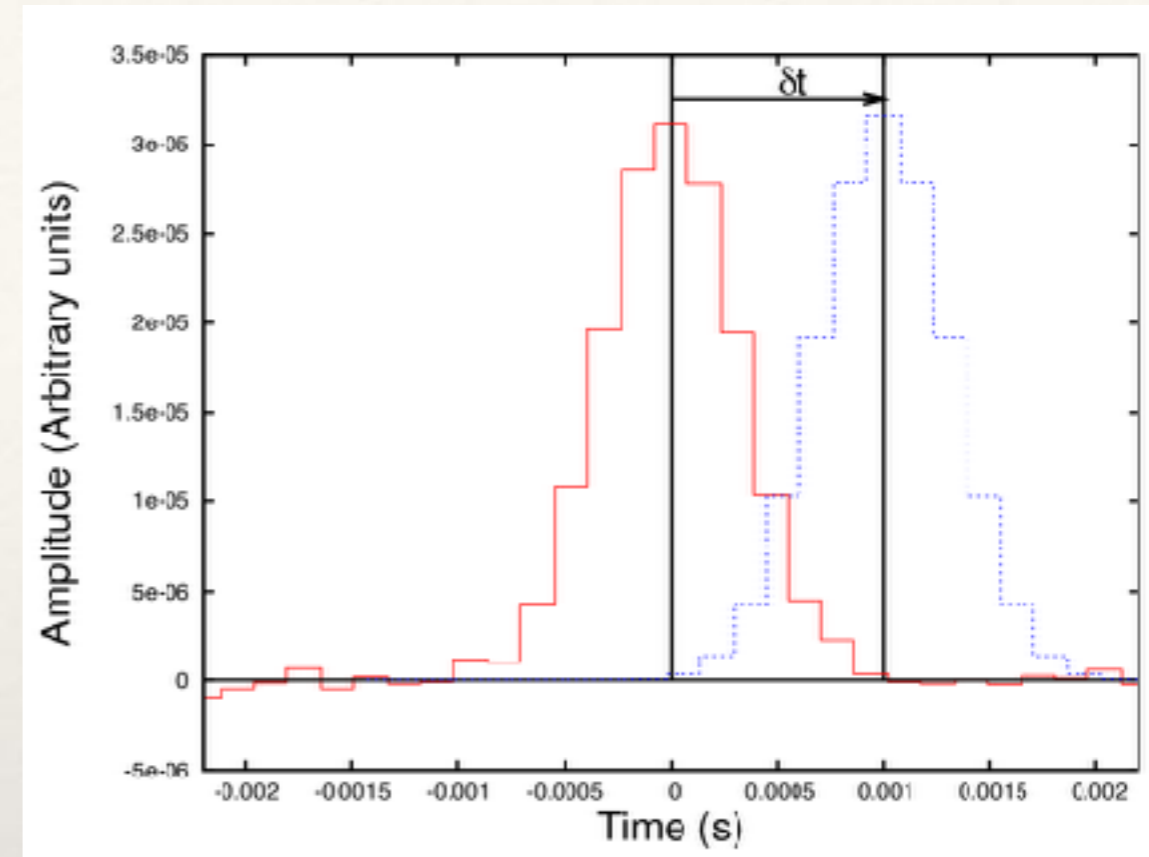
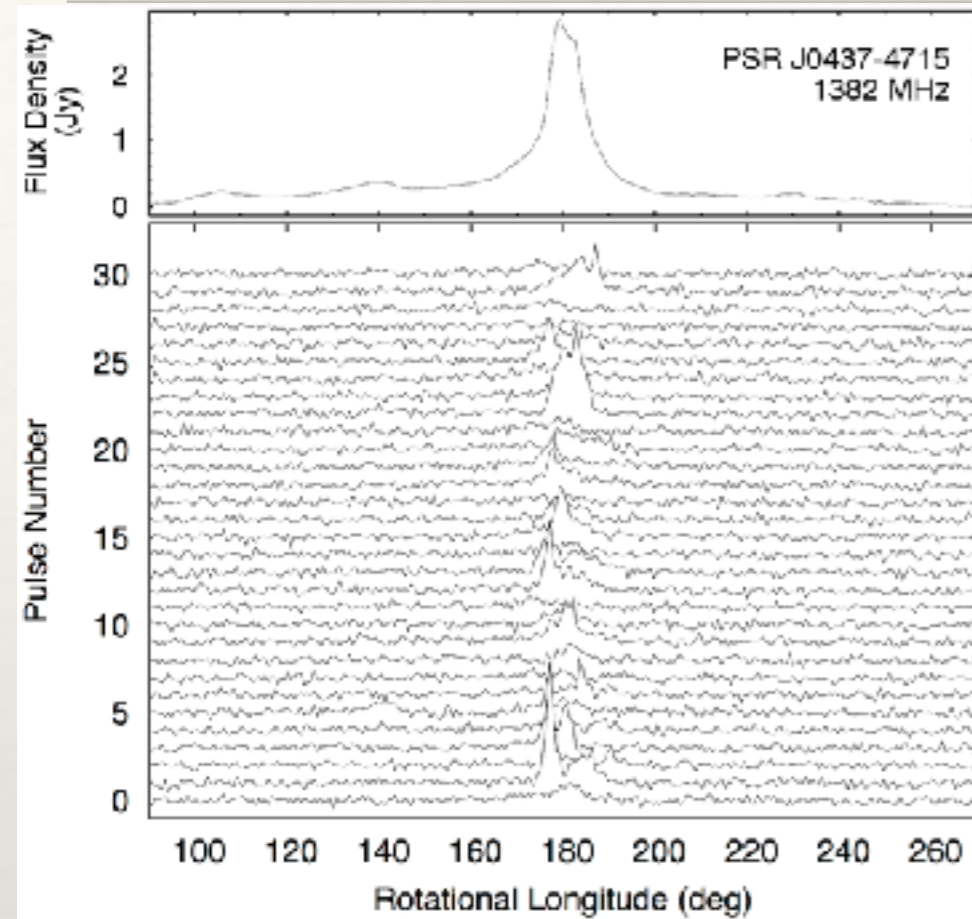
[Credits: NASA]



