

Exploring the scale dependence of Lagrangian bias

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Motivation

Scale dependence of halo bias can be a powerful probe of cosmology, potentially testing the initial conditions (e.g., Dalal et al. 08; Slosar et al. 08; many others) or the nature of gravity (e.g., Parfrey, Hui & Sheth 11; Lam & Li 12).

Problem is, halo bias in vanilla LCDM is *also* scale dependent due to the nature of Einstein/Newtonian gravity (e.g., Chan et al. 12; Baldauf et al. 12).

Analytical predictions traditionally rely on excursion set approach and “peak-background split” argument. Recent work on excursion sets with correlated steps (see Marcello’s talk) shows that scale-dependence is expected even in the *initial* conditions.

We explore this here.

Caveat: Everything in this talk is Lagrangian!

Outline

- Fourier space versus real space
- Scale dependence from random walks
- (Detour: Excursion Set Peaks)
- Numerical tests and comparisons with N -body simulations

Traditional estimates of bias

$$b_1^2(k) \equiv \frac{P_{hh}(k)}{P_{mm}(k)} \quad \text{or} \quad b_1(k) \equiv \frac{P_{hm}(k)}{P_{mm}(k)}$$

where $P_{hh}(k) = \langle \delta_h^2 \rangle$, $P_{mm}(k) = \langle \delta^2 \rangle$ and $P_{hm}(k) = \langle \delta_h \delta \rangle$.

These typically show scale-dependence $\sim k^2$ with **constant** value at **large scales**, usually compared with “peak-background split”

(Kaiser 84; BBKS 86; Mo & White 96)

$$b_{n0} = f^{-1} \left(-\frac{\partial}{\partial \delta_c} \right)^n f,$$

with halo mass function $f(\delta_c; m)$.

Estimates beyond linear bias require measuring higher order correlations (e.g., quadratic bias requires bispectrum) and/or assumptions regarding locality/stochasticity/scale dependence.

From Fourier space to Real space

Peak-bg split argument works in real space. Not immediately obvious why/how this should be compared with Fourier space measurement.

Consider a toy example:

$$\delta_0(\mathbf{k}) = \delta(\mathbf{k})W(kR_0) \quad ; \quad \delta_h(\mathbf{k}) = b(k|m)\delta(\mathbf{k})W(kR)$$

with $m \propto R^3$ and $W(y) = e^{-y^2/2}$.

Natural real-space definition of linear bias could be $b(m, R_0) = \langle \delta_h \delta_0 \rangle / \langle \delta_0^2 \rangle$. Then,

$$b(k|m) = b_{10}(m) \implies b(m, R_0) = (S_\times / S_0) b_{10}$$

$$\left[S_0 = \int d \ln k \Delta(k) W(kR_0)^2 \quad ; \quad S_\times = \int d \ln k \Delta(k) W(kR) W(kR_0) \right]$$

AP & Sheth 12

So even constant Fourier-space bias will lead to scale dependence in real space.

Real-space bias from random walks

Excursion set approach makes the peak-bg argument rigorous (*Mo & White 96*).

Calculation proceeds by writing a conditional mass function

$$f(\delta_c, m | \delta_0, R_0) = f(\delta_c, s | \delta_0, S_0)$$

and then Taylor expanding. $\left[s \equiv \int d \ln k \Delta(k) W(kR)^2 ; m \propto R^3 \right]$

For random walks with uncorrelated steps $f(\delta_c, s | \delta_0, S_0) = f(\delta_c - \delta_0, s - S_0)$ so Taylor expansion precisely recovers peak-bg results provided $S_0 \ll s$.

Two issues:

- (1.) We'd like to compute cross-correlations, not Taylor coefficients.
- (2.) We'd like to use random walks with correlated steps.

It's possible to do (1.) alone.
But it's easier to do (1.) and (2.) simultaneously.

