



2453-14

School on Modelling Tools and Capacity Building in Climate and Public Health

15 - 26 April 2013

Expectile and Quantile Regression and Other Extensions

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# Expectile and Quantile Regression and Other Extensions

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A presentation at School on Modelling Tools and Capacity Building in Climate and Public Health

ICTP, Trieste, Italy

**Objectives and Questions** 

Objective:

Introduce other regression beyond the mean

- Are there median regression models
- What about regression at the mode?
- Can we fit regression model at any other locations too?
- What of the variance, skewness and kurtosis?

#### **Prelimaries**

- Given a r.v.  $y \sim f(\cdot)$  then  $-E(Y) = \mu$  defines the mean;  $-Var(Y) = \sigma^2$  is the variance.
- General assumption: two measures completely define the distribution (**stationarity** concept)
- Classical regression is often characterized by relating to the mean  $-E(y|x) = \beta x$  if x is the set of covariates.  $-y \sim N(\mu, \sigma^2)$ , then  $\mu = \beta x$

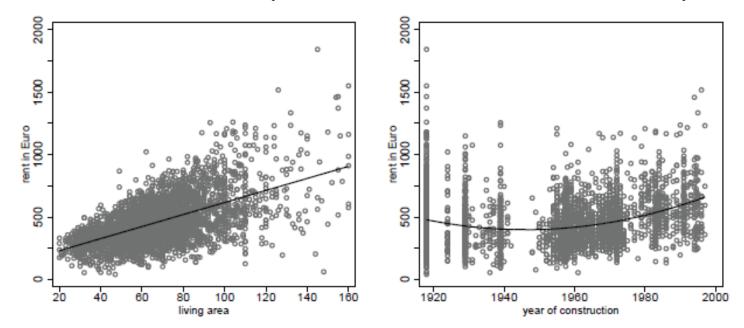
$$-y \sim Bern(\pi)$$
, then  $\log\left(\frac{\pi}{1-\pi}\right) = \beta x$   
 $-y \sim Pois(\lambda)$ , then  $\log(\lambda) = \beta x$ .

- Statisticians are mean lovers (Friedman, Friedman & Amoo)
- It goes like:

We are "mean" lovers. Deviation is considered normal. We are right 95% of the time. Statisticians do it discretely and continuously. We can legally comment on someone's posterior distribution.

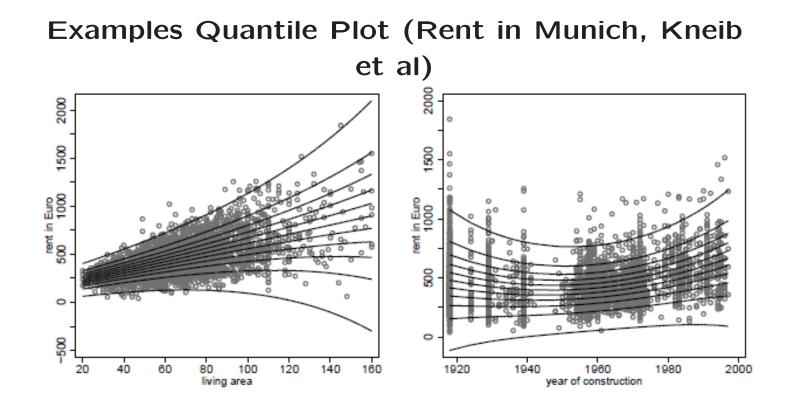
• Why the mean?

-Mean regression reduces complexity -However, the mean is not sufficient to describe a distribution.



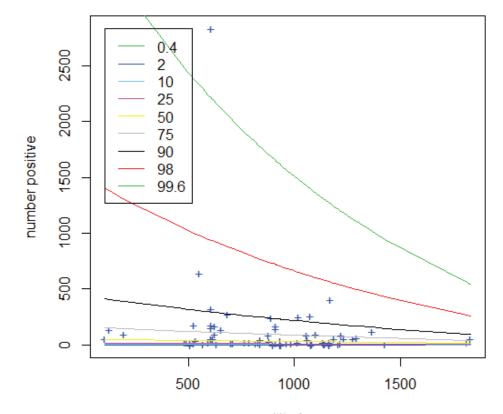
# Examples Plot (Rent in Munich, Kneib et al)

4



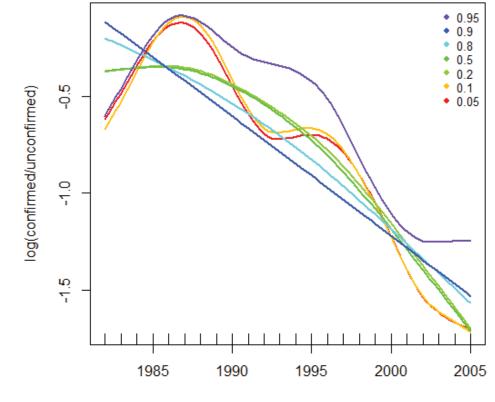
## Example (Malaria vs Altitude , Kazembe et al)

centile curves



altitude

6



year

7

### Motivating Examples

• Epidemiology and Public Health:

-Investigating height, weight and body mass index as a function of different covariates e.g. age (Wei et al. 2006)

-Exploring stunting curves in African and Indian children (Yue et al. 2012)

-Generating age-specific centile charts (Chitty et al. 1994).