[Wikimedia]



# Pulsar timing



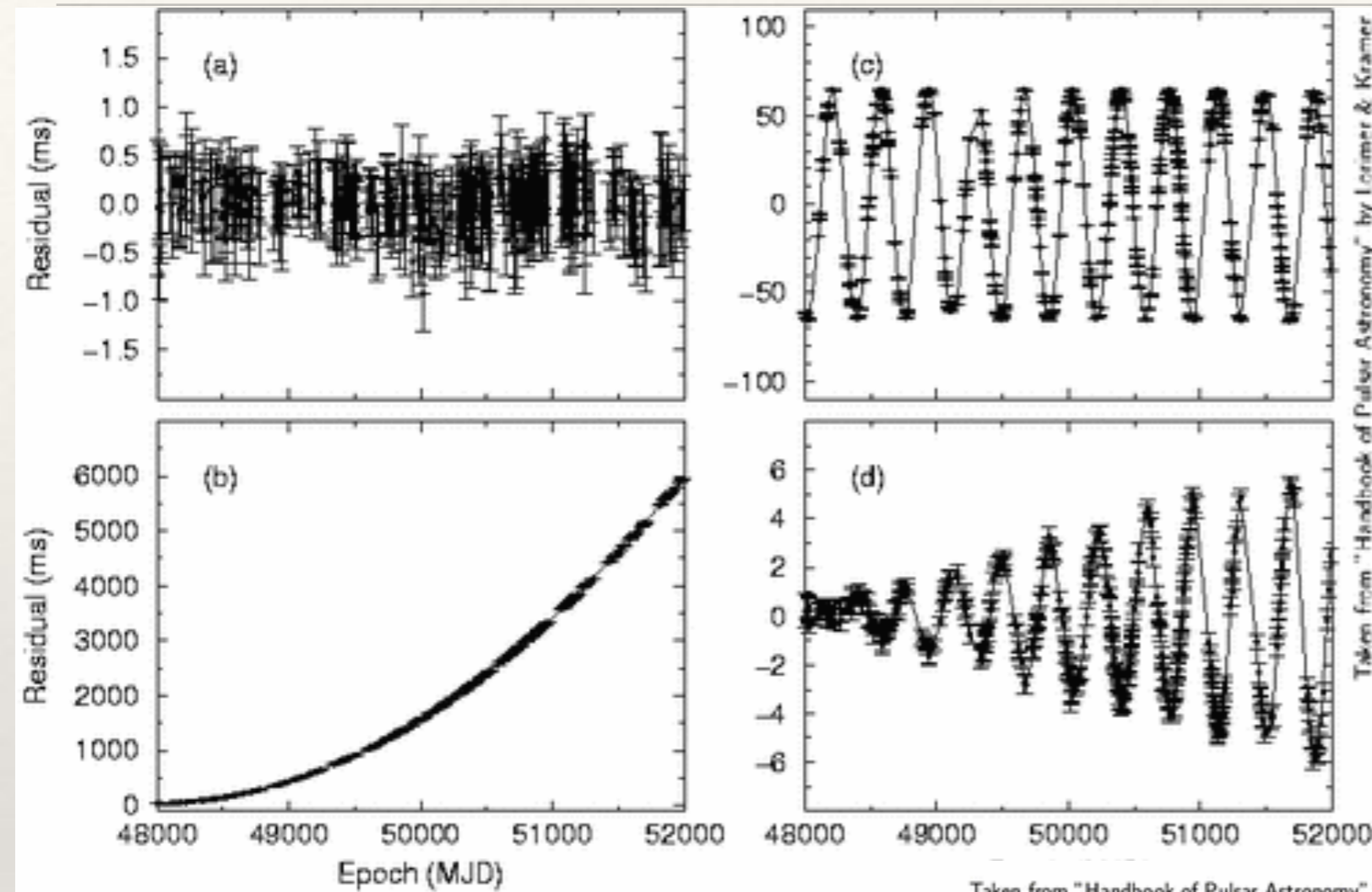
[Figs: credits  
S. Burke-Spolar & L. Lentati]

- Each observed radio pulse profile has a lot of micro-structure. If we average over  $\sim$ hour the (average) profile is very stable
- We can use the average pulse profile to estimate the time-of-arrival (TOA) of the pulses.
- The idea is to measure the TOA, and compare to the expected TOA. We know the spin of the pulsars, so we can predict the TOA. The difference between measured and expected TOA: *residuals*