Scale dependence from correlated steps

From Marcello's talk we know an accurate analytical first crossing distribution

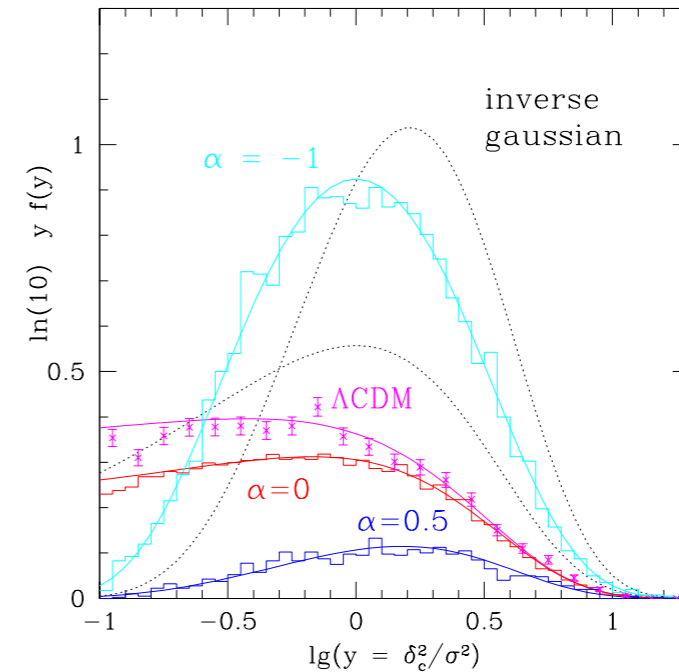
$\nu \equiv \delta_c / \sqrt{s}$; $x \sim$ walk slope

$\gamma = \langle x \nu \rangle \sim$ width of power spectrum

$p_G(x - \mu; \Sigma^2)$: Gaussian with mean μ
and variance Σ^2

$$f_{MS}(\nu) = \frac{e^{-\nu^2/2}}{\sqrt{2\pi}} \frac{1}{\gamma\nu} \int_0^\infty dx x p_G(x - \gamma\nu; 1 - \gamma^2)$$

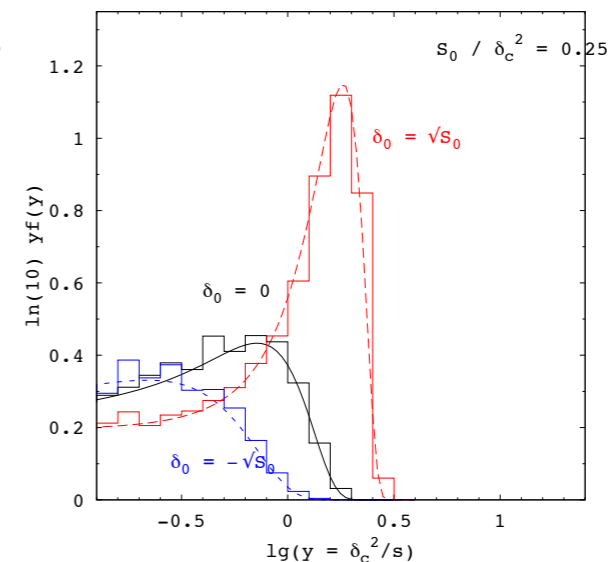
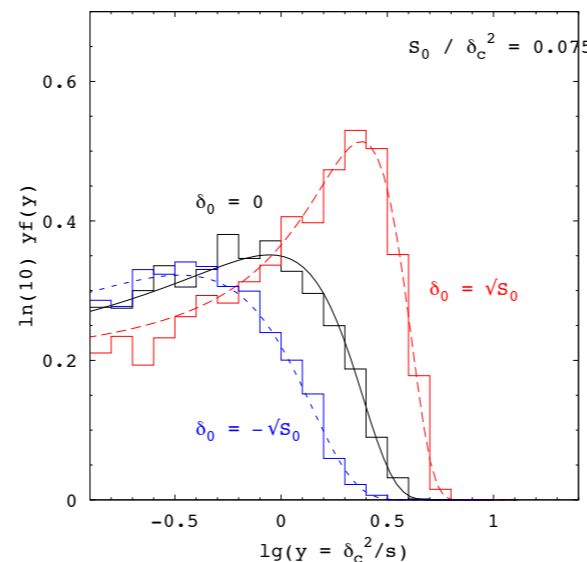
$$= \int_0^\infty dx \frac{x}{\gamma\nu} p(x, \nu) \quad (\text{Musso \& Sheth 12})$$



which also leads to an accurate conditional distribution

$$f(\nu|\delta_0) = \int_0^\infty dx \frac{x}{\gamma\nu} p(x, \nu|\delta_0)$$

Musso, AP & Sheth 12



Scale dependence from correlated steps

Musso, AP & Sheth 12

Define $\langle \rho_h | \delta_0; S_0 \rangle \equiv f(\nu | \delta_0; S_0) / f(\nu)$

