

Parameter inference: $\overset{\text{model}}{\uparrow}$

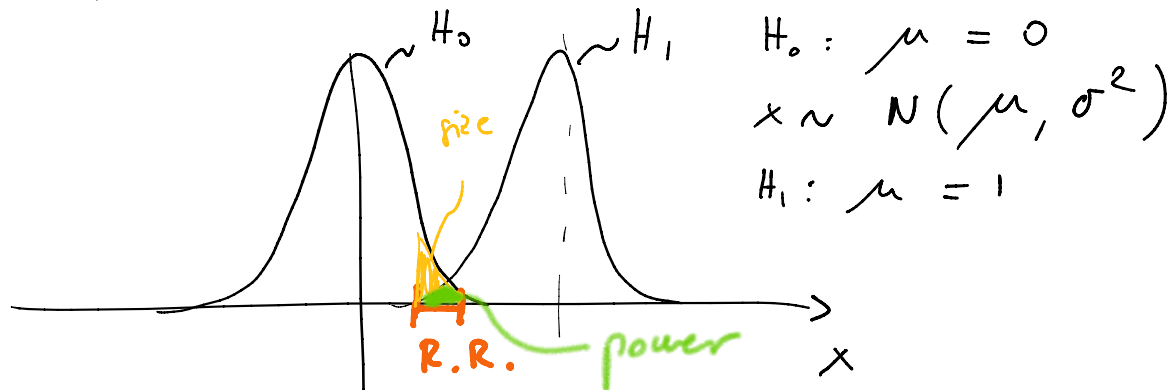
$$P(\Theta | d, M) = \frac{P(d | \Theta, M) P(\Theta | M)}{P(d | M)}$$

evidence,
model likelihood

Aside: Hypothesis testing

Null hypothesis H_0 ; alternative H_1

- Rejection region: a region where we reject H_0 and accept H_1
- The size: the prob. of an outcome in the rejection region if H_0 is true (low size)
- The power of the rejection region: prob. of getting an outcome in the rejection region if H_1 is true (high power)



Type I error: if H_0 is true, reject H_0 .
II = if H_1 is true, accept H_0 .

Level of significance of the test α : prob. of committing a type I error

$$\alpha(R) = P_{H_0}(\text{accept } H_1) = P_{H_0}(x \in R)$$

Power of the test = $1 - \text{Prob}(\text{type II error})$

Neyman - Pearson Lemma :

$$L(x) = \frac{P_{H_1}(x)}{P_{H_0}(x)}$$

Set the significance α ($\alpha = 0.05$)

Accept H_0 if $L(x) < \alpha$

Reject H_0 if $L(x) > \alpha$

This is the most powerful test.

For α :

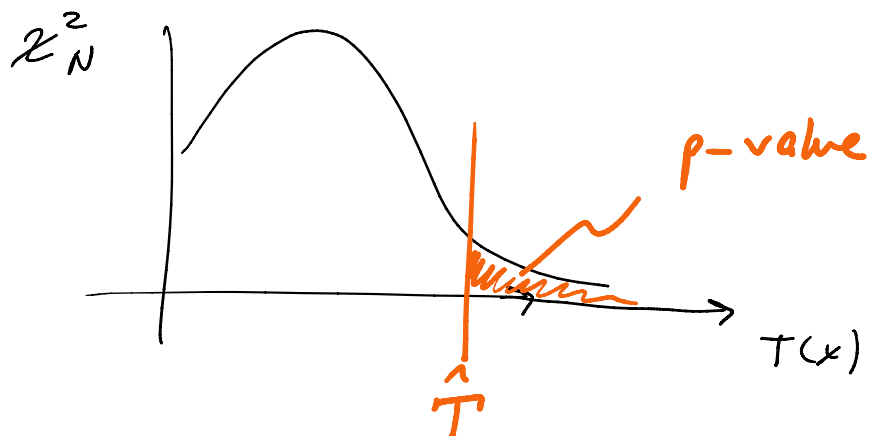
$$H_0 : \mu = 0$$

$x_i \sim N(\mu, \sigma^2)$, N draws

$$T(x) = \sqrt{N} \frac{\bar{x}}{\sigma}, \text{ distributed as } \chi_N^2$$

Size of the test :

$$P_{\mu=0} (T(x) \geq c) = P_{\mu=0} \left(\frac{\sqrt{N} \bar{x}}{\sigma} \geq c \right)$$



