

Gaussian-ology

Start with $P_G(x;1) \equiv \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$

Normalisation $\int_{-\infty}^{\infty} dx P_G(x;1) = 1$

Mean: $\langle x \rangle = \int_{-\infty}^{\infty} dx \cdot x P_G(x;1) = 0$

Variance: $\langle x^2 \rangle - \langle x \rangle^2 = \int_{-\infty}^{\infty} dx x^2 P_G(x;1) = 1$

Odd moments: $\langle x^{2n+1} \rangle = 0$

Even moments: $\langle x^{2n} \rangle = (2n-1)!!$

$$\left. \begin{aligned} & \left[N^2 = \left(\int_{-\infty}^{\infty} dx P_G(x;1) \right)^2 \right. \\ & = \int dx dy \frac{1}{2\pi} e^{-(x^2+y^2)/2} \\ & = \int_0^{\infty} dr r \int_0^{2\pi} d\theta \frac{1}{2\pi} e^{-r^2/2} \\ & = \int_0^{\infty} dz e^{-z} \left[z = r^2/2 \right] \\ & = 1 \end{aligned} \right\}$$

How to prove the above?

→ Fourier representation

$$\begin{aligned} \mathcal{F}[P_G](k) &\equiv \int_{-\infty}^{\infty} dx e^{-ikx} P_G(x;1) \\ &= \langle e^{-ikx} \rangle = e^{-k^2/2} \end{aligned}$$

$$\begin{aligned} \text{So } \langle x^n \rangle &= \int_{-\infty}^{\infty} dx \cdot x^n P_G \\ &= \left(i \frac{\partial}{\partial k} \right)^n \mathcal{F}[P_G](k) \Big|_{k=0} \end{aligned}$$

→ $n=0$: $\langle 1 \rangle = 1$ ✓

$n=1$: $\langle x \rangle = -ik e^{-k^2/2} \Big|_{k=0} = 0$

$n=2$: $\langle x^2 \rangle = -i^2 \frac{\partial^2}{\partial k^2} (k e^{-k^2/2}) \Big|_{k=0} = 1$

⋮ so on

Another way of writing this:

$$\begin{aligned} \left(-\frac{\partial}{\partial x}\right)^n P_G(x;1) &\equiv H_n(x) P_G(x;1) \\ &= \int_{-\infty}^{\infty} \frac{dk}{2\pi} e^{ikx} (-ik)^n e^{-k^2/2} \end{aligned}$$

$$\S (-ik)^n e^{-k^2/2} = \int_{-\infty}^{\infty} dx e^{-ikx} P_G(x;1) H_n(x)$$

$$\# \left[\text{Recall } \int_{-\infty}^{\infty} dx e^{-ikx} = 2\pi \delta_D(k) \right]$$

Note that $\langle H_n(x) \rangle = 0$ for all $n \geq 1$.

Introducing variance $\sigma^2 \neq 1$
→ dimensional analysis

$$\begin{aligned} P_G(x;1) \rightarrow P_G(x;\sigma^2) &= \frac{1}{\sigma} P_G(x/\sigma;1) \\ &= \frac{1}{\sqrt{2\pi}\sigma} e^{-x^2/2\sigma^2} \end{aligned}$$

$$\langle 1 \rangle = 1$$

$$\langle x \rangle = 0, \quad \langle x^2 \rangle = \sigma^2$$

$$\langle x^{2n+1} \rangle = 0; \quad \langle x^{2n} \rangle = \sigma^{2n} (2n-1)!!$$

$$\mathcal{F}[P_G](k) = \langle e^{-ikx} \rangle = e^{-k^2\sigma^2/2}$$

$$(-ik\sigma)^n e^{-k^2\sigma^2/2} = \int_{-\infty}^{\infty} dx e^{-ikx} P_G(x;\sigma^2) H_n(x/\sigma)$$

$$P_G(x;\sigma^2) H_n(x/\sigma) = \int_{-\infty}^{\infty} \frac{dk}{2\pi} e^{ikx} (-ik\sigma)^n e^{-k^2\sigma^2/2}$$