Cross-correlation $\langle \rho_h \delta_0 \rangle = \int d\delta_0 p_G(\delta_0; S_0) \delta_0 \langle \rho_h | \delta_0; S_0 \rangle$ is analytic, and has nice properties:

- **Structure of linear bias**

$$b_1 \equiv \langle \rho_h \delta_0 \rangle / S_0 = (S_\times / S_0) [b_{10} + \epsilon_\times b_{11}]$$

where $\epsilon_\times = 2 d \ln S_\times / d \ln s$; $b_{10} = -\partial \ln f / \partial \delta_c$; $\delta_c b_{11} = \nu^2 - \delta_c b_{10}$

- **Interpretation of ϵ_\times**

Suppose that in Fourier space: $b_1(k) = b_{10} + (k^2 s / \sigma_1^2) b_{11}$ $\left[\sigma_1^2 \equiv \int d \ln k \Delta(k) k^2 W(kR)^2 \right]$

Then in real space: $b_1 = (S_\times / S_0) [b_{10} + \epsilon_\times b_{11}]$

Scale dependence from correlated steps

Musso, AP & Sheth 12

Define $\langle \rho_h | \delta_0; S_0 \rangle \equiv f(\nu | \delta_0; S_0) / f(\nu)$

Cross-correlation $\langle \rho_h \delta_0 \rangle = \int d\delta_0 p_G(\delta_0; S_0) \delta_0 \langle \rho_h | \delta_0; S_0 \rangle$ is analytic, and has nice properties:

- **Extension to nonlinear bias**

$b_n \equiv S_0^{-n/2} \langle \rho_h H_n(\delta_0 / \sqrt{S_0}) \rangle$ has similar properties:

$b_n = (S_\times / S_0)^n \sum_{r=0}^n \binom{n}{r} b_{nr} \epsilon_\times^r$ where $b_{n0} = f^{-1} \left(-\frac{\partial}{\partial \delta_c} \right)^n f$, and linear relations between b_{nr}

- **Suggests simple measurement prescription** $\hat{b}_n = S_0^{-n/2} \sum_{i=1}^N H_n(\delta_{0i} / \sqrt{S_0}) / N$

- **Peak-background split** is the large scale limit of Fourier-space bias, and can be recovered from *finite scale* measurement using linear relations.

- **Extends to Excursion Set Peaks** with same structure, different details.

Detour: Excursion Set Peaks

Mass Function

- **Constant threshold:** $B = \delta_c$

(Appel & Jones 90)

$$f_{\text{ESP}}(\nu) = \frac{m}{\bar{\rho}V_*} \int_0^\infty dx \frac{x}{\gamma\nu} F(x) p(x, \nu) \xrightarrow{\nu \gg 1} \frac{m}{\bar{\rho}V_*} \gamma^3 \nu^3 f_{\text{MS}}(\nu)$$

x = peak curvature; $F(x)$ = BBKS curvature weight; V_* = characteristic peak volume

- **“Moving” threshold:** $B(\sigma) \quad \left[\sigma^2 = s \right]$

AP & Sheth 12

$$f_{\text{ESP}}(\nu) = \frac{m}{\bar{\rho}V_*} \frac{1}{\gamma\nu} \int_{\gamma B'}^\infty dx (x - \gamma B') F(x) p(x, B/\sigma)$$

- **“Ellipsoidal” threshold:** $B = \delta_c + \beta\sigma$ with stochastic β

$$\begin{aligned} f_{\text{ESP}}(\nu) &= \int d\beta p(\beta) f_{\text{ESP}}(\nu|\beta) \\ &= \frac{m}{\bar{\rho}V_*} \frac{1}{\gamma\nu} \int d\beta p(\beta) \int_{\beta\gamma}^\infty dx (x - \beta\gamma) F(x) p(x, \nu + \beta) \end{aligned}$$

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Detour: Excursion Set Peaks

Bias

Bias calculation is identical to that for traditional excursion sets, but keeping track of barrier stochasticity.