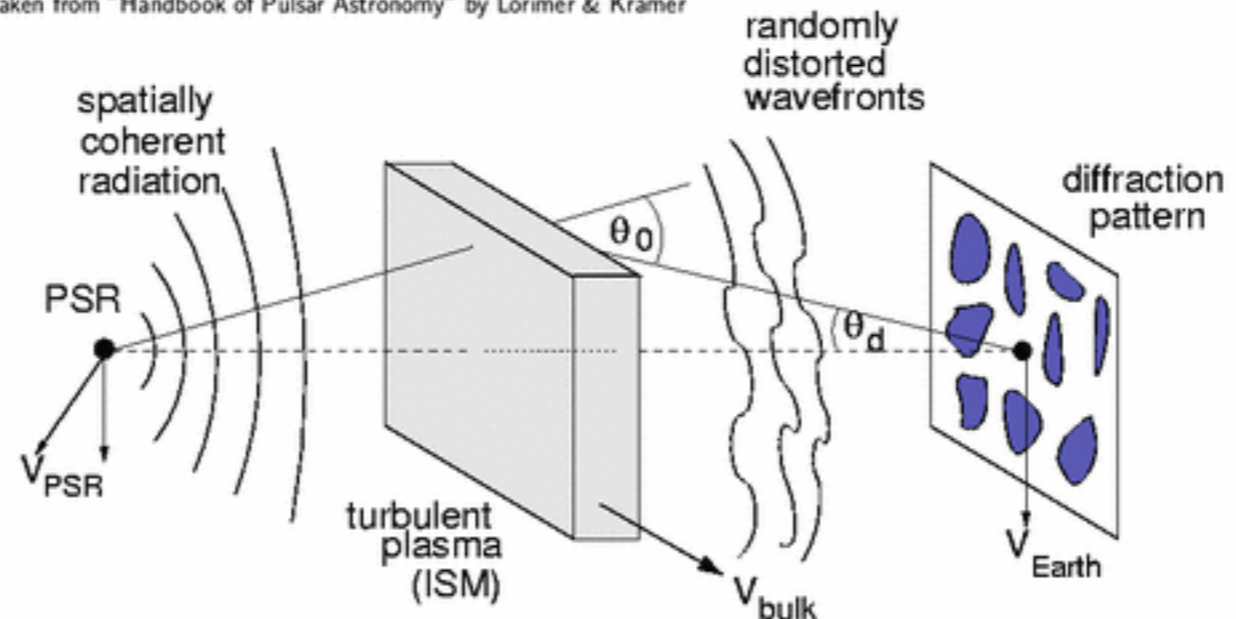
# Timing pulsars



- We need to build a timing model to make accurate prediction for TOAs - take into account various physical effects
- Dispersion of e/m wave and its time dependence
- Rate of change of rotation (b)
- Sky position of the pulsar (c)
- Proper motion of a pulsar (d)

○ Timing model could be quite complex if pulsar is in the binary

Taken from "Handbook of Pulsar Astronomy" by Lorimer & Kramer



# Residuals

- Building the timing model: depends on many parameters

$$t_{toa} = t_{toa}(P, \dot{P}, \ddot{P}, \Delta_{clock}, \Delta_{DM}(L), \Delta_{\odot-\oplus}, \Delta_E, \Delta_S)$$

$P, \dot{P}, \ddot{P}$  period of pulsar' rotation and its derivatives: spin-down

$\Delta_{clock}$  difference in the local clock and terrestrial standrad

$\Delta_{DM}(L)$  delays caused by propagation in the interstellar medium

$\Delta_{\odot-\oplus}$  Transformation from the local frame to the solar system barycentre

$\Delta_E$  Accounts for relative motion (Doppler) + gravitational redshift caused by the Sun, plantes or binary companion.

$\Delta_S$  Extra time required to trave in the curved spacetime containng Sun/companion (if in binary)

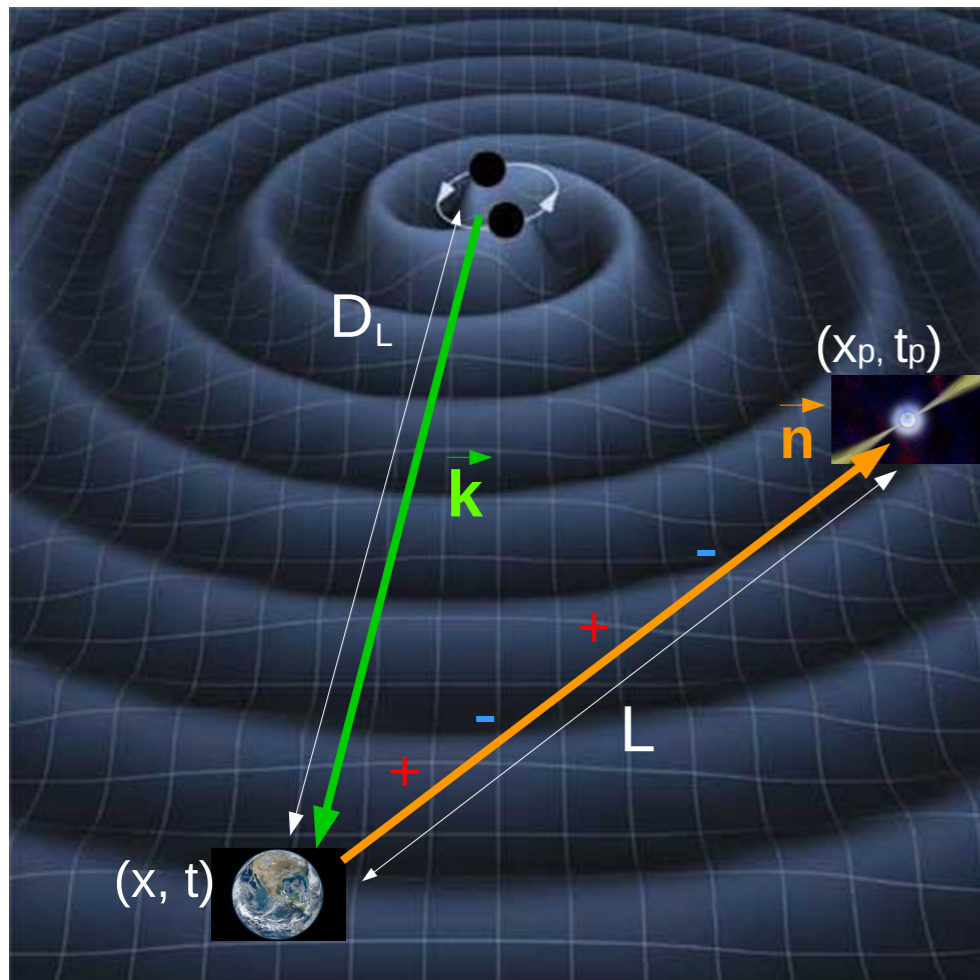
$$dt = t_{toa}^p - t_{toa}^o = dt_{errors} + \delta\tau_{GW} + noise$$