Introducing mean $\langle x \rangle = \mu \neq 0$

$$x \rightarrow x - \mu \text{ in } P_G.$$

Fourier transforms pick up phase $e^{-ik\mu}$

Generating functions

[i.e., function whose Taylor coefficients are the quantity of interest]

$$M(k) = \int [p_x] e^{ikx} = \langle e^{kx} \rangle = e^{\mu k} e^{k^2 \sigma^2 / 2}$$

is the moment generating function

$$\left[\text{Since } \langle x^n \rangle = \left(i \frac{\partial}{\partial k} \right)^n e^{-k^2/2} \Big|_{k=0} \right.$$

(e.g., with $\mu=0$)

→ with $k \rightarrow ik$

$$\left. \rightarrow e^{k^2 \sigma^2 / 2} = \sum_{n=0}^{\infty} \frac{1}{n!} k^n \langle x^n \rangle \right]$$

Cumulant generating function

$$C(k) \equiv \ln M(k)$$

So in this case

$$C(k) = \mu k + \frac{1}{2} k^2 \sigma^2$$

→ Defining property of Gaussian distribution: all cumulants of order $n \geq 3$ are zero.

→ So any p.d.f. with non-vanishing higher order cumulant(s) is non-Gaussian.

Multivariate Gaussian

Generalisation to $N > 1$ variables.

Now variance σ^2 is replaced with covariance matrix

$$C_{ij} = \langle x_i x_j \rangle - \langle x_i \rangle \langle x_j \rangle$$

So diagonal elements are individual variances $\sigma_j^2 = C_{jj}$.

while off-diagonals are cross-correlations.

Full expression:

$$P_{(N)}(x_1, \dots, x_N) = \frac{1}{(2\pi)^{N/2} |\det C|^{1/2}} e^{-\frac{1}{2}(x-\mu)^T C^{-1} (x-\mu)}$$

$\mu = (\mu_1, \dots, \mu_N)$ is mean vector.

Making sense of this

If C is diagonal = $\text{diag}(\sigma_1^2, \dots, \sigma_N^2)$

$$\text{then } P_{(N)}(x_1, \dots, x_N) = \prod_{j=1}^N P_1(x_j; \mu_j, \sigma_j^2)$$

→ obviously normalised correctly

If C is not diagonal, find rotation R such that

$$\Lambda = R C R^T = \text{diag}(\lambda_1^2, \dots, \lambda_N^2) \quad [R^T R = \mathbb{1}_N]$$

Define $y = R \cdot (x - \mu)$, so $x - \mu = R^T y$

Then

$$P_{(N)}(x_1, \dots, x_N) = \frac{1}{(2\pi)^{N/2} \lambda_1 \dots \lambda_N} e^{-\frac{1}{2} y^T \Lambda^{-1} y}$$

&

$$\left| \frac{\partial y}{\partial x} \right| = |\det R| = 1$$

$$\begin{aligned} & \text{[because} \\ & C = R \Lambda R^T \\ & \Rightarrow C^{-1} = R \Lambda^{-1} R^T \end{aligned}$$

$$\Rightarrow d^N x P_{(N)}(x_1, \dots, x_N) = \prod_{j=1}^N dy_j P_C(y_j; \lambda_j^2)$$

→ again, correctly normalised.

Marginalisation

Say we want to ignore x_N .

What is the resulting distribution $P_{(N-1)}(x_1, \dots, x_{N-1})$?