E.g., linear bias is given by

$$\delta_c b_1 = \left(\frac{S_x}{S_0} \right) \frac{\int d\beta p(\beta) \mathcal{B}_{1,\text{ESP}}(\nu, \epsilon_x | \beta)}{\int d\beta p(\beta) f_{\text{ESP}}(\nu | \beta)},$$

where

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$$\begin{aligned} \mathcal{B}_{1,\text{ESP}}(\nu, \epsilon_x | \beta) \equiv & \left(\frac{m}{\bar{\rho} V_*} \right) \frac{e^{-(\nu+\beta)^2/2}}{\sqrt{2\pi}} \\ & \times \frac{1}{\gamma\nu} \int_{\beta\gamma}^{\infty} dx (x - \beta\gamma) F(x) p_G(x - \beta\gamma - \gamma\nu; 1 - \gamma^2) \\ & \times \left[\nu(\nu + \beta) - (1 - \epsilon_x) \frac{\gamma\nu}{1 - \gamma^2} (x - \beta\gamma - \gamma\nu) \right] \end{aligned}$$

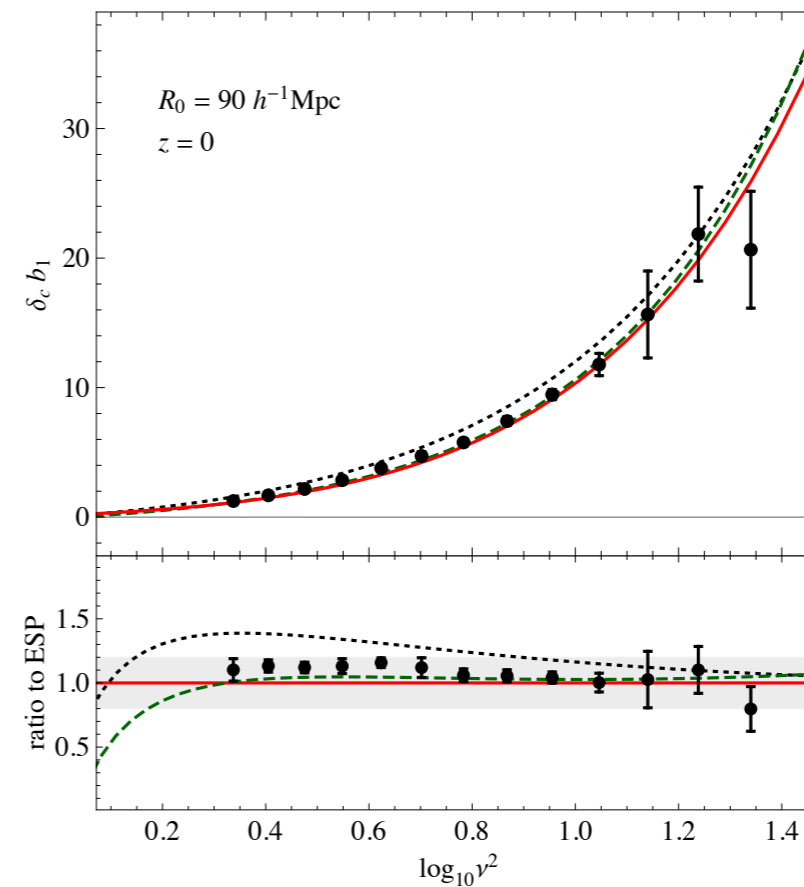
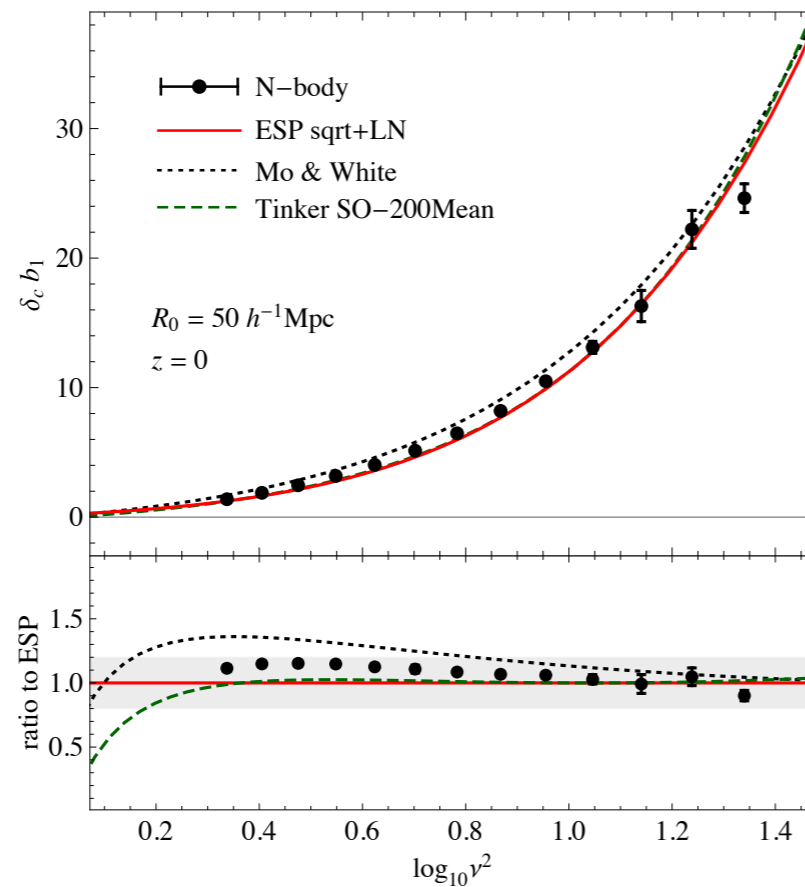
Numerical tests: N-body simulations