Errors in fitting the model

due to GWs



# Response to GW signal



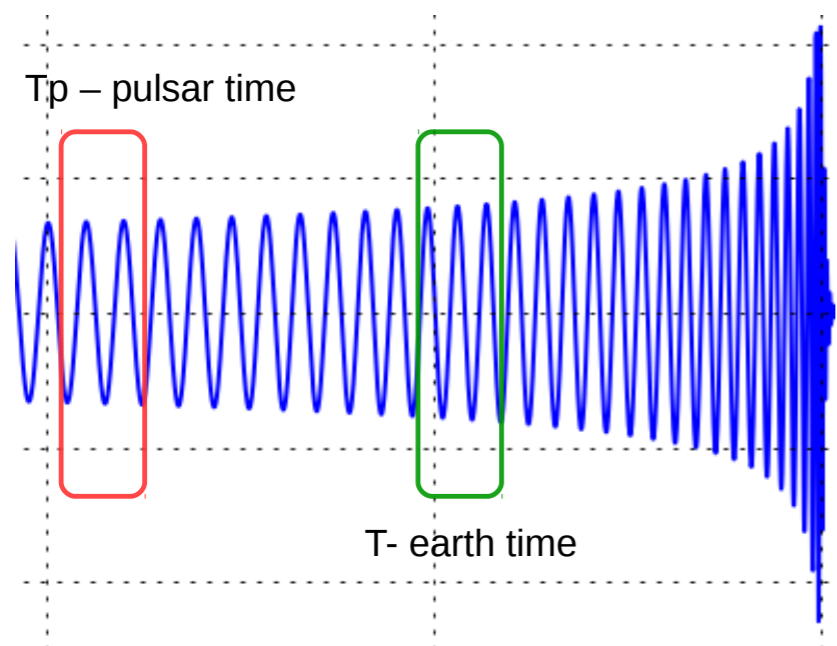
Important quantity which characterizes the response of any GW observatory is  $\epsilon = (2\pi f_* L/c)$

size of GW detector

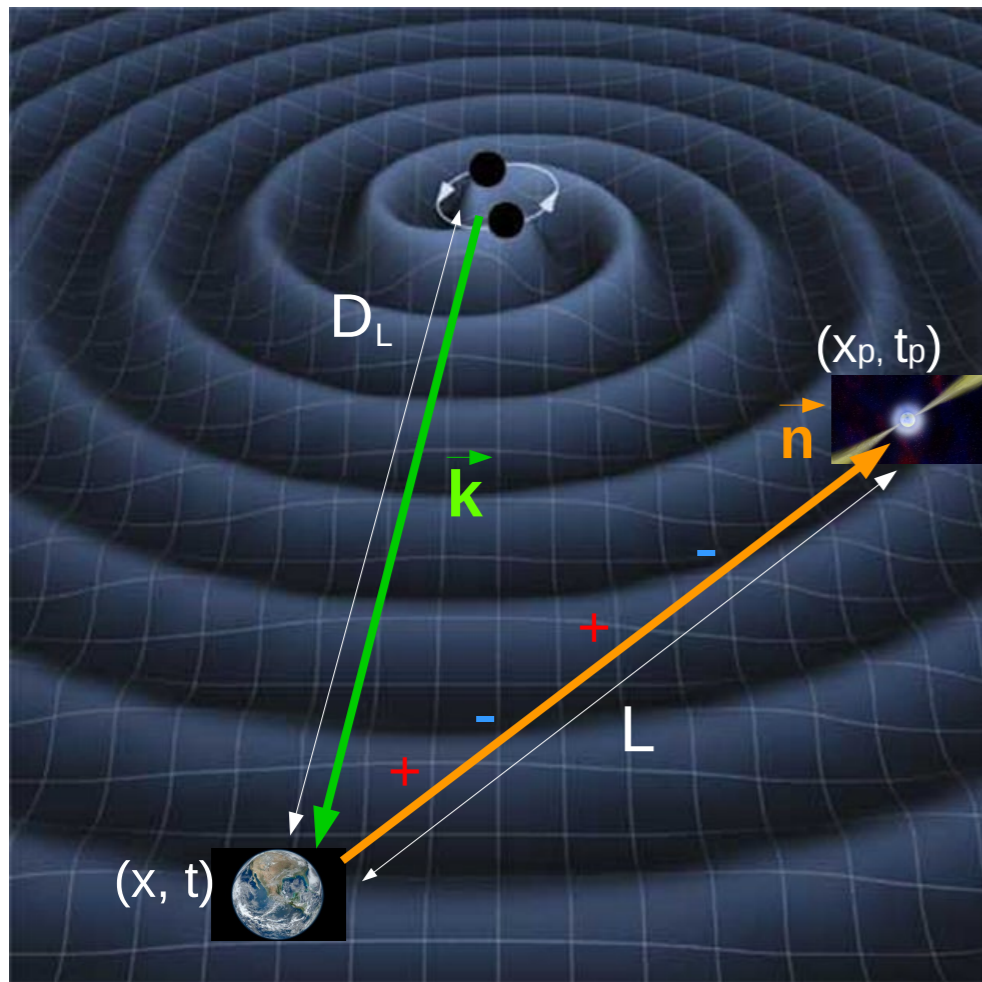
$$\epsilon \ll 1 \rightarrow R \propto h_{ij} n^i n^j$$

long wavelength approximation: already seen it with LIGO/Virgo

$$\epsilon = 1 \rightarrow \text{LIGO: } f^* \sim 12 \text{ kHz, LISA: } f^* \sim 0.05 \text{ Hz, PTA: } f^* \sim 0.002 \text{ nHz}$$