Bayesian approach : (BMC)

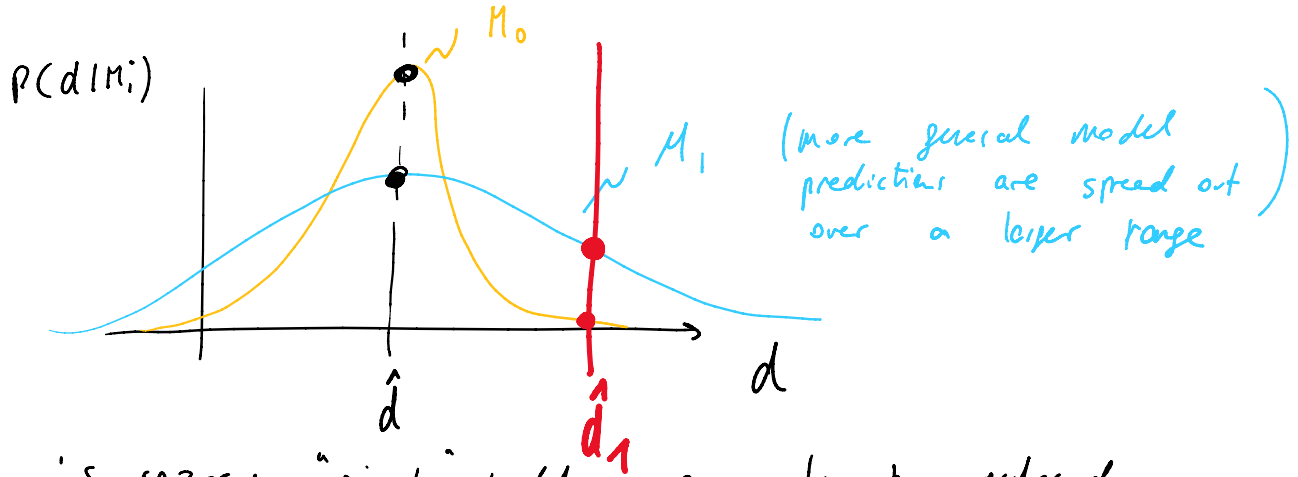
Bayesian approach : (BMC)

$$P(M|d) = \frac{P(d|M)P(M)}{P(d)}$$

$$P(d|M) = \int p(d|\theta, M) p(\theta|M)$$

$$\underbrace{\frac{P(M_0|d)}{P(M_1|d)}}_{\text{posterior odds}} = \underbrace{\frac{P(d|M_0)}{P(d|M_1)}}_{\text{BAYES FACTOR } B_{01}} \cdot \underbrace{\frac{P(M_0)}{P(M_1)}}_{\text{prior odds}}$$

Evidence is a predictive prob. for d :



Occam's razor: "simple" models are to be preferred.

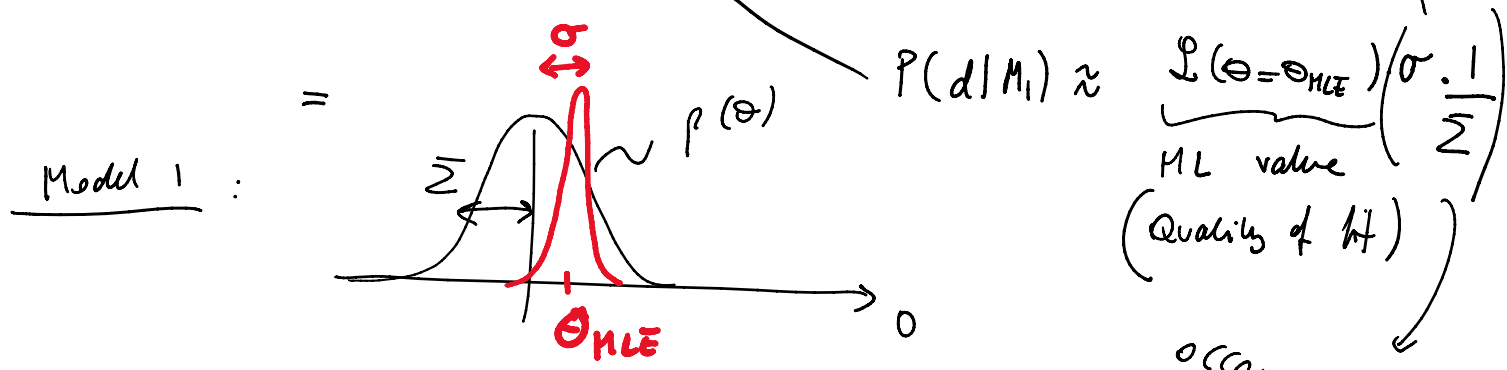
1D example : M_0 no free parameters : $\theta = 0 \rightarrow \delta(\theta)$
 M_1 with 1 parameter, $p(\theta)$, $x \sim N(\theta, \sigma^2)$
 observed \hat{x}