Again, Fourier transforming helps

$$\text{Use } P_C(y_j; \lambda_j^2) = \int_{-\infty}^{\infty} \frac{dq_j}{2\pi} e^{iq_j y_j} e^{-q_j^2 \lambda_j^2 / 2}$$

$$= d^N x P_{(N)}(x_1, \dots, x_N) = d^N y \int \frac{d^N q}{(2\pi)^N} e^{2q^T y} e^{-q^T \Lambda q / 2}$$

$$\text{But } y = R \cdot (x - \mu) \text{ \& } \Lambda = R C R^T$$

$$\text{So define } k = R^T q \Rightarrow d^N q \rightarrow d^N k.$$

Then

$$P_{(N-1)}(x_1, \dots, x_{N-1}) = \int \frac{d^N k}{(2\pi)^N} e^{2k^T (x - \mu)} e^{-k^T C k / 2}$$

Now marginalising is easy.

$$\begin{aligned}
P_{(N-1)}(x_1, \dots, x_{N-1}) & \equiv \int_{-\infty}^{\infty} dx_N P_{(N)}(x_1, \dots, x_N) \\
& = \int dx_N \int \frac{d^N k}{(2\pi)^N} e^{i k^T (x - \mu)} e^{-k^T C k / 2} \\
& = \int \frac{d^N k}{(2\pi)^N} e^{-k^T C k / 2} e^{i k_{(N-1)}^T (x_{(N-1)} - \mu_{(N-1)})} \\
& \quad \int dx_N e^{i k_N (x_N - \mu_N)} \\
& \quad \underbrace{\hspace{10em}}_{(2\pi) \delta_D(k_N)} \\
& = \int \frac{d^{N-1} k}{(2\pi)^{N-1}} e^{i k_{(N-1)}^T (x_{(N-1)} - \mu_{(N-1)})} e^{-k_{(N-1)}^T C_{(N-1)} k_{(N-1)} / 2}
\end{aligned}$$

$\underbrace{\hspace{10em}}_{\substack{\text{N-1 dimensional} \\ \text{Gaussian}}}$ where $C_{(N-1)ij} = C_{ij}$, $1 \leq i, j \leq N-1$

i.e., simply ignore the row & column in C that corresponds to the variable being integrated over.

$$P_{(N-1)}(x_1, \dots, x_{N-1}) = \frac{1}{(2\pi)^{\frac{N-1}{2}} |\det C_{(N-1)}|^{\frac{1}{2}}} e^{-\frac{1}{2} (x - \mu)^T \underbrace{C_{(N-1)}}_{\substack{\text{N-1 dimensional}}} (x - \mu)}$$

Conditional distributions:

Consider bivariate Gaussian

$$P(x_1, x_2) = \int \frac{d^2k}{(2\pi)^2} e^{i k^T (x - \mu)} e^{-k^T C k / 2}$$

$$C = \begin{pmatrix} \sigma_1^2 & c_{12} \\ c_{12} & \sigma_2^2 \end{pmatrix}$$

Bayes' thm says

$$P(x_2 | x_1) = \frac{P(x_2, x_1)}{P(x_1)}$$

Here $P(x_1) = \int P(x_1, x_2) dx_2 = P_1(x_1 - \mu_1, \sigma_1^2)$

Note that $-\ln P(x_2 | x_1)$ is the difference of a quadratic form in (x_2, x_1) and a quadratic form in x_1 .

$\Rightarrow \ln P(x_2 | x_1)$ is a quadratic form in x_2 , parametrised by x_1 (which also appears quadratically).

$\Rightarrow P(x_2 | x_1)$ is Gaussian!