$$b_n \equiv S_0^{-n/2} \left\langle \rho_h H_n(\delta_0 / \sqrt{S_0}) \right\rangle$$

Measurement:

- Find all N **halos** in mass bin $(m, m + dm)$
- Choose S_0 (i.e., R_0)
- Estimate, e.g., $\hat{b}_1 = S_0^{-1} \sum_{i=1}^N \delta_{0i} / N$

AP, Sefusatti, Chan et al. 13



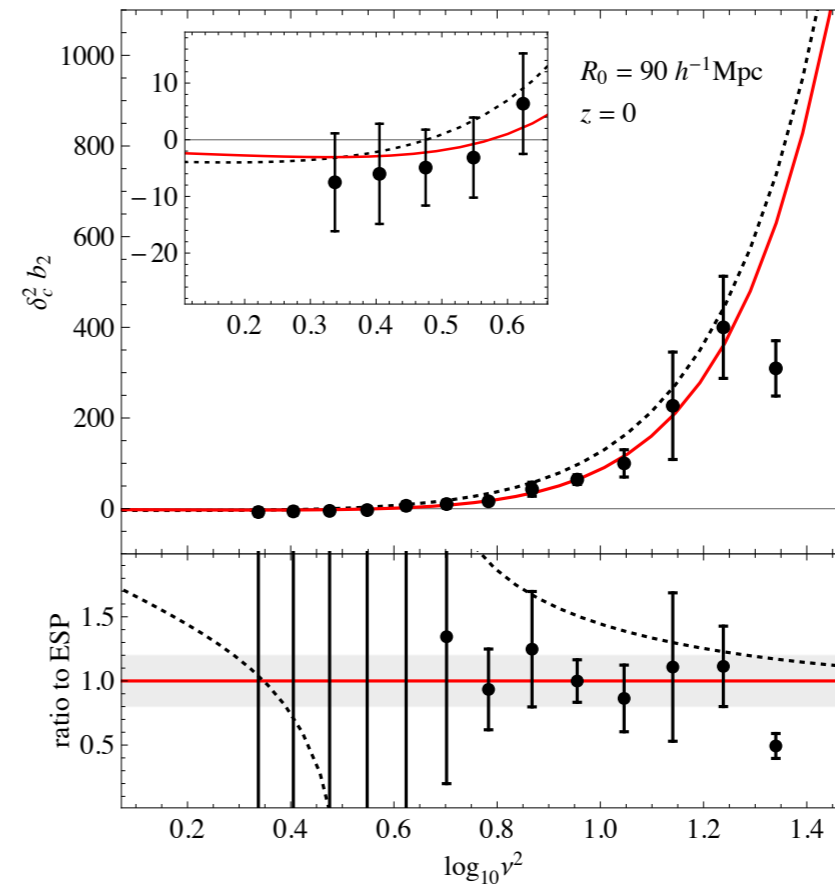
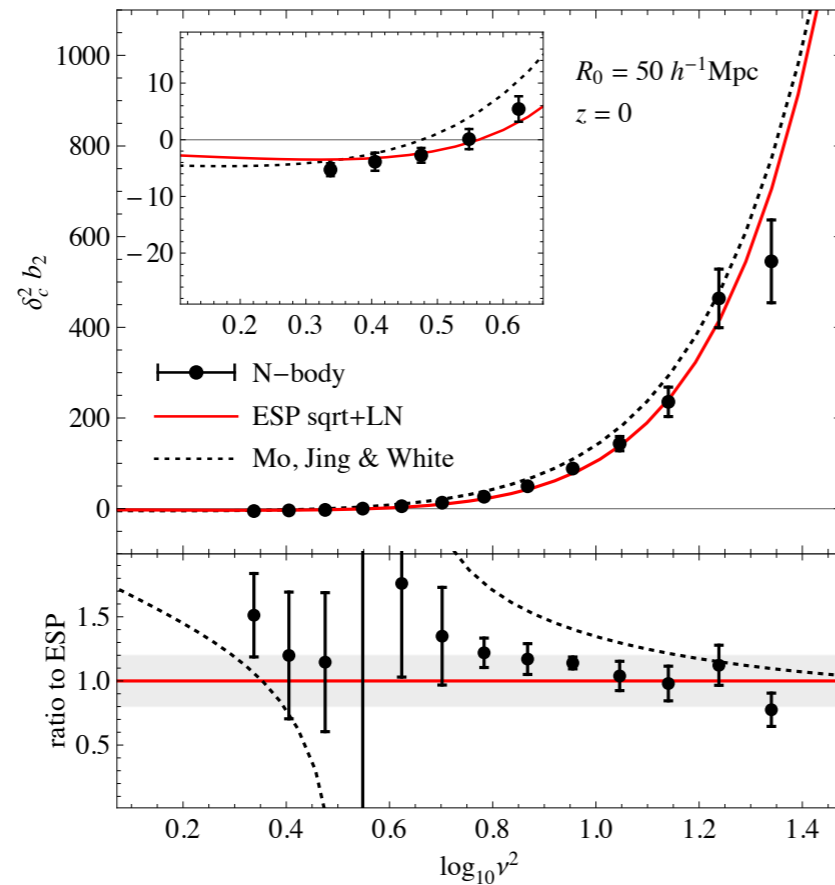
Numerical tests: N-body simulations