$$B_{01} = \frac{e^{-\frac{1}{2}(0-\hat{x})^2/\sigma^2}}{\int d\theta e^{-\frac{1}{2}(\theta-\hat{x})^2/\sigma^2} p(\theta)}$$

($\frac{\text{size } d}{\text{size of prior}}$)

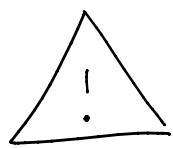
$$B_{01} = \frac{e^{-\lambda^2}}{\int d\theta e^{-\frac{1}{2} \frac{(\theta - \hat{\theta})^2}{\sigma^2}} f(\theta)} \quad (2)$$

(size of prior)



Occam's factor

Prior dependency in BMC never goes away!



penalizes "wasted" parameter space

$$\ln(B_{01}) = \ln \left[\sqrt{1 + \left(\frac{\sigma}{\Sigma}\right)^{-2}} \exp \left(-\frac{1}{2} \frac{\lambda^2}{1 + \left(\frac{\sigma}{\Sigma}\right)^2} \right) \right]$$

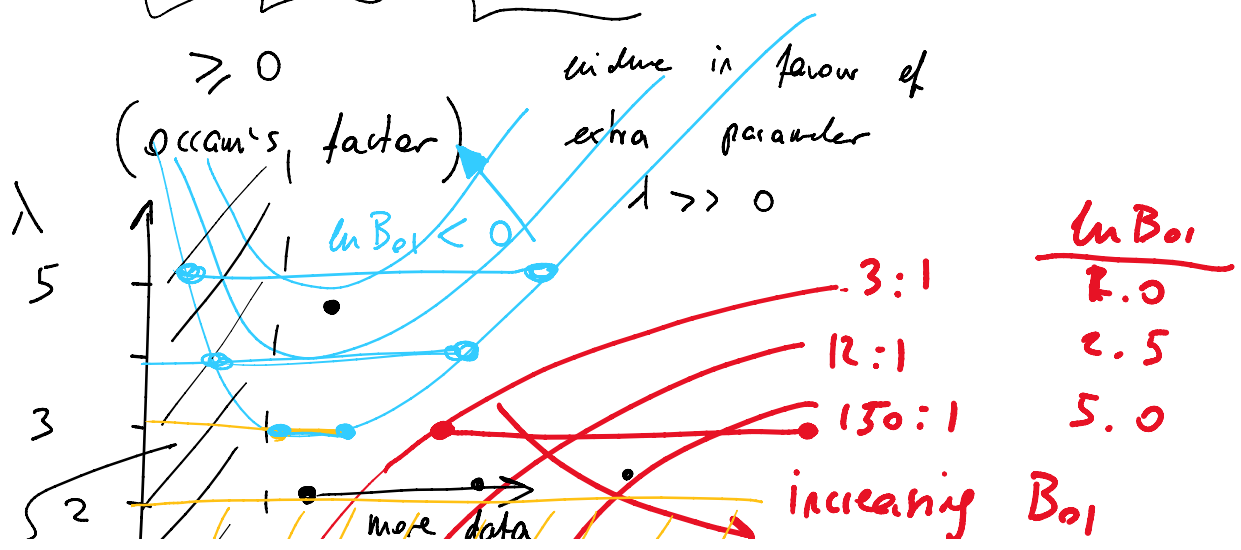
informative data

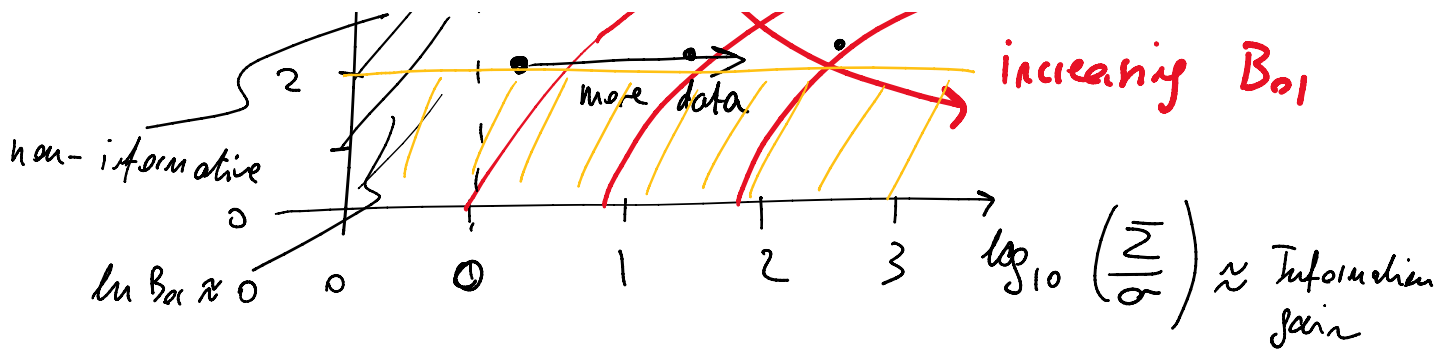
$$\lambda = \frac{|\theta_{MLE}|}{\sigma} \quad (\# \text{ of } \sigma \text{ "detection"})$$

$\left(\frac{\sigma}{\Sigma} \ll 1\right)$