\Rightarrow we only need to compute mean & variance

$$\begin{aligned}
\langle x_2 | x_1 \rangle &= \int dx_2 \cdot x_2 P(x_2 | x_1) \\
&= \int dx_2 \cdot \frac{1}{P_1(x_1)} \int \frac{d^2 k}{(2\pi)^2} e^{i k^T C k / 2} e^{i k^T (x - \mu)} x_2 \\
&= \frac{1}{P_1(x_1)} \int dx_2 \int \frac{d^2 k}{(2\pi)^2} e^{-k^T C k / 2} (\mu_2 - i \frac{\partial}{\partial k_2}) e^{i k^T (x - \mu)} \\
&= \frac{1}{P_1(x_1)} \int dx_2 \int \frac{d^2 k}{(2\pi)^2} e^{i k^T (x - \mu)} (\mu_2 + i \frac{\partial}{\partial k_2}) e^{-k^T C k / 2} \\
&= \frac{1}{P_1(x_1)} \int \frac{d^2 k}{(2\pi)^2} \cdot (2\pi) \delta_0(k_2) e^{i k_1 (x_1 - \mu_1)} \\
&\quad \times e^{-k^T C k / 2} (\mu_2 - i k_2 (C_{22} - i k_1 C_{12})) \\
&= \frac{1}{P_1(x_1)} \int \frac{dk_1}{(2\pi)} e^{i k_1 (x_1 - \mu_1)} e^{-k_1^2 \sigma_1^2 / 2} (\mu_2 - i k_1 C_{12}) \\
&= \frac{1}{P_1(x_1)} \cdot P_1(x_1 - \mu_1, \sigma_1^2) \left[\mu_2 + H_1 \left(\frac{x_1 - \mu_1}{\sigma_1} \right) \cdot \frac{C_{12}}{\sigma_1} \right]
\end{aligned}$$

$$\Rightarrow \left\langle x_2 | x_1 \right\rangle = \mu_2 + \frac{C_{12}}{\sigma_1^2} (x_1 - \mu_1) = \langle x_2 \rangle + (x_1 - \langle x_1 \rangle) \frac{\text{Cov}(x_1, x_2)}{\text{Var}(x_1)}$$

Similar calculation gives

$$\begin{aligned}
\text{Var}(x_2 | x_1) &= \langle x_2^2 | x_1 \rangle - \langle x_2 | x_1 \rangle^2 \\
&= \text{Var}(x_2) - \left(\frac{\text{Cov}(x_1, x_2)}{\text{Var}(x_1)} \right)^2 \\
&= \sigma_2^2 - C_{12}^2 / \sigma_1^2
\end{aligned}$$

[Verify this]

Bardeen et al. (1986) [BBKS] Appendix D
generalises this to multivariate conditional
Gaussians

If $p(\vec{Y}_A, \vec{Y}_B)$ is Gaussian with $\langle \vec{Y}_A \rangle = 0 = \langle \vec{Y}_B \rangle$

then

$p(\vec{Y}_B | \vec{Y}_A)$ is multivar. Gaussian

with

$$\langle \vec{Y}_B | \vec{Y}_A \rangle = \langle Y_B \otimes Y_A \rangle \langle Y_A \otimes Y_A \rangle^{-1} \vec{Y}_A^T$$

$$\begin{aligned} \& \langle \Delta Y_B \otimes \Delta Y_B | \vec{Y}_A \rangle \\ &= \langle Y_B \otimes Y_B \rangle - \langle Y_B \otimes Y_A \rangle \langle Y_A \otimes Y_A \rangle^{-1} \langle Y_A \otimes Y_B \rangle \end{aligned}$$

where $\Delta Y_B \equiv Y_B - \langle Y_B | Y_A \rangle$.

Random Fields & Correlations

Inflation tells us that "initial" density field had tiny fluctuations whose Fourier modes obeyed (nearly) Gaussian statistics.

$\delta_{\vec{k}} \sim$ Gaussian means that its real & imaginary parts are independent Gaussian variables.

Start with cube of side L [$V = L^3$]

$$\delta_{\vec{k}} = \int_V d^3x e^{-i\vec{k}\cdot\vec{x}} \delta(\vec{x})$$

$$\delta(\vec{x}) = \frac{1}{V} \sum_{\vec{k}} e^{i\vec{k}\cdot\vec{x}} \delta_{\vec{k}}$$

(Later we'll take the continuum limit

$$\frac{1}{V} \sum_{\vec{k}} \rightarrow \int \frac{d^3k}{(2\pi)^3})$$

Note that $\delta_{\vec{k}} = a_{\vec{k}} + i b_{\vec{k}}$ is a weighted sum of Gaussians