$$b_n \equiv S_0^{-n/2} \left\langle \rho_h H_n(\delta_0 / \sqrt{S_0}) \right\rangle$$

Measurement:

- Find all N **halos** in mass bin $(m, m + dm)$
- Choose S_0 (i.e., R_0)
- Estimate, e.g., $\hat{b}_2 = S_0^{-1} \sum_{i=1}^N (\frac{\delta_{0i}^2}{S_0} - 1) / N$

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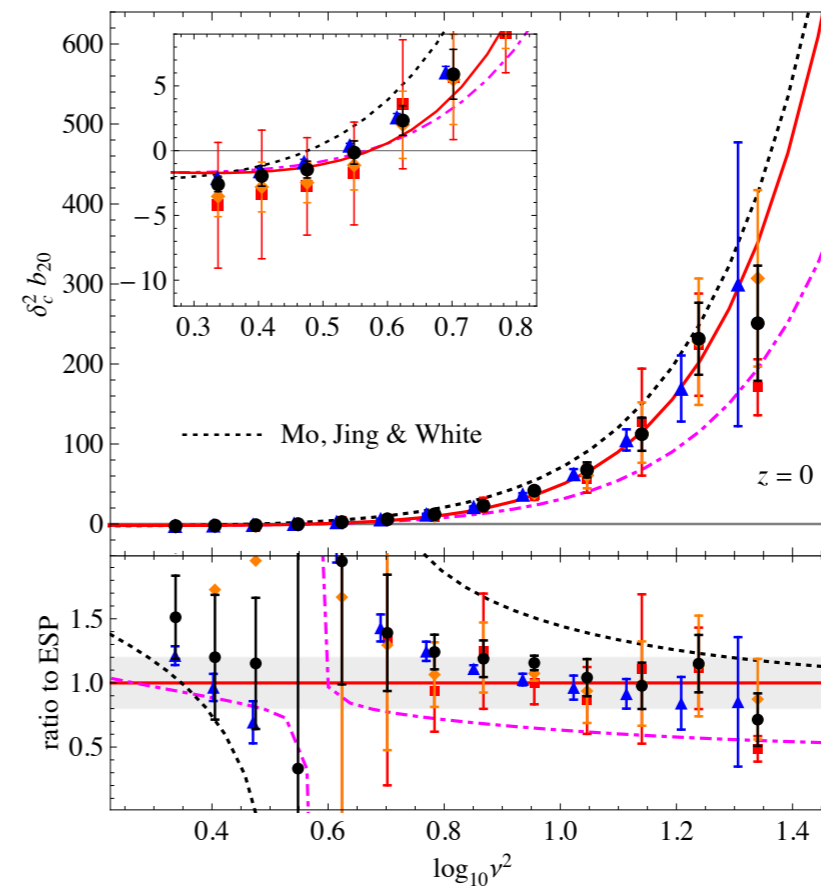
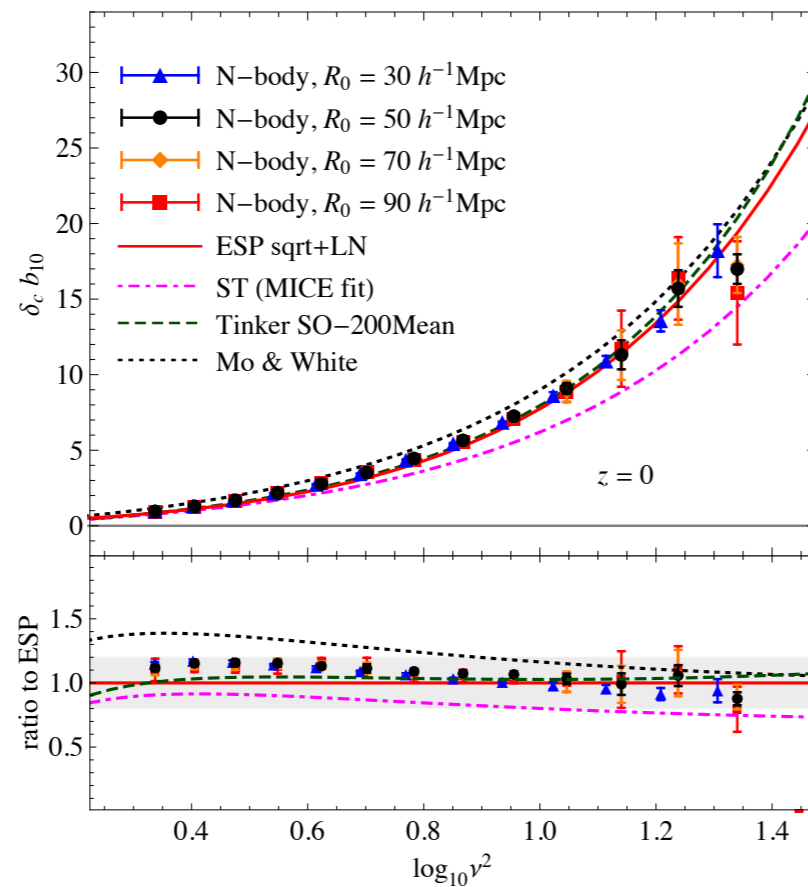
Numerical tests: N-body simulations

$$b_n \equiv S_0^{-n/2} \left\langle \rho_h H_n(\delta_0 / \sqrt{S_0}) \right\rangle$$

Measurement:

- Find all N **halos** in mass bin $(m, m + dm)$
- Choose S_0 (i.e., R_0)
- Estimate, e.g., \hat{b}_1, \hat{b}_2 and reconstruct $\hat{b}_{10}, \hat{b}_{20}$