$$\approx \ln \left(\frac{\Sigma}{\sigma} \right) - \frac{\lambda^2}{2} = \ln B_{01}$$

(Lindley Paradox)





Computation :

For nested models : $M_0 : \theta, p_0(\theta)$
 $M_1 : \{\theta, w\}, p_1(\theta, w) \rightarrow M_0 \text{ for } w = w_*$

Assuming separable priors : $p_1(\theta, w) = p_1(w) \cdot p_0(\theta)$

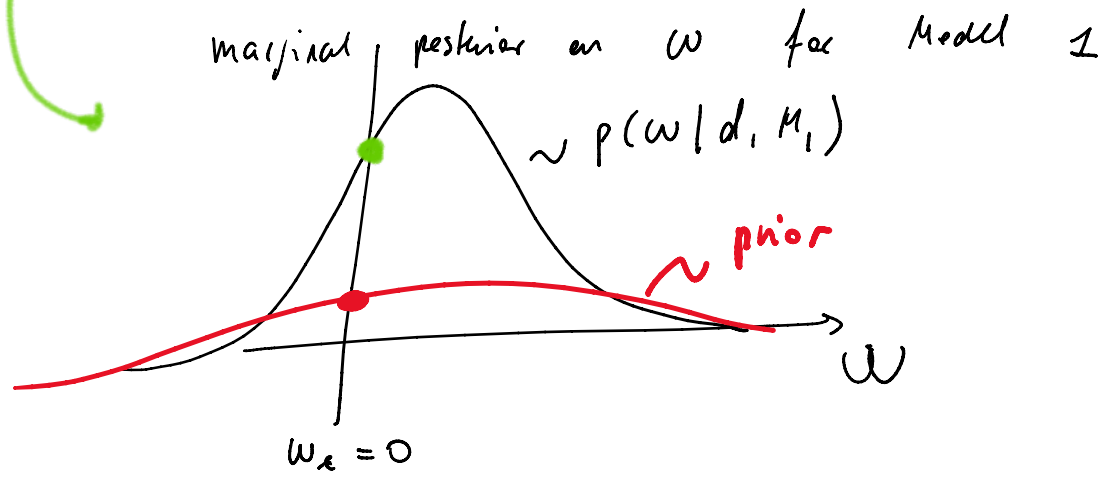
Then :

$$B_{01} = \frac{p(w = w_* | d, M_1)}{p_1(w = w_* | M_1)}$$

prior on additional parameter

Savage-Dickey density ratio (SDDR)

doesn't depend on θ ; does NOT depend on $p_0(\theta)$



Numerical tools : nested sampling; designed to compute $p(d | M) = \int p(\theta | d) p(\theta) d\theta$

$D \approx 30$
(ellipsoidal)

For $D \gg 1$: (PolyChord)

MCMC samples reweighting : $D \approx 10$

$$P(d|M) = \int P(\theta|d) P(\theta) d\theta$$
$$\rightarrow \int_0^1 \mathcal{I}(x) dx$$