# Response to GW signal



$$dt = t_{toa}^p - t_{toa}^o = dt_{errors} + \delta\tau_{GW} + noise$$

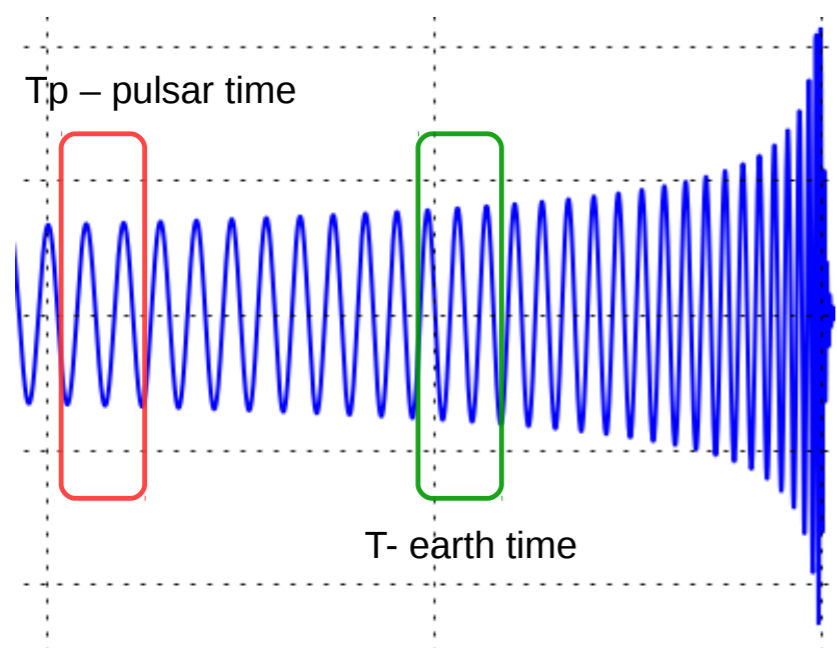
$$\delta\tau_{GW} = r(t) = \int_0^t \frac{\delta\nu}{\nu_0}(t') dt'; \quad \frac{\delta\nu}{\nu_0} = \frac{1}{2} \frac{\hat{n}^i \hat{n}^j \Delta h_{ij}}{1 + \hat{n} \cdot \hat{k}}$$

Familiar from LISA

$$\Delta h_{ij} = h_{ij}(t_p = t - L(1 + \hat{n} \cdot \hat{k})) - h_{ij}(t)$$

$t_p$  — pulsar time,  $\sim$  time of emission of the radio pulse:

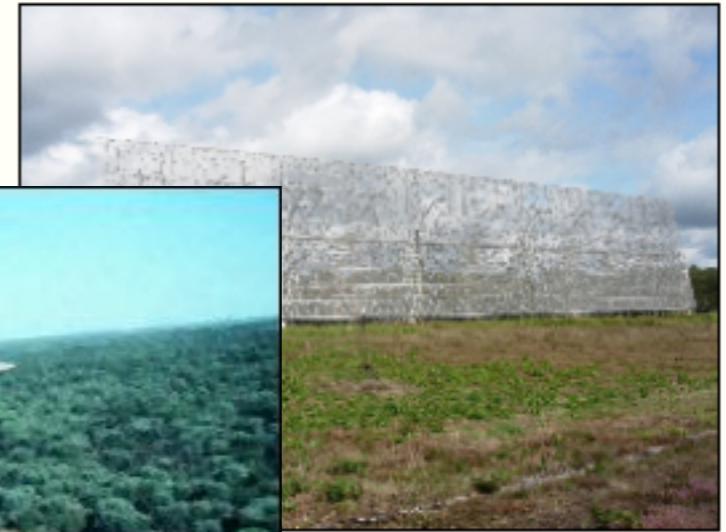
- depends on the relative position of a pulsar and GW source
- depends on the distance to the pulsar  $L$
- $L \sim$  few kpc  $\sim 10^5$  years — “pulsar” term  $h(t_p)$  contains info about the system  $10^5$  years in the past as compared to the “earth” term
- pulsar term depends on the pulsar.