• Economics:

-Study of determinants of wages (Koenker, 2005).

### • Education:

-The performance of students in public schools (Koenker and Hallock, 2001).

## • Climate data:

- Spatiotemporal analysis of Boston temperature (Reich 2012)

#### Double GLM

 Classical regression assumes a homogeneous variance

$$egin{aligned} y | m{x}, arepsilon \sim N(m{eta}m{x}, \sigma^2) \ & arepsilon \sim N(0, \sigma^2) \end{aligned}$$

- However variance heterogeneity (heteroscedasticity) is an order in real-life problems
- The variance of the response may depend on the covariates

Normal regression example:
 Regression for mean and variance of a normal distribution

$$y_i = \eta_{i1} + \exp(\eta_{2i})\varepsilon_i, \quad \varepsilon_i \sim N(0, 1)$$

such that

$$E(y_i | x_i) = \eta_{1i}$$
$$Var(y_i | x_i) = \exp(\eta_{2i})^2$$

Regression for location, scale and shape

- In general: Specify a distribution for the response on all parameters and relate to the predictors.
- The generalized additive model location, scale and shape (GAMLSS) is a statistical model developed by Rigby and Stasinopoulos (2005).
- For a probability (density) function  $f(y_i|\mu_i, \sigma_i, \nu_i, \tau_i)$ conditional on  $(\mu_i, \sigma_i, \nu_i, \tau_i)$  a vector of four distribution parameters, each of which can be a function

to the explanatory variables  $\left(X
ight)$ 

$$g_1(\mu) = \eta_1 = \beta_1 X_1 + \sum_{j=1}^{J_1} h_{j1}(x_{j1})$$
 (1)

$$g_2(\sigma) = \eta_2 = \beta_2 X_2 + \sum_{j=1}^{J_2} h_{j2}(x_{j2})$$
 (2)

$$g_3(\nu) = \eta_3 = \beta_3 X_3 + \sum_{j=1}^{J_3} h_{j3}(x_{j3})$$
 (3)

$$g_4(\tau) = \eta_4 = \beta_4 X_4 + \sum_{j=1}^{J_4} h_{j4}(x_{j4})$$
 (4)

- GAMLSS assumes different models for each parameter
  - -Model 1: Mean regression
  - -Model 2: Dispersion regression
  - -Model 3: Skewness regression
  - -Model 4: Kurtosis regression
- Other summary measures are also permissible

### **Quantile Regression**

- Quantile, Centile, and Percentile are terms that can be used interchangeably
   -A 0.5 quantile ≡ 50 percentile, which is a median.
- Related terms are quartiles, quintiles and deciles: divides distribution into 4, 5, and 10 parts
- Quantiles are related to the median.
- Suppose Y has a cumulative distribution  $F(y) = P(Y \le \tau)$ . Then  $\tau$ th quantile of Y is defined to be

$$Q(\tau) = \int_{-\infty}^{\tau} f(y) dy = \inf\{y : \tau \le F(y)\}$$

for  $0 < \tau < 1$ .

- The quantile regression drops the parametric assumption for the error/ response distribution.
- Fit separate models for different asymmetries  $\tau \in [0, 1]$ :

$$y_i = \eta_{i\tau} + \varepsilon_{i\tau}$$

• Instead of assuming  $E(y_i \leq \eta_{i\tau} = 0, \text{ one assumes})$ 

$$P(\varepsilon_{i\tau} \le 0) = \tau$$

or

$$F_{y_i}(\eta_{i\tau}) = P(\varepsilon_{i\tau} \le 0) = \tau$$

- This gives a set of regression function at any assumed quantile value.
- Assumptions:
  - -it is distribution-free since it does not make any specific assumption on the type of errors
    -it does not even require i.i.d errors
    -it allows for heterogeneity.

#### **Expectile Regression**

• Expectiles are related to the mean, as are quantiles related to the median.

$$\begin{array}{ll} \sum_{i=1}^{n} |y_i - \eta_i| \to \min & \sum_{i=1}^{n} w_\tau(y_i, \eta_{i\tau}) |y_i - \eta_{i\tau}| \to \min \\ \text{median regression} & \text{quantile regression} \end{array}$$

 $\begin{array}{ll} \sum_{i=1}^{n} |y_i - \eta_i|^2 \to \min & \sum_{i=1}^{n} w_\tau(y_i, \eta_{i\tau}) |y_i - \eta_{i\tau}|^2 \to \min \\ \text{mean regression} & \text{expectile regression} \end{array}$ 

where  $w_{\tau}$  is the check function defined by

$$w_{\tau} = \begin{cases} \tau & \text{if } y_i > \mu(\tau) \\ 1 - \tau & \text{if } y_i \le \mu(\tau) \end{cases}$$

for some population expectile  $\mu(\tau)$  for different values of an asymmetric parameter  $0 < \tau < 1$ .