$\Rightarrow \delta(\vec{x})$ is also a weighted sum of Gaussians

$\Rightarrow \delta(\vec{x})$ is itself Gaussian distributed

Translational invariance (Statistical homogeneity)

$$\Rightarrow \langle \delta(\vec{x}) \delta(\vec{x} + \vec{s}) \rangle \text{ independent of } \vec{x}$$

$$\Rightarrow \sum_{\vec{k}} \sum_{\vec{k}'} e^{i\vec{k} \cdot \vec{x}} e^{-i\vec{k}' \cdot (\vec{x} + \vec{s})} \langle \delta_{\vec{k}} \delta_{\vec{k}'}^* \rangle$$

independent of \vec{x}

$$\Rightarrow \langle \delta_{\vec{k}} \delta_{\vec{k}'}^* \rangle \sim \delta_{\vec{k}, \vec{k}'}$$

\Rightarrow modes with different \vec{k} are uncorrelated or independent.

\rightarrow Important exception:

$$\delta_{-\vec{k}} = \delta_{\vec{k}}^*$$

because we want $\delta(\vec{x})$ to be real.

But this simply means that all d.o.f. are contained in the right half cube " $\vec{k} > 0$ " [Actually, e.g., $k_z > 0$]

There's no contradiction since each $\delta_{\vec{k}}$ has 2 d.o.f ($a_{\vec{k}}$ & $b_{\vec{k}}$), so total number of d.o.f = total number of modes.

How to count d.o.f in Gaussian field?

Fix \vec{k} (" >0 "), $\delta_{\vec{k}} = a_{\vec{k}} + i b_{\vec{k}}$

Then

$$\begin{aligned} da_{\vec{k}} db_{\vec{k}} g_{\vec{k}}(a_{\vec{k}}, b_{\vec{k}}) &= \text{Gaussian} \\ &= \frac{1}{2\pi\mu_{\vec{k}}^2} e^{-\frac{1}{2\mu_{\vec{k}}^2}(a_{\vec{k}}^2 + b_{\vec{k}}^2)} da_{\vec{k}} db_{\vec{k}} \end{aligned}$$

P.d.f. of modes is

$$\begin{aligned} \mathcal{P}[\{\delta_{\vec{k}}\}] &= \prod_{\vec{k} > 0} g_{\vec{k}}(\delta_{\vec{k}}) \\ &\propto e^{-\sum_{\vec{k} > 0} (a_{\vec{k}}^2 + b_{\vec{k}}^2) / 2\mu_{\vec{k}}^2} \end{aligned}$$

But $\delta_{\vec{k}^*} = \delta_{-\vec{k}} \Rightarrow a_{-\vec{k}} = a_{\vec{k}}, b_{-\vec{k}} = -b_{\vec{k}}$
 i.e. $a_{-\vec{k}}^2 = a_{\vec{k}}^2$ & $b_{-\vec{k}}^2 = b_{\vec{k}}^2$

$$\Rightarrow \mathcal{P}[\{\delta_{\vec{k}}\}] \propto \exp\left(-\sum_{\text{all } \vec{k}} \frac{1}{2} \frac{a_{\vec{k}}^2 + b_{\vec{k}}^2}{2\mu_{\vec{k}}^2}\right)$$

or

$$\boxed{\mathcal{P}[\{\delta_{\vec{k}}\}] \propto \exp\left(-\sum_{\text{all } \vec{k}} \frac{|\delta_{\vec{k}}|^2}{2P(\vec{k})}\right)}$$

$$\boxed{P(\vec{k}) = 2\mu_{\vec{k}}^2 = \langle |\delta_{\vec{k}}|^2 \rangle}$$

(statistical isotropy)

Rotational invariance demands

$$P(\vec{k}) = P(k)$$

$P(k)$ is called the power spectrum.