# Radiotelescopes: EPTA



**The Effelsberg Radio Telescope**  
Effelsberg, Germany



**Jodrell Bank Radio Telescope**  
Cheshire, England



**SQUARE KILOMETER ARRAY \***



**The Sardinia Radio Telescope**  
Pranu Sanguni, Italy

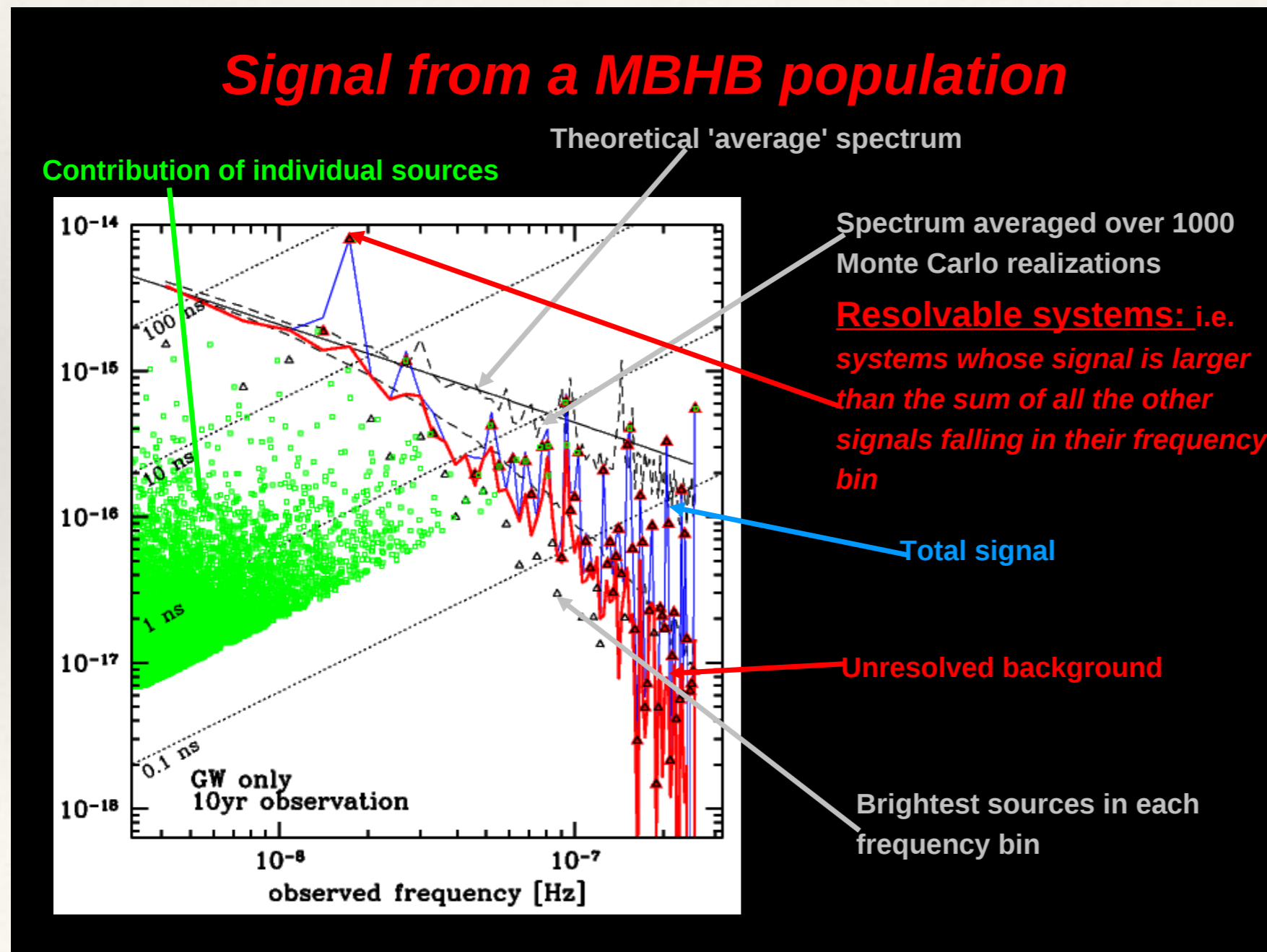


**The Westerbork Synthesis Radio Telescope**  
Westerbork, The Netherlands



# Supermassive black hole binaries

- Main sources are supermassive black hole binaries (mass  $10^7$  —  $10^{10}$  solar) on very broad orbit (period  $\sim$  year(s))
- The orbital evolution due to GW emission is very slow:  $\frac{dE}{dt} \propto \eta(M/r)^5$   
signal is (almost) monochromatic over period of observations



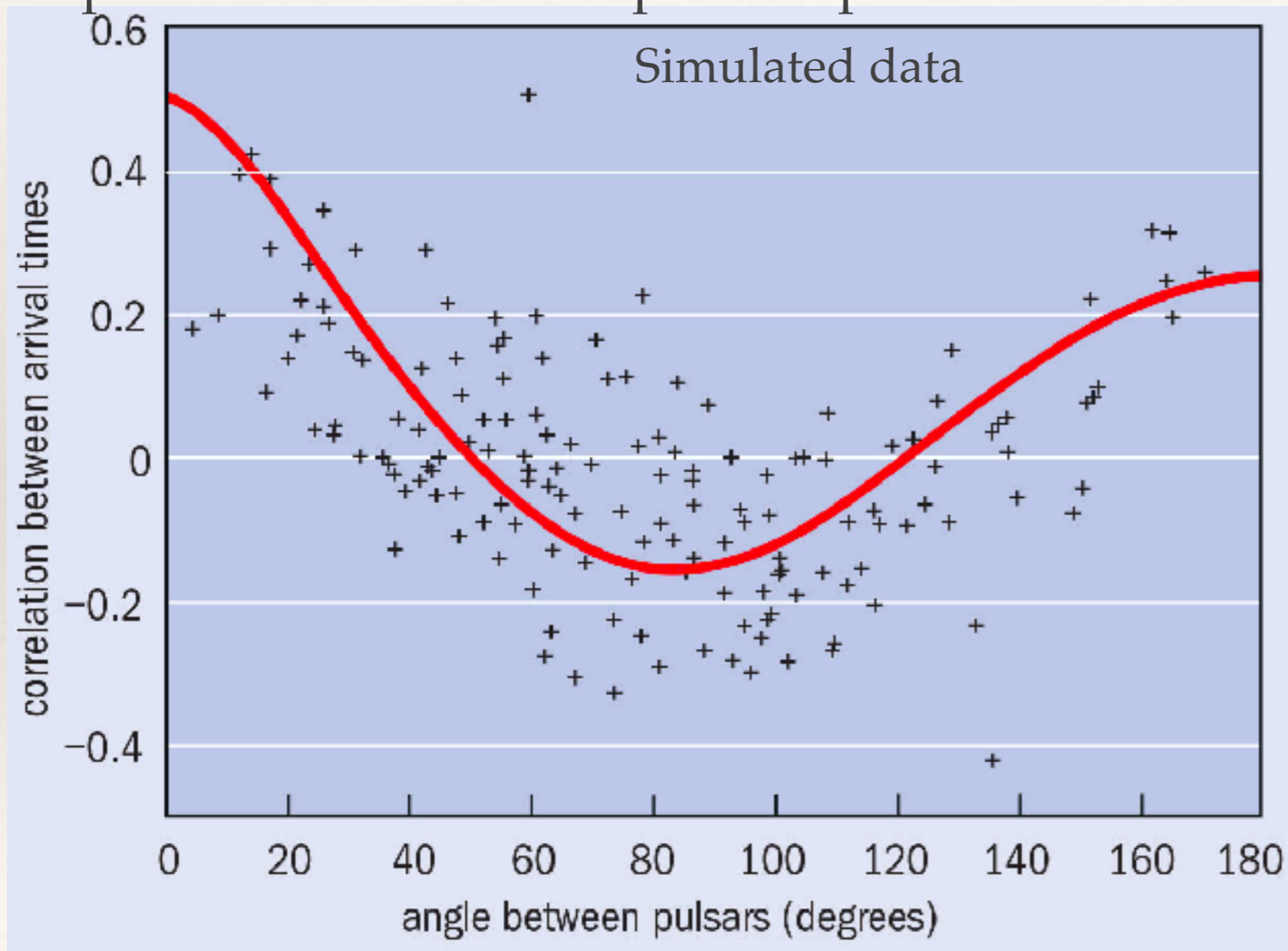
GW signal from the population of SMBH binaries: forms a stochastic signal at low freqs. (similar to Galactic binaries in LISA)