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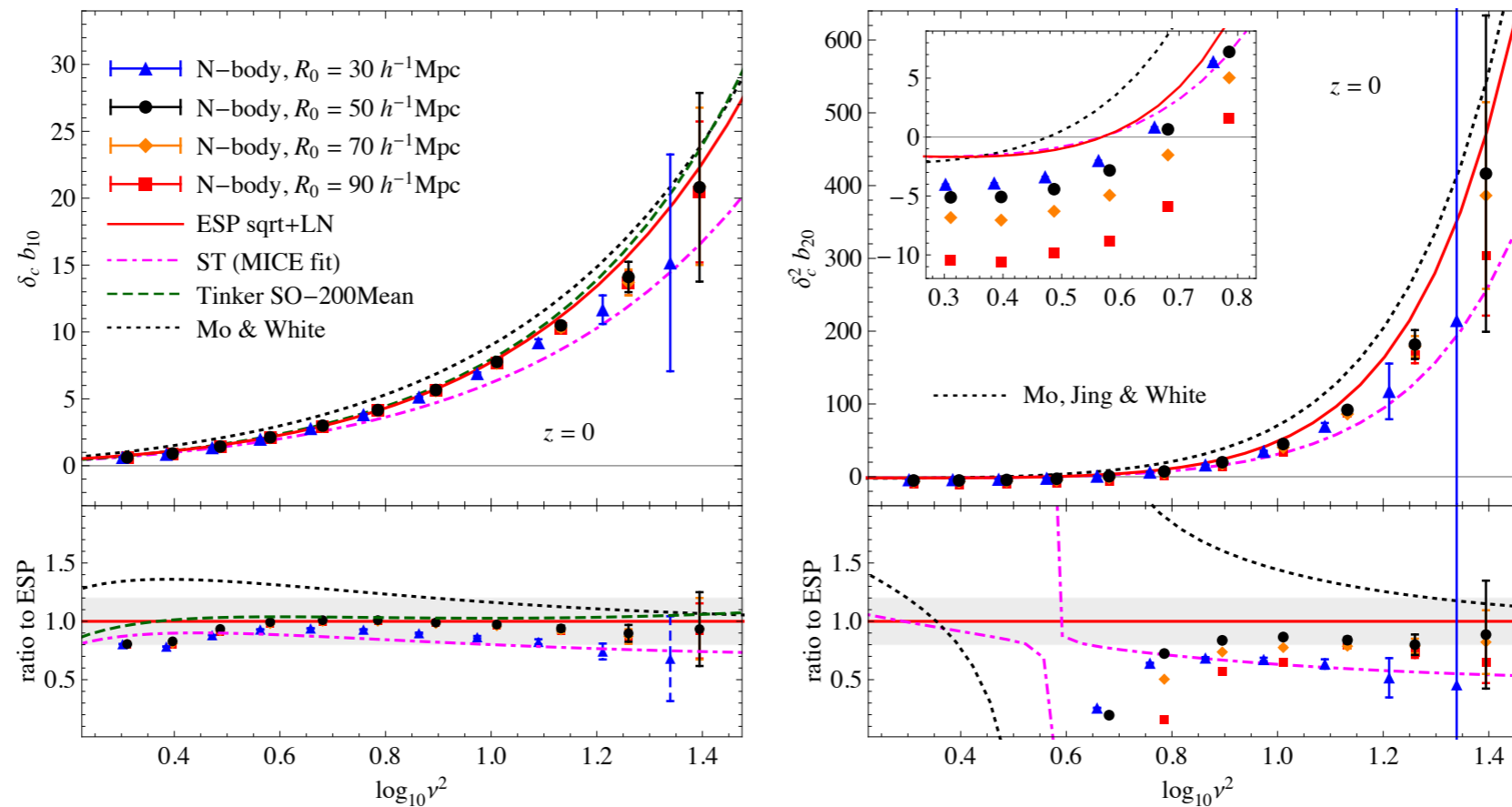
Numerical tests: PINOCCHIO

$$b_n \equiv S_0^{-n/2} \left\langle \rho_h H_n(\delta_0 / \sqrt{S_0}) \right\rangle$$

Measurement:

- Find all N **halos** in mass bin $(m, m + dm)$
- Choose S_0 (i.e., R_0)
- Estimate, e.g., \hat{b}_1, \hat{b}_2 and reconstruct $\hat{b}_{10}, \hat{b}_{20}$

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Conclusions

- Random walks with correlated steps predict scale-dependent Lagrangian bias.
- This extends to excursion set peaks (ESP) calculations too.
- A simple prescription leads to measurements in simulations that are directly comparable to the analytical prediction.

$$\hat{b}_n = S_0^{-n/2} \sum_{i=1}^N H_n(\delta_{0i}/\sqrt{S_0})/N$$

- ESP makes accurate predictions for linear and quadratic bias coefficients.
- Measurement prescription is a fast, useful consistency check for “semi-analytic” algorithms (we tested PINOCCHIO).

Still to do:

- Scale-dependence beyond density
- Predictions for Eulerian bias
- Relevance for observations (?)

Thank you!