• Expectiles are obtained by solving

$$\tau = \frac{\int_{-\infty}^{e_{\tau}} |y - e_{\tau}| f_y(y) dy}{\int_{-\infty}^{\infty} |y - e_{\tau}| f_y(y) dy} = \frac{G_y(e_{\tau}) - e_{\tau} F_y(e_{\tau})}{2(G_y(e_{\tau}) - e_{\tau} F_y(e_{\tau})) + (e_{\tau} - \mu)}$$

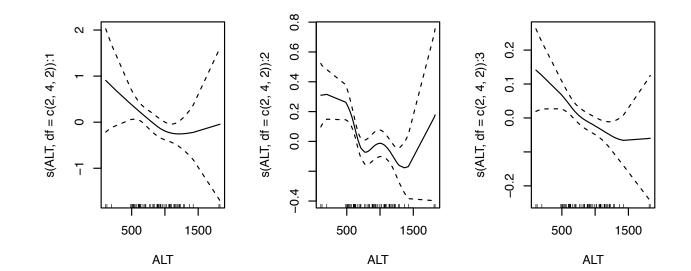
where

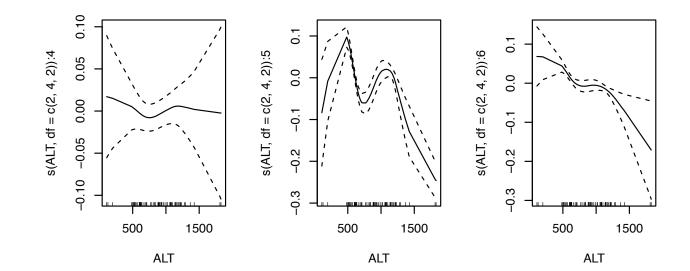
 $-f_y(\cdot)$  and  $F_y(\cdot)$  denote the density and cumulative distribution function of y.

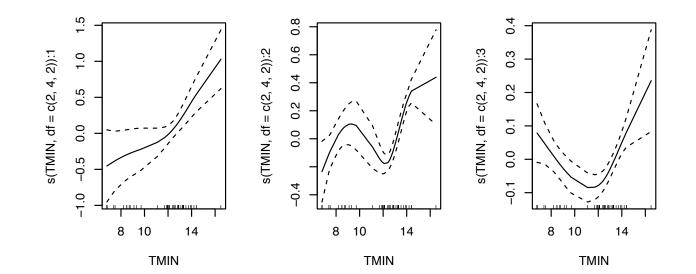
 $-G_y(e) = \int_{-\infty}^e y f_y(y) dy$  is the partial moment function of y and

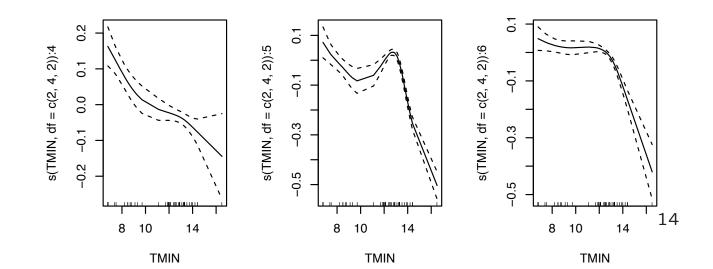
 $-G_y(\infty) = \mu$  is the expectation of y.

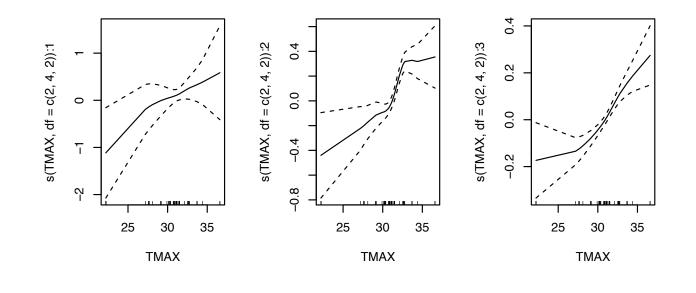
Worked example - binomial data

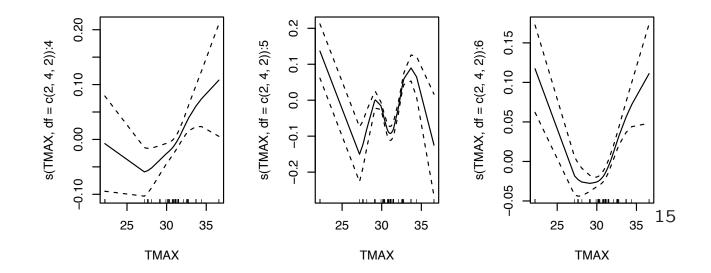












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