[Credits: A. Sesana]



# Correlation

- Stochastic GW signal — noise like signal which is correlated in observation of a pulsars. The correlation due to GW is very specific: Hellings-Downs curve.
- Correlation for the isotropic stochastic GW signal depends only on the angular separation between the pairs of pulsars.



[Fig. from IOP, Physics World]



# Data analysis: likelihood function

- The likelihood function (likelihood of the signal with given parameters is present in the observed data) is given as

$$P(\vec{\delta t} | \vec{\theta}) = \frac{1}{\sqrt{(2\pi)^n \det(C)}} \exp \left[ -\frac{1}{2} (\vec{\delta t} - \vec{s})^T C^{-1} (\vec{\delta t} - \vec{s}) \right]$$

data (TOA residuals)  
 unevenly sampled

deterministic  
 signal

Noise variance-  
 covariance matrix

- We pack *all* observed data (from all telescopes and all pulsars) into a single array of size n
- The GW signal (template) is located in “s” vector (resolvable binaries) and in noise matrix C

greek: pulsar index,  
 latin: data index

$$C_{\alpha i, \beta j} = C^{wn} \delta_{\alpha\beta} \delta_{ij} + C_{ij}^{rn} \delta_{\alpha\beta} + C_{ij}^{dm} \delta_{\alpha\beta} + C_{\alpha i, \beta j}^{GW}$$

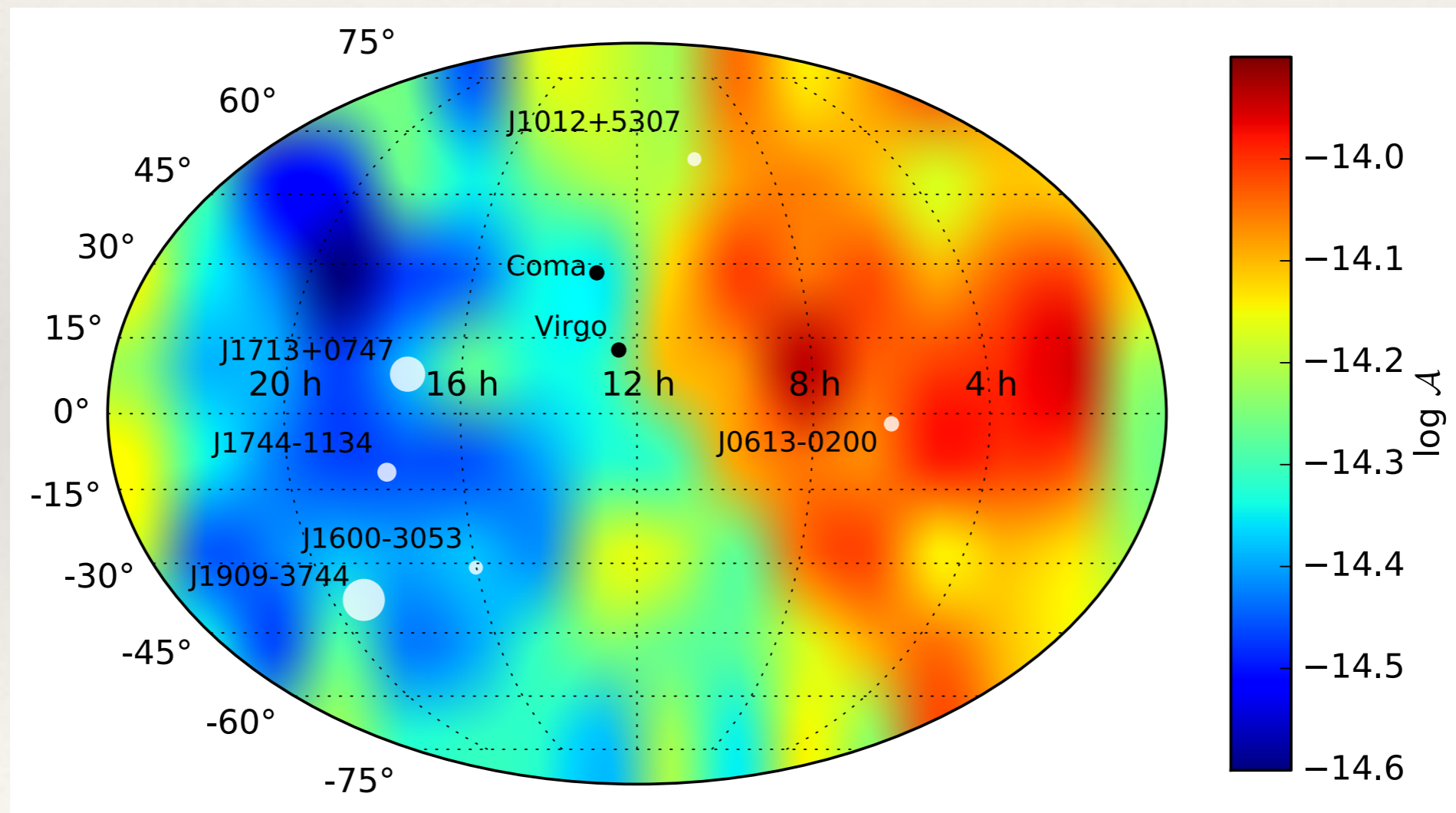
white noise      red noise      dispersion variation      stochastic GW signal