Alternative counting

Recast $a_{\mathbf{k}}, b_{\mathbf{k}}$ as amplitude $r_{\mathbf{k}}$ & phase $\phi_{\mathbf{k}}$

$$\delta_{\mathbf{k}} = r_{\mathbf{k}} e^{i\phi_{\mathbf{k}}}$$

$$\begin{aligned} \rightarrow & g_{\mathbf{k}}(r_{\mathbf{k}}, \phi_{\mathbf{k}}) dr_{\mathbf{k}} d\phi_{\mathbf{k}} \\ &= \frac{1}{2\pi\mu_{\mathbf{k}}^2} e^{-r_{\mathbf{k}}^2/2\mu_{\mathbf{k}}^2} r_{\mathbf{k}} dr_{\mathbf{k}} d\phi_{\mathbf{k}} \\ &= \frac{d\phi_{\mathbf{k}}}{(2\pi)} \cdot \frac{d(r_{\mathbf{k}}^2)}{P(\mathbf{k})} e^{-r_{\mathbf{k}}^2/P(\mathbf{k})} \end{aligned}$$

I.e., $\phi_{\mathbf{k}}$ is uniformly distributed in $[0, 2\pi)$ and $r_{\mathbf{k}}^2 = |\delta_{\mathbf{k}}|^2$ is exponentially distributed with $\langle r_{\mathbf{k}}^2 \rangle = P(\mathbf{k})$.

Uncorrelated modes

$$\Rightarrow P(\delta_{\mathbf{k}}, \delta_{\mathbf{k}'}) \sim g_{\mathbf{k}}(\phi_{\mathbf{k}}, r_{\mathbf{k}}) g_{\mathbf{k}'}(\phi_{\mathbf{k}'}, r_{\mathbf{k}'})$$

\Rightarrow Amplitudes as well as phases of different modes are uncorrelated.

\rightarrow So non-Gaussianity can show up as non-uniformity of phases, or phase correlations, or amplitude correlations (despite translational invariance)

In real space,

$$|\delta_{\mathbb{R}^2}|^2 = \int d^3x d^3y e^{-i\vec{k}_0 \cdot (\vec{x} - \vec{y})} \delta(\vec{x}) \delta(\vec{y})$$

$$= \mathcal{P}[\{\delta(\vec{x})\}]$$

$$\propto \exp -\frac{1}{2} \int_{\mathbb{E}} d^3x d^3y \delta(\vec{x}) \delta(\vec{y}) \frac{e^{-i\vec{k}_0 \cdot (\vec{x} - \vec{y})}}{P(k)}$$

Define

$$F(\vec{x} - \vec{y}) = \int_{\mathbb{E}} \frac{e^{-i\vec{k}_0 \cdot (\vec{x} - \vec{y})}}{P(k)}$$

$$= \mathcal{P}[\{\delta(\vec{x})\}] \propto \exp -\frac{1}{2} \int d^3x d^3y \delta(\vec{x}) F(\vec{x} - \vec{y}) \delta(\vec{y})$$

Note that $F(\vec{x} - \vec{y}) \neq \delta_0(\vec{x} - \vec{y})$,

i.e., uncorrelated modes

\Rightarrow real space 2-point function is non-trivial.

$F(\vec{x} - \vec{y})$ is like an inverse covariance matrix.

Continuum limit.

$$\frac{1}{V} \sum_{\vec{k}} \rightarrow \int \frac{d^3k}{(2\pi)^3} ; P(k) |_{\text{discrete}} \rightarrow V P(k)$$

$$\delta_{\vec{k}, \vec{k}'} \rightarrow \frac{(2\pi)^3}{V} \delta_D(\vec{k} - \vec{k}')$$

$$\text{So } \langle \delta_{\vec{k}} \delta_{\vec{k}'}^* \rangle = \delta_{\vec{k}, \vec{k}'} P(k)$$

$$\boxed{\langle \delta_{\vec{k}} \delta_{\vec{k}'}^* \rangle = (2\pi)^3 \delta_D(\vec{k} - \vec{k}') P(k)}$$

1-point p.d.f of $\delta(\vec{x})$:

$$P(\delta) = P_G(\delta; \sigma^2)$$

where

$$\sigma^2 \equiv \langle \delta(\vec{x})^2 \rangle = \int \frac{d^3k d^3k'}{(2\pi)^6} e^{i(\vec{k}' - \vec{k}) \cdot \vec{x}} \langle \delta_{\vec{k}} \delta_{\vec{k}'}^* \rangle$$

$$= \int \frac{d^3k}{(2\pi)^3} P(k) \equiv \int dk k \Delta^2(k)$$

$$\Delta^2(k) = \frac{k^3 P(k)}{2\pi^2}$$

For CDM-like spectrum, $P(k)$ has too much 'UV' power

→ δ needs to be smoothed

$$\delta_R(\vec{x}) = \int d^3x' \delta(\vec{x}') W_R(\vec{x} - \vec{x}') = \int \frac{d^3k}{(2\pi)^3} e^{i\vec{k} \cdot \vec{x}} \delta_{\vec{k}} W_R(k)$$

$$P(\delta_R) = P_G(\delta_R; \sigma^2(R))$$

$$\sigma^2(R) = \int dk k \Delta^2(k) W_R(k)^2$$

usually
spherically
symmetric

[Try proving this

$$\text{using } P(\delta_R) = \int \mathcal{D}\xi \mathcal{D}\zeta P(\delta_R, \xi \delta_{\vec{k}} \zeta) = \int \mathcal{D}\xi \mathcal{D}\zeta \delta_D(\delta_R - \frac{1}{V} \sum_{\vec{k}} \xi_{\vec{k}} W_R(k) e^{i\vec{k} \cdot \vec{x}}) P(\xi \delta_{\vec{k}} \zeta)$$

E.g. common quantity is σ_8 where

$$\sigma_8^2 = \int dk k \Delta_{\text{lin}}^2(k) W_{R=8\text{mpc}/h}^{\text{(TopHat)}}(k)^2$$

Correlation function

- Recall $F(\vec{x}-\vec{y})$ was analog of inverse covariance matrix.
- Analog of covariance matrix is the correlation function

$$\boxed{\xi(\vec{r}) \equiv \langle \delta(\vec{x}) \delta(\vec{x}+\vec{r}) \rangle}$$

Simple calculation shows

$$\xi(\vec{r}) = \int \frac{d^3k}{(2\pi)^3} e^{i\vec{k}\cdot\vec{r}} P(k)$$

So that, due to rotational invariance

$$\boxed{\xi(\vec{r}) = \xi(r) = \int d^3k \Delta^2(k) \frac{\sin kr}{kr}}$$

Examples:

- $P(k) \propto k^n$, $-3 \leq n \leq -1$

$$\Delta^2(k) \propto k^{n+3}$$

$$\xi(r) \propto r^{-(n+3)} \left| \Gamma(n+2) \sin \frac{n\pi}{2} \right|$$

- $P(k) \propto k e^{-kR_c}$ [toy CDM (\sim WDM)]

$$\Delta^2(k) \propto k^4 e^{-kR_c}$$

$$\xi(r) \propto \frac{(3R_c^2 - r^2)R_c^4}{(R_c^2 + r^2)^3} = \frac{(3 - (r/R_c)^2)}{(1 + (r/R_c)^2)^3}$$

As $r \rightarrow 0$: $\xi(r) \rightarrow \text{constant}$

As $r \rightarrow \infty$: $\xi(r) \rightarrow 0^-$, having crossed
zero at ~~$r = R_c/\sqrt{3}$~~
 $r = \sqrt{3} R_c$.

- Full CDM: $P(k) = A (k/k_b)^{n_s} T(k)^2$ → primordial slope, ≈ 1

$T(k)$ describes decay of modes that entered the Hubble radius during radiation domination.

$$T(k \rightarrow 0) \rightarrow 1$$

$$T(k \gg k_{eq}) \rightarrow \frac{12}{(k/k_{eq})^2} \ln(k/8k_{eq})$$

where $k_{eq} = a_{eq} H(a_{eq})$, $\Omega_{m0} a_{eq}^{-3} = \Omega_{rad,0} a_{eq}^{-4}$

$\xi(r)$ needs numerical evaluation.