# Upper limit on GW in the PTA band

- GW not yet detected with PTA: GW are weak and need to be integrated out the noise (decades of observation time). We also need more good “timers”.
- Give a current sensitivity we can set upper limit on GW strain: for individual binaries and for the stochastic GW signal.

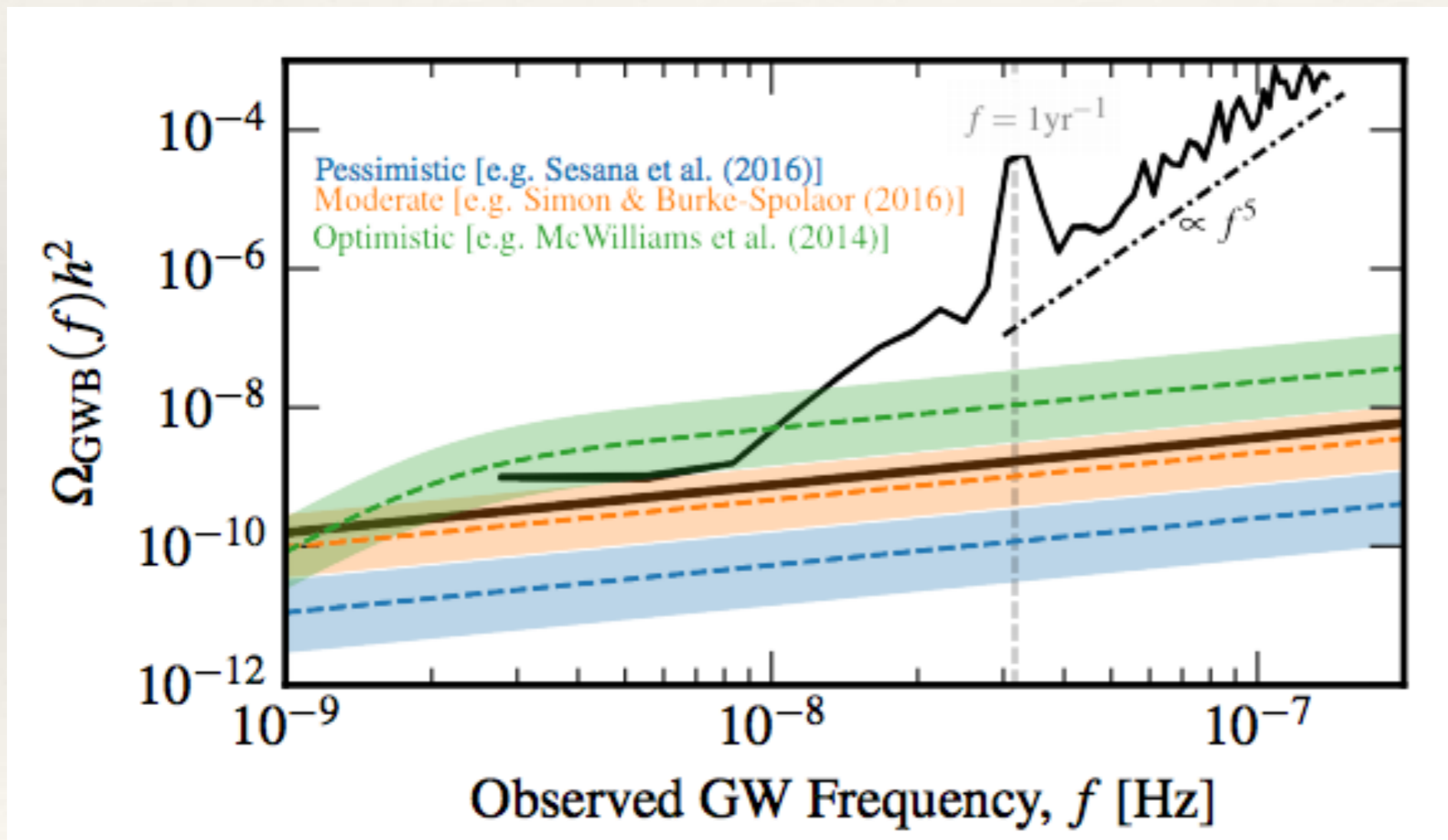


[Babak+ MNRAS (2015), EPTA]



# Upper limit on GW in the PTA

Current observational data allows us close some (over)optimistic astrophysical models and constrain some parameters



[Nanograv, arXiv: 1801.